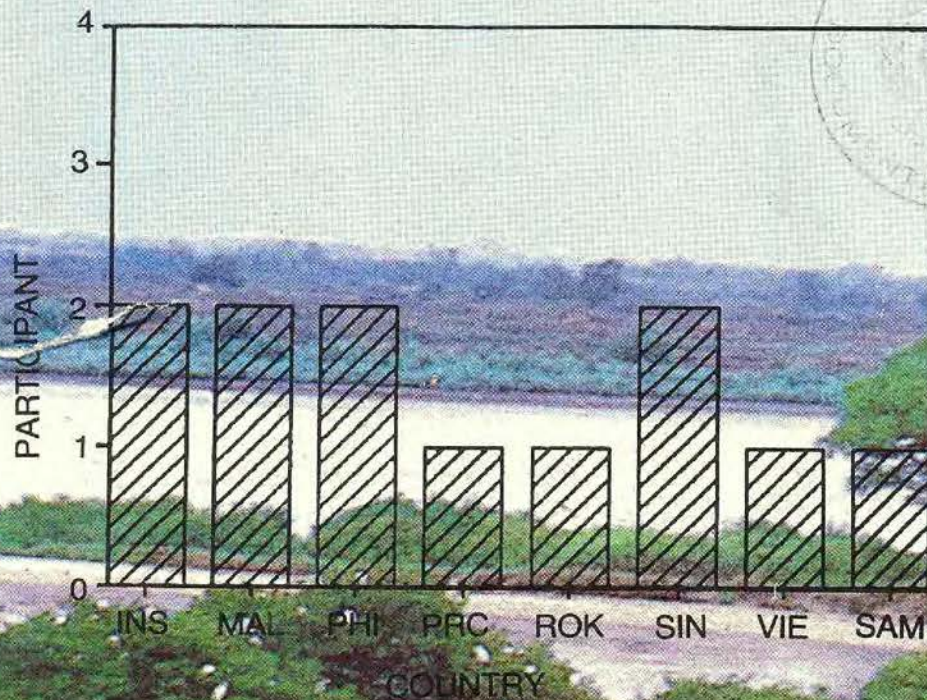


# Computer Analysis for Environmental Biologists

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The cover photograph illustrates the mangrove ecosystem on the Pulau Dua Nature Reserve, West Java, Indonesia. Reproduced from an Indonesia MAB poster on "Cintaku Negeriku" series with permission.

**REPORT**  
**ON**  
**THE UNESCO-MAB/UNEP/NUS REGIONAL TRAINING COURSE**  
**ON**  
**COMPUTER - BASED QUANTITATIVE METHODS FOR**  
**ENVIRONMENTAL BIOLOGISTS**

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## PREFACE

It is appropriate that a training course in computer-based quantitative methods for environmental biologists be held at this time and it is especially appropriate that it be held in Singapore. For developing countries, "development" must include the ability to use computers for many purposes, and one of the areas where computer skills are essential is certainly environmental biology. With the rapid and continuing increase in availability of microcomputers and softwares, there is no longer any barrier to the use of computer-based methods. Furthermore, it is advisable that training be given in use of state-of-the art systems rather than on the microcomputer systems of several years ago. Obsolescence is a problem in any area, but nowhere does it advance more rapidly than in the area of microcomputer technology and software development. We should not compound the problem by training for the use of out-of-date systems. In any case the newer systems and the software available to run on them is usually "simpler" for the user than the earlier systems were.

Singapore is an appropriate venue for such a course because as a country it has recognised the importance of the computer-based technological revolution. This is particularly true at the National University of Singapore where an excellent mainframe computer system and microcomputers linked in a network connected with the main system are available to students and staff. We hope that this training course will be just the first of a series of similar courses to be given in this region.

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# TRAINING COURSE ON COMPUTER-BASED QUANTITATIVE

## METHODS FOR ENVIRONMENTAL BIOLOGISTS

### 1. GENERAL INTRODUCTION

#### 1.1 General objectives and approach

The emphasis is on biological modelling, especially the fitting of bivariate models and the use of them to test meaningful biological hypotheses.

Practical experience in application of the techniques is emphasized. Theory taught in lectures should be applied by doing the practical exercises. It should be emphasized that attendance at lectures without also taking time to do the practical exercises will be of limited value.

The biological modelling will be computer-based because it is the new micro and mainframe computer technology that allows biologists to effectively and efficiently model and analyze their data in ways that were impossible even a decade or two ago. However, this is not a "computer course", micro or mainframe, nor is it a course in computer programming. It is on modelling and it is aimed at biologists. A variety of computer hardware and software will be used to illustrate the diverse "tools" that are available to do biological modelling. But we shall attempt not to lose sight of the forest (biological modelling) as we become surrounded by trees (computer hardware and software).

The technology re. both hardware and software is advancing rapidly. You will always be out-of-date in the sense that someone, somewhere, has probably developed or is developing a better system for doing the job a biologist wants to do. Therefore present availability "back home" will not be the overriding criterion for the choice of hardware and software to be used in this course. In any country or region the system



available now is not what will be available in a few years time, so that an attempt to train for use of present facilities would very soon become training for the past. Statistical packages that required large mainframe computers a few years ago now run on microcomputers. The "networking" of microcomputers so they can be used as terminals to access mainframe computers will more and more become common practice. This course will utilize a wide variety of hardware and software so that participants can gain experience with different systems, and then back in their countries they can help influence the direction of development of systems most useful to biologists. In summary, this course will train participants in use of hardware/software systems which will be available in the near future, even if they are not available in all countries now.

#### 1.2. Structure of course

This report is based on a 3-week training course. Each day, Monday to Friday, was spent in 2 hr 15 min of lecture theory) in the morning and 3 hr of practical ("hands on") work in the afternoon. A library consisting of most of the references cited in the bibliography was available throughout the course. Presentations by participants were held on the one public holiday (May 1) and on the second Saturday.

## 2. INTRODUCTION TO HARDWARE AND SOFTWARE SYSTEMS

### 2.1 General remarks

As are noted in section 1.1, sophisticated languages and statistical packages are rapidly becoming available on microcomputers. One is no longer limited to the BASIC language, or to amateurishly written (sometimes incorrect) statistical programs. Most languages (FORTRAN, PASCAL, APL included) are now available for IBM PC-compatible systems, and some of them are available for APPLE-compatible systems. All of the major statistical packages used on mainframe computers (MINITAB, SAS, SPSS, BMD) are now available in "micro" versions, as well as a number of others especially designed for microcomputers.

However we are aware that many participants in the course, and readers of this report, are probably naive regarding languages other than BASIC, regarding use of statistical packages, and regarding use of mainframe computers. Therefore we have taken care to describe the general availability of the software we use in the course, in addition to commonly available hardware systems. Then we describe in great detail how to get started using the various hardware/software combinations.

The choices of languages and software are personal and subjective, and based on the senior author's experience. BASIC is ubiquitous on microcomputers and could not be omitted. FORTRAN is the original profession scientific programming language, and many excellent programs are available. APL is not known as well as it should be, and is an excellent language for statistics and modelling. MINITAB is a good "friendly" beginning package for the most commonly used univariate statistical methods. SAS complements MINITAB by providing a wide variety of specialized methods (e.g., probit analysis) and multivariate methods (e.g., cluster analysis, principal components analysis, discriminant analysis). Other languages or packages might have been chosen, but these are appropriate ones.

2.2 Hardware & software Information2.2.1. COMMENTS ON GENERALITY AND AVAILABILITY

Languages & Packages	OPERATING SYSTEM			
	Apple/DOS 3.3	Apple/CPM	IBM PC/MSPDOS	IBM mainfram/VM CMS
BASIC	This is the APPLE specific operating system. It is not on other micros. It uses a 6502 microprocessor. Some BASIC commands are APPLE-specific. 40 columns only.	This is a more general operating system, based on a Z80 microprocessor and available on a variety of micros. 80 columns.	The IBM-standard operating system BASIC is available in several versions. We will use MICROSOFT BASIC.	The VM/CMS operating system. own IBM BASIC.
FORTRAN		Available (but we will not run FORTRAN on the APPLE)	Available in several versions	Several versions of FORTRAN are on this system — we will use VS FORTRAN
APL			Available in several versions (we can run APL on the IBM PC)	IBM VS APL
MINITAB			Available for IBM PC with >256K RAM (but we will not use MINITAB on the PC)	Created originally by Dept. of Statistics, Pennsylvania State University U.S.A.
SAS			Available for IBM PC with >256K RAM (but we will not run SAS on the PC)	Created originally by Dept. of Statistics North Carolina State University U.S.A.

2.2.2. REGRESSION AND SCATTERPLOT PROGRAM & PROCEDURES (+ NECESSARY UTILITIES)

Languages & Packages	OPERATING SYSTEM			
	Apple/DOS 3.3	Apple/CPM	IBM PC/MS DOS	IBM mainfram/VM CMS
BASIC	MAKE TEXT CREATE TEXT * REGRESSION PLOT *programs by Orloci and Kenkel	CREATE LINREG *  *program CREATE by Green program LINREG by Somers	LINREG *  *program by Somers	LINREG *  * program by Somers
			REGR * PLOT	REGR PLOT *
FORTRAN			*programs created or greatly modified by Green	* programs created or greatly modified by Green
APL				GLM * SCATTERPLOT  *program GLM by Simillie, modified by Bailey  program SCATTERPLOT by Anscombe, modified by Green.
				READ and SET procedures.  REGR procedure PLOT procedure
MINITAB				DATA step
SAS				GLM procedure  PLOT procedure

2.2.3. ANALYSIS OF COVARIANCE & MULTIVARIATE ANALYSES.

Languages & Packages	OPERATING SYSTEM			
	Apple/Dos 3.3	Apple/CPM	IBM PC/MS DOS	IBM mainframe/VM CMS
BASIC	Orloci & Kenkel programs: PCAR (PCA) ALC (avr link clustering) SSA (sum of sqrs clustering) WEIGHING/SCP (variable subset selection, as done by RSLCT FORTRAN)	ANOVA (by K. Somers, modified by R. Green)	ANOVA (by K. Somers, modified by R. Green)	ANCOVA (by K. Somers, modified by R. Green)
FORTRAN				RSLCTIBM (by R. Green, from an algorithm by L. Orloci)
APL			APL operators can be used to do multivariate analyses	MATFORM R. Bailey COVAR  PDET Ramsey & Musgrave ISOTROPY F. Anscombe  GEIG R. Green and APL operators.
MINITAB				EIGEN, matrix algebra commands, and other commands.
SAS				PROC PRINCOMP PROC CLUSTER PROC CANDISC and other procedures.

## 2.3 Basic operation instructions

### 2.3.1 To start-up

#### 2.3.1.1 IBM PC

Insert the DOS 2.0 diskette into drive A.  
 Switch on the computer switch at the right hand side.  
 Wait.  
 The following will appear:

```
A>wtdatim
Current date (DD-MM-YY):01-01-1980
Enter new date:   Press ENTER

Current time:00:01:00
Enter new time:   Press ENTER

A>REM The IBM Personal Computer DOS
A>REM Version 2.00 (C)Copyright IBM Corp 1981,1982,1983
A>
```

You can now proceed.

#### 2.3.1.2 APPLE - DOS 3.3

Insert the DOS 3.3 diskette into drive A.  
 Switch on the computer switch at the back on the left side, and the monitor knob on top.  
 Adjust the switch attached to the right hand side of the monitor to the 40 column-number.  
 The following will appear:

```
DOS VERSION 3.3
APPLE II PLUS OR ROMCARD          SYSTEM MASTER

(LOADING INTEGER INTO LANGUAGE CARD)
```

]

Insert diskette containing required programmes into drive B.  
 Type in CATALOG, D2 to obtain a list of the files.

### 2.3.1.3 APPLE - CP/M

Adjust the switch on the right of the monitor to 80 column-number.

Insert the CP/M diskette into drive A.

Switch on the computer.

The following will appear:

```
Apple ][ CP/M
56K Ver. 2.20B
(C) 1980 Microsoft
A>
```

Insert diskette containing required programmes into drive B.

To obtain a list of the files in this diskette, type A>dir B:

### 2.3.1.4 MAINFRAME - IBM 3081 VM/CMS

The terminal and two of the IBM PCs are connected to the mainframe over at the Computer Centre.

Before you can use the system, you must have a user identification (userid) and password.

#### 2.3.1.4.1 The TERMINAL

Switch on the terminal. After a short while, the logo:

```
NUS
VM/SP
```

should appear. Press ENTER key and logo disappears.

If, for example, your userid and password are DEMO1, log on to the system with the LOGON command, as follows:

```
Type LOGON DEMO1
ENTER PASSWORD: DEMO1      Press ENTER
```

For security purposes, the password you enter is not displayed on the screen.

After logging-on, type the following:

CP DEF STOR 1500K	Press ENTER
IPL CMS	Press ENTER
	Press ENTER A second time

Then you can carry on with XEDIT, MINITAB, SAS, BASIC, FORTRAN, APL, etc.

#### 2.3.1.4.2 IBM PC used as Terminal

Insert the DOS 2.0 diskette into drive A and the IRMA Rev 1.10 diskette into drive B.

After starting-up, type:

A>B:

and then B>e78 Press ENTER

The NUS logo will appear. Log on as described above.

#### 2.3.1.4.3 Some commands for IBM VM/CMS Operating System

LOGON acctname	The log-on procedure - the system will respond with a request for your password
DEF STOR 1500K	A request for temporary memory available to you to be increased to 1500K. This is necessary for running SAS, APL & certain other software. It is best to always do it. Following it you must return to CMS by entering 'I CMS' and depressing ENTER twice.
LIST	Lists all your files.
LIST fn ft	Lists a particular file, with name = fn and type = ft. For example, if you have a BASIC program in a file named 'REGR', then fn = REGR and ft = BASIC.
LIST fn	Lists all files with that fn. For example, if there is a file 'REGR BASIC' with a BASIC program in it and a file 'REGR DATA' with the data to be analysed in it, then both files would be listed if you entered 'LIST REGR'.



LIST * ft	Lists all files with that ft. For example, if you entered 'LIST * BASIC', then all files containing BASIC programs would be listed.
TYPE fn ft	Types the contents of that file on the terminal screen.
ERASE fn ft	Erases that file from your disk area. Use with care!
ALT + CLEAR	Clears the screen so the next screenful can be shown.
PRINT fn ft	Prints out the contents of that file (at the NUS Computer Centre).
LPRINT B08 fn ft	Prints out the contents of the file in the Computer Science Dept. lab room we will use (Comp Sci S15 02-11).
LPRINT C08 fn ft	Prints out the contents of the file at the Computer science Dept. printer that is operator-covered (one floor below 'our' lab room).
LOGOFF	The log off procedure ( <u>and</u> switch off the terminal!).

### 2.3.2 To back-up files

#### 2.3.2.1 IBM PC

After starting-up with DOS, type as follows:

```
A>diskcopy a: b:
```

The message will appear:

```
Insert source diskette in drive A:
Insert target diskette in drive B:
Strike any key when ready.
```

#### 2.3.2.2 APPLE - DOS 3.3

Insert DOS 3.3 into drive A and empty diskette into drive B.  
To copy files, issue the command:

]RUN COPYA, D1

The following will appear:

APPLE DISKETTE DUPLICATION PROGRAMME

ORIGINAL SLOT: DEFAULT=6                    Press RETURN

                  DRIVE: DEFAULT=1            Press RETURN

DUPLICATE SLOT: DEFAULT=6                  Press RETURN

                  DRIVE: DEFAULT=2            Press RETURN

--- PRESS 'RETURN' KEY TO BEGIN COPY ---

### 2.3.3 Formatting a new disk

#### 2.3.3.1 IBM PC

To format new diskettes:

A>format b:

Insert new diskette into drive B:

and strike any key when ready.

#### 2.3.3.2 APPLE - DOS 3.3

To initialise new diskettes, insert new diskette into drive B.

]INIT HELLO, D2

### 2.3.4. To run a BASIC program

#### 2.3.4.1 IBM PC

Start-up with DOS 2.0. Type in:

A>basic                    Press ENTER

The screen will show:

The IBM Personal Computer Basic  
 Version D2.00 Copyright IBM Corp. 1981, 1982, 1983  
 61330 Bytes free

OK

-

1LIST 2RUN 3LOAD" 4SAVE" 5CONT 6LPT1 7TRON .....

You can now start running the program by typing in:

LOAD" (name of the file)

RUN (name of the file)

The file will be retrieved and the program will run.

To save time typing in the command LOAD and RUN, the function keys on the left of the keyboard can be used. With reference to the line printed at the bottom of the screen,

F3 is for the command LOAD

F2 is for the command RUN

#### 2.3.4.2 APPLE - DOS 3.3

Type:

LOAD (name of the file)

RUN (name of the file)

#### 2.3.4.3 APPLE - CP/M

Type: (note use of quotation marks)

LOAD " (name of the file)

RUN " (name of the file)

#### 2.3.4.4 MAINFRAME - IBM 3081 VM/CMS

File with BASIC program in it must have filetype = BASIC

To go into BASIC mode:

```
basic
IBM BASIC/VM Version 1 Release 1.1.....
* run (name of the file)
```

To leave BASIC mode:

```
quit
```

#### 2.3.5. To run a FORTRAN program on the MAINFRAME - IBM 3081 VM/CMS

File with FORTRAN program in it must have filetype = FORTRAN.

Data file to be used by program must have same filename, and filetype = DATA.

To run FORTRAN program that is in the file "fn FORTRAN":

```
fortvs fn
```

Lots of output follows, but the last three lines on the screen should be:

```
DMSLI0740I EXECUTION BEGINS.....
&EXIT
R;T=.....
```

Your output is in file "fn OUTPUT". To see it, enter

```
type fn output
```

#### 2.3.6 To run APL

##### 2.3.6.1 IBM PC

Start up as before, using the Dos 2.0 diskette.

Insert the IBM APL diskette into drive B.

When A> appears, type B: to change drive.

Type:

```
B> APL
```

The following will appear

```
IBM Personal Computer APL
Version 1.00 (C) Copyright IBM Corp. 1983
Produced by IBM Madrid Scientific Center
```

```
CLEAR WS
```

Press the Control key `Ctrl` and the backspace key `<---` (at the top row, right side of the keyboard), to invoke the APL character set. You may now proceed.

#### 2.3.6.2 MAINFRAME - IBM 3081 VM/CMS

The terminal must have an APL character set.

Invoke the APL character set by depressing the ALT key and at the same time the `<---` key that says "APL ON/OFF" on the front of it. The words APL should appear in the middle, under the horizontal line at the bottom of the screen.

To enter APL mode, type:

```
APL
```

To leave APL mode,

```
)OFF
```

## 2.3.7 MINITAB ON THE MAINFRAME - IBM 3081 VM/CMS

### 2.3.7.1. Some commands for using MINITAB

You have logged on and done 'DEF STOR 1500K'. To run MINITAB

- (a) in interactive mode, enter 'MINITAB'. To get out of MINITAB enter 'STOP'. In this mode each entered command produces immediate response on the screen, but you can not get hard copy of your session -- not easily anyway.
- (b) in batch mode, enter 'MINITAB fn' where the MINITAB commands are in a file named fn and with ft = MINITAB. The last command in that file must be 'STOP'. When the job has run, the output will be in a file with the same fn and ft = OUTPUT. You can use the 'PRINT' command to produce hard copy of those two files. (Or you can use 'LPRINT' - see section 2.3.1.4.3)

Please follow the following procedure in doing exercises. Do the complete exercise in interactive mode until you know you have it correct, exactly the way you want it. Write down the MINITAB commands that produce this perfect MINITAB run (including 'STOP' as the last command!), and enter them into a file (use XEDIT) with any fn you want (but it has to be ft=MINITAB). Then run it in batch mode, and then 'PRINT' out the 'fn MINITAB' and 'fn OUTPUT' files (see section 2.3.7.2 for more details).

Two useful things to remember about MINITAB commands are:

- (a) MINITAB only pays attention to the first 4 letters of the command.
- (b) MINITAB pays no attention to words included in the command statement.

This means that you can enter 'REGRESS SIZE IN C1 ON 2 PREDICTOR VARIABLES TEMPERATURE IN C2 AND SALINITY IN C3, STORE STANDARDIZED RESIDUALS IN C4, PREDICTED VALUES IN C5'. Or you

can enter 'REGR C1 2 C2 C3, C4, C5'. Both will work equally well. Why use the longer, wordier version? At least in the beginning it is better because:

- (a) It will help you remember what you are doing and why, so the commands will make more sense and you will remember them more easily.
- (b) When you go back to look at the hard copy, months or years later, it will be much easier to understand.

There will be MINITAB manuals around for you to refer to. Here are a very few commands to get you started:

REGR ---- you have just had this command described.

REGR C1 1 C2 would be the short version, for a simple linear regression if variable Y was stored in column C1 and variable X in column C2. Neither residuals nor predicted Y-values would be stored.

READ C1-C3 would indicate that you will enter a 3-variable data set, one row at a time, and variables 1-3 will be stored in C1-C3, respectively.

SET C4 would indicate that you will enter a column of data, as a string of numbers, and it will be stored in C4.

PRINT C1-C4 would display the contents of C1-C4.

DESCRIBE C1-C4 would provide summary statistics for C1-C4 (number of elements, mean, standard deviation).

COPY C1 INTO C5 would copy the contents of C1 into C5, leaving C1 unchanged.

ERASE C1-C2 would erase the contents of C1 and C2 and leave them empty.

PLOT C1 VERSUS C2 would produce a scatterplot of the variable in C1 versus the variable in C2.

READ FROM 'fn' into C1-C3 would read data from a file named 'fn DATA' into columns C1-C3.

#### 2.3.7.2 Sequence of steps for doing assignments in MINITAB

It is important that you follow this sequence:

1. Set up a blank sheet of paper as shown:
2. If you are going to read in data from a data file, then before entering the MINITAB interactive mode you must enter 'FI 8 DISK fn DATA(PERM)'.
3. Go into the MINITAB interactive mode. (Enter the command 'MINITAB'.)
4. Do the calculations step-by-step, using the MINITAB commands. In interactive mode you get an immediate response to each command. Look at each response to decide whether the correct command was entered. If it was, then write it down on your sheet under "Commands". Use the 'PRINT' command frequently (and write it down each time as well) to see what values are stored in columns, matrices, or constants, before they are used in commands or after they are produced by commands. (If you read in data from a data file, then you must use the 'DISKREAD' command.)



5. After you have done the entire assignment successfully in interactive mode, and have written down on your sheet all the commands needed to do an error-free MINITAB analysis run, then leave MINITAB interactive mode (enter 'STOP').
6. You are now back in CMS. Enter 'XEDIT fn MINITAB' (use whatever fn is appropriate), and you will go into the editor to create a file 'fn MINITAB'. Enter 'INPUT' and you will go into the INPUT mode within the editor. Enter the MINITAB commands that you wrote down when you did the assignment in interactive mode. (Remember to use READ rather than DISKREAD.)
7. When you have finished entering the MINITAB commands, leave INPUT mode by depressing 'ENTER' twice in succession without entering anything. Now you are out of the INPUT mode but still in the editor. Enter 'TOP' to see the file from its beginning. Carefully check what you have entered for errors, and correct any errors using the up, down, left, and right arrows. To get data from a data file ('fn DATA'), position the "active line" (brightly lit) on the "READ --" statement line, and then enter 'GET fn DATA'. The data from 'fn DATA' will be inserted just below the READ -- statement line. Then enter 'FILE' to store the file and get out of the editor.
8. Now enter 'MINITAB fn'. The first response will say that your output is going into 'fn OUTPUT'. The second response (wait for it!) will be 'R;--'. Then continue.
9. Enter 'TYPE fn OUTPUT'. Examine your output on the terminal screen, making sure that it is the same results you obtained when you did the analysis in interactive mode. Notice that each command line or data entry line in the 'fn MINITAB' file produces dashes on a line in the 'fn OUTPUT' file. You may want to get rid of these lines. If so go into the editor again (by entering 'XEDIT fn OUTPUT'), delete the lines, and then enter 'FILE' to leave the editor.
10. Now you probably want a printed copy of both the MINITAB job command file ('fn MINITAB') and the output file ('fn OUTPUT'). Enter 'LPRINT B08 fn MINITAB' immediately

following that enter 'LPRINT BO8 fn OUTPUT'

N.B.: You can not start an assignment involving MINITAB by attempting to create a 'fn MINITAB' file to use for a "batch mode" run!

### 2.3.7.3 MINITAB runs (interactive or batch) with file I/O

Before entering/running MINITAB the input and/or output data files, if there are any, must be identified. To identify an input data file, enter FI 8 DISK fn DATA(PERM and to identify an output data file, enter FI 7 DISK fn DATA(PERM.

The MINITAB command 'DISKREAD' (p. 34 of manual) must be used for input, and the command 'FPUNCH' (p.34 of manual) must be used for output.

It may be easier, when running MINITAB in batch mode, to follow the example given for running SAS in batch mode. That is, use XEDIT to incorporate the data file into the command file, and use 'READ' as if you were in interactive mode. But if you are running MINITAB in interactive mode, and are analysing a large data set that is in a file 'fn DATA', then you have little choice other than to follow the above instructions.

### 2.3.8 Some commands for using XEDIT (the editor on the IBM VM/CMS system)

XEDIT fn ft                      Puts you into the editor, to edit the named file if it already exists, or to create it if it doesn't.

INPUT                              Puts you into INPUT mode. Enter each line. When you depress the ENTER key, you are automatically given a new line to enter. To leave INPUT mode, depress the ENTER key twice in succession.

TOP	Moves the active line to the top line of the file.
BOTTOM	Moves the active line to the bottom line of the file.
UP 10	Moves the active line up 10 lines.
DOWN 15	Moves the active line down 15 lines.
PF8	Moves the active line down one screenfull.
PF7	Moves the active line up one screenfull.
up, down, left and right arrow keys	Use these buttons to move the cursor around and make changes wherever you want. The changes are not stored until you depress the ENTER key the next time.
FILE	To leave the editor and store the file with all the changes you have made. Be careful! You are overwriting the file that existed when you went into the editor!
FILE fn ft	To leave the editor and store the file under the name fn and with the filetype ft. If you give a new fn, then you create a new file, and you do not overwrite the old file.
QUIT	To leave the editor, abandoning all the changes you have made.

SAVE fn ft            To stay in the editor but save all the changes you have made so far.

GET fn ft            Inserts the named file into the file you are editing. It will be inserted just below the active line.

2.3.9 Some commands for running APL on the IBM VM/CMS system  
(most of these commands also work on the IBM PC with APL)

In APL look at the red letters and symbols. Look at the black ones only where there is no red symbol.

2.3.9.1

)LOAD ws            Loads the named workspace from disk to your active area. This workspace will contain programs (called functions in APL) and variables (which can represent vectors or matrices of numbers in APL).

)FNS                This causes the functions in the active workspace to be listed.

)VARS              This causes the variables in the active workspace to be listed.

)WSID              This obtains the name of the active workspace (in case you forget it).

)SAVE              This saves the active workspace under the same name. Be careful! You are overwriting the workspace that you loaded from disk!

)SAVE ws            This saves the active workspace under the name ws. If you give a new name

for ws, then you create a new workspace and do not overwrite the old one.

)ERASE n1 n2 n3 ---where 'n1 n2 n3 ---' are names of functions and/or variables. This erases the named functions and/or variables from the active workspace. (Before doing a 'SAVE' you should do 'FNS' and 'VARS' to see what garbage you have accumulated, and then do an 'ERASE' to get rid of it.)

vn where vn is a variable name. The contents of the variable will be displayed. For example if the variable X contains the vector '1 2 3 4', and you enter 'X', then '1 2 3 4' will be displayed.

vn ← ---- This sets the variable named vn equal to whatever is to the right of the arrow. For example if you enter 'Y ← 4 6 7' and then enter 'Y', the response will be '2 4 6 7'. If X contains '1 2 3 4', and you now enter 'Z ← X, Y' then Z will contain '1 2 3 4 2 4 6 7'.

vn3 ← vn1, vn2 Therefore, as just described, two vectors are put together. That is, they are "catenated".

vn3 vn1+vn2 This adds the two vectors, element by element. Obviously they must be the same length. For example if you entered 'Z ← X+Y', then Z would contain '3 6 9 11'. Subtraction ('-') works the same way, and so does

multiplication ('x') and division ('/'). N.B.: The minus sign for subtraction is at the upper right of the keyboard, whereas the "negative" sign to put in front of a negative number is at the upper left of the keyboard. They are different symbols in APL.

vn [3 4]

This will display the 3rd and the 4th elements of a vector whose name is vn. Obviously vn must be a vector, and it must have at least 4 elements. For example if you enter 'Y [3 4]'; the response will be '6 7' if Y contains '2 4 6 7'.

vn 4 2 / ----

This creates a 4-by-2 data matrix from the vector of numbers represented by '----'. There must be  $4 \times 2 = 8$  elements in the vector. For example if you enter 'D ← 4 2 / 1 2 2 4 3 6 4 7', then D will contain the matrix '1 2'.

```

      2 4
      3 6
      4 7

```

vn3, vn1, vn2

where vn1 and vn2 contain matrices rather than vectors. This concatenates the matrices. For example if you enter 'D ← X, Y' where X has been defined by 'X ← 4 1 / 1 2 3 4', and Y by 'Y ← 4 1 / 2 4 6 7', then D will contain the matrix '1 2'.

```

      2 4
      3 6
      4 7

```

`D[3;2]`

This would result in the display of the '6' which is the element in the 3rd row and the 2nd column of D as defined above.

`D[;2]`

This would result in the display of the 2nd column of matrix D. If you entered '`D[;1 2]`' or '`D[1 2 3 4;]`', then you would get a display of all of matrix D. (That would be silly of course you could just enter 'D' and get the same thing. But if you enter '`D[3 4;]`', then you will get a display of '3 6'.)

4 7

`Y←XP`

where '`Y←-4 1 2 4 6 7`', and '`XP←-1,X`' where '`X←-4 1 1 2 3 4`'. The result is a display of 0.5 and 1.7, the intercept and slope of the regression of Y on X.

`)LIB`

This displays the names of all your APL workspaces stored in your disk area.

`)OFF HOLD`

This causes you to leave APL mode, but stay logged on the system. (Be sure to SAVE your active workspace first if you have created something you want to keep!)

2.3.9.2 Examples of APL

ADDITION

8 7 + 7 3

15 10

SUBTRACT

4 6 - 2 3

2 3

RECIPROCAL

÷ 5 2

0.2 0.5

ABSOLUTE VALUE

3 -6 -5

3 6 5

NATURAL LOGARITHM

1 10

0 2.303

MULTIPLY

2 6 x 1 4

2 24

DIVIDE

2 6 - 1 4

2 1.5

SHAPE

2 2 1 2 3 4  
/

1 2

3 4

TRANSPOSE

If A is  $\begin{matrix} 1 & 2 \\ 3 & 4 \end{matrix}$  then A

1 3

2 4

CATENATE

1 3 , 4 5

1 3 4 5



## EXPONENT

\* 0 2.303

1 10

## FACTORIAL

!1 2 3 4

2 6 24

## MATRIX INVERSE

if B is  $\begin{bmatrix} 2 & 1 \\ 1 & 3 \end{bmatrix}$  then  $\begin{bmatrix} \cdot & \cdot \\ \cdot & \cdot \end{bmatrix} B$ 

$$\begin{bmatrix} 0.6 & -0.2 \\ -0.2 & 0.4 \end{bmatrix}$$

## INDEXING

 $[A \ 2; ]$ 

3 4

 $B[2; 2]$ 

3

### 3. INTRODUCTION TO LINEAR REGRESSION

#### 3.1. General remarks

Linear regression analysis is treated extensively in standard textbooks on introductory statistics and biometrics, some of which are cited in the bibliography. This section is not intended to substitute for a knowledge of regression analysis that should have been obtained from an introductory statistics course using such a textbook. Indeed, such a background was explicitly stated as a prerequisite for this course. This section is intended to be a review, and also to introduce participants to some new concepts and techniques. New concepts include matrix notation, matrix algebra, and derivation of the least squares regression formulae from first principles. New techniques include the use for regression analysis, of the software (BASIC, FORTRAN and APL languages; MINITAB and SAS statistical packages) and hardware (APPLE DOS 3.3, APPLE CP/M, IBM PC, and IBM 3081 mainframe) which will be used throughout the course.

#### 3.2 A linear regression analysis on a small data set

X :     -2     0     +2                    and n=3  
 Y :     +3    +1    -4

(1) First we need the X and Y deviations, but in these data  $\bar{X} = 0$  and  $\bar{Y} = 0$  so the data are X and Y deviations.

$$\begin{aligned} \text{Therefore} \quad \Sigma Y^2 &= \Sigma y^2 = 3^2 + 1^2 + (-4)^2 &= 26 \\ \Sigma X^2 &= \Sigma x^2 = (-2)^2 + 0^2 + 2^2 &= 8 \\ \Sigma XY &= \Sigma xy = (-2)(3) + (0)(1) + (2)(-4) &= -14 \end{aligned}$$

(2) slope  $b = \Sigma xy / \Sigma x^2 = -14/8 = -1.75$

$$\bar{y} = 0$$

$$\bar{x} = 0$$

Therefore  $\bar{Y} = a + b\bar{X}$   
 $0 = a + (-1.75)(0)$  and  $a = 0$

The equation is  $y = -1.75X$

(3) ANOVA of regression table:

Source	df	SS	MS	F
Regression	1	$(\sum xy)^2 / \sum x^2$ $= (-14)^2 / 8$	24.5	$24.5 / 1.5$ $= 16.3$
Error	$n-2$ $= 3-2$ $= 1$	$26 - 24.5 = 1.5$	1.5	
Total	$n-1$ $= 3-1$ $= 2$	$\sum y^2 = 26$		

$$r^2 = \frac{\text{regr. SS}}{\text{tot. SS}} = \frac{(\sum xy)^2 / \sum x^2}{\sum y^2} = \frac{(\sum xy)^2}{\sum x^2 \sum y^2} = \frac{24.5}{26} = 0.942$$

$$r = \sqrt{r^2} = \sqrt{0.942} = 0.971$$

### 3.3 Examples of Matrix algebra.

Addition: 
$$\begin{bmatrix} 1 & 5 \\ 5 & 4 \end{bmatrix} + \begin{bmatrix} 3 & 2 \\ 2 & 4 \end{bmatrix} = \begin{bmatrix} 4 & 7 \\ 7 & 8 \end{bmatrix}$$

Subtraction: 
$$\begin{bmatrix} 1 & 5 \\ 5 & 4 \end{bmatrix} - \begin{bmatrix} 3 & 2 \\ 2 & 4 \end{bmatrix} = \begin{bmatrix} -2 & 3 \\ 3 & 0 \end{bmatrix}$$

$$\text{Transpose : } \begin{bmatrix} 6 & 2 \\ 1 & 3 \end{bmatrix}^{-1} = \begin{bmatrix} 6 & 1 \\ 2 & 3 \end{bmatrix} \quad (\text{rows} \rightarrow \text{columns})$$

$$\text{Multiply : } \begin{bmatrix} 2 & 3 & 1 \\ 1 & 5 & 2 \end{bmatrix} \begin{bmatrix} 1 & 8 \\ 4 & 1 \\ 3 & 6 \end{bmatrix} = \begin{bmatrix} 17 & 25 \\ 27 & 25 \end{bmatrix} \quad (\text{rows by columns})$$

$$\text{Inverse : } \begin{bmatrix} 2 & 1 \\ 1 & 3 \end{bmatrix}^{-1} = \begin{bmatrix} 0.6 & -0.2 \\ -0.2 & 0.4 \end{bmatrix}$$

$$\text{Check: } \begin{bmatrix} 2 & 1 \\ 1 & 3 \end{bmatrix} \begin{bmatrix} .6 & -.2 \\ -.2 & .4 \end{bmatrix} = \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix}$$

$$\text{"Divide": } \begin{bmatrix} 2 & 1 \\ 1 & 3 \end{bmatrix}^{-1} \begin{bmatrix} 5 & 2 \\ 2 & 6 \end{bmatrix} = \begin{bmatrix} 2.6 & 0 \\ -0.2 & 2.0 \end{bmatrix}$$

(Note that the result is not a symmetric matrix, though the originals were)

$$\text{Determinant: } \begin{vmatrix} 2 & 1 \\ 1 & 3 \end{vmatrix} = (2)(3) - (1)(1) = 5$$

Roots and vectors:  $\begin{bmatrix} 2 & 6 \\ .5 & 0 \end{bmatrix} = M$  Find roots and vectors of matrix  $M$ .

Solve the equation  $(\lambda I - M)X = 0$ ,

$$\text{which is } \left| \lambda \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix} - \begin{bmatrix} 2 & 6 \\ .5 & 0 \end{bmatrix} \right| \begin{bmatrix} X_1 \\ X_2 \end{bmatrix} = \begin{bmatrix} 0 \\ 0 \end{bmatrix} \quad (1)$$

$$\text{or } \left| \begin{bmatrix} \lambda & 0 \\ 0 & \lambda \end{bmatrix} - \begin{bmatrix} 2 & 6 \\ .5 & 0 \end{bmatrix} \right| \begin{bmatrix} X_1 \\ X_2 \end{bmatrix} = \begin{bmatrix} 0 \\ 0 \end{bmatrix} \quad (2)$$

$$\text{or } \left| \begin{bmatrix} \lambda-2 & -6 \\ -.5 & \lambda \end{bmatrix} \right| \begin{bmatrix} X_1 \\ X_2 \end{bmatrix} = \begin{bmatrix} 0 \\ 0 \end{bmatrix} \quad (3)$$

Divide both

$$\text{sides by } \begin{bmatrix} X_1 \\ X_2 \end{bmatrix} : \begin{bmatrix} \lambda-2 & -6 \\ -.5 & \lambda \end{bmatrix} = \begin{bmatrix} 0 \\ 0 \end{bmatrix} = \begin{matrix} (\lambda-2)\lambda - (-6)(-.5) = 0 \\ \lambda^2 - 2\lambda - 3 = 0 \\ (\lambda-3)(\lambda+1) = 0 \end{matrix}$$

So the roots are :  $\lambda_1 = 3$  and  $\lambda_2 = -1$

Note that  $\begin{bmatrix} 3 & 0 \\ 0 & -1 \end{bmatrix} = \begin{bmatrix} \lambda_1 & \\ & \lambda_2 \end{bmatrix}$  and  $M = \begin{bmatrix} 2 & 6 \\ .5 & 0 \end{bmatrix}$  are equivalent

For one thing, note that the sum of the diagonal is the same.

vector associated with each root can be found by substituting the root into the equation  $[\lambda I - M][X] = [0]$  (see equation (3) above). The vector for  $\lambda_1$  is  $\begin{bmatrix} X_1 \\ X_2 \end{bmatrix} = \begin{bmatrix} 6 \\ 1 \end{bmatrix}$  or any vector proportional to it.

To do matrix algebra using MINITAB (where M1 and M2 are matrices):

```

Addition      : 'ADD M1 TO M2, PUT IN M3'
Subtraction   : 'SUBTRACT M2 FROM M1, PUT IN M3'
Transpose     : 'TRANPOSE M1, PUT IN M2'
Multiply      : 'MULTIPLY M1 BY M2, PUT IN M3'
Inverse       : 'INVERT M1, PUT IN M2'
"Divide"      : 'INVERT M1, PUT IN M2' and
                'MULT M1 BY M2, PUT IN M3'
Determinant   : 'EIGEN M1, C1, M2' & 'LET C2 = LOG E(C1)'
                and 'LET K1 = EXPO(SUM(C2))'
Roots & vectors: 'EIGEN M1, PUT ROOTS IN C1,
                VECTORS IN M2'

```

(N.B.): This will only work for a symmetric matrix. There is another way of doing it for a nonsymmetric matrix.

To do matrix algebra using APL (where vn1 and vn2 are matrices):

```

Addition      : 'vn3 ← vn1 + vn2'
Substraction  : 'vn3 ← vn1 - vn2'
Transpose     : 'vn2 ← ⍉vn1'
Multiply      : 'vn3 ← vn1 +.x vn2'
Inverse       : 'vn2 ← ⍉ vn1'
"Divide"      : 'vn3 ← vn2 ⍉ vn1'
Determinant   : Use function 'PDET vn'
Roots & vectors: Use function 'GEIG vn'. Works for a
                symmetric or nonsymmetric
                matrix so long as the roots are
                fairly distinct (that is, none of the
                roots are approximately equal to each
                other).

```

### 3.4 Derivation of least squares regression formula

In matrix notation,  $Y = XB + e$ , where the  $e_i$  are independent estimates of  $\epsilon_i$ , which (for t & F-tests) are normally distributed with 0 mean.

which is

$$\begin{bmatrix} Y_{i=1} \\ Y_{i=2} \\ \vdots \\ Y_{i=n} \end{bmatrix} = \begin{bmatrix} 1 & X_{i=1} \\ 1 & X_{i=2} \\ \vdots & \vdots \\ 1 & X_{i=n} \end{bmatrix} \begin{bmatrix} b_0 \\ b_1 \end{bmatrix} + \begin{bmatrix} e_{i=1} \\ e_{i=2} \\ \vdots \\ e_{i=n} \end{bmatrix}$$

where  $b_0$  = intercept and  
 $b_1$  = slope.

Finally,

$$\begin{bmatrix} Y_{i=1} \\ Y_{i=2} \\ \vdots \\ Y_{i=n} \end{bmatrix} = \begin{bmatrix} b_0 + b_1 X_{i=1} \\ b_0 + b_1 X_{i=2} \\ \vdots \\ b_0 + b_1 X_{i=n} \end{bmatrix} + \begin{bmatrix} e_{i=1} \\ e_{i=2} \\ \vdots \\ e_{i=n} \end{bmatrix}$$

$$= \begin{bmatrix} b_0 + b_1 X_{i=1} + e_{i=1} \\ b_0 + b_1 X_{i=2} + e_{i=2} \\ \vdots \\ b_0 + b_1 X_{i=n} + e_{i=n} \end{bmatrix}$$

We want to find B such that  $e'e$  is a minimum, where

$$e'e = \begin{bmatrix} e_{i=1} & e_{i=2} & \dots & e_{i=n} \end{bmatrix} \begin{bmatrix} e_{i=1} \\ e_{i=2} \\ \vdots \\ e_{i=n} \end{bmatrix} = \sum_i e_i^2$$

This is the "least squares solution".

$$\begin{aligned}
 \text{If } e &= Y - XB, \text{ then} \\
 e'e &= (Y - XB)'(Y - XB) \\
 &= Y'Y - (XB)'Y - Y'XB + (XB)'XB \\
 &= Y'Y - 2X'YB + X'XB'B, \\
 &\quad \text{because } X'Y = Y'X
 \end{aligned}$$

To minimize something, we differentiate it with respect to the parameter(s) we are trying to estimate and set it equal to zero,

$$\text{so } \frac{\partial e'e}{\partial B} = 0 = -2X'Y + 2X'XB,$$

$$\begin{aligned}
 \text{and } X'XB &= X'Y, \\
 \text{and } \hat{B} &= (X'X)^{-1} X'Y.
 \end{aligned}$$

In subscripted, rather than matrix, notation

$$X'XB = X'Y \quad \text{is}$$

$$\begin{bmatrix} 1 & 1 & \dots & 1 \\ X_{i=1} & X_{i=2} & \dots & X_{i=n} \end{bmatrix} \begin{bmatrix} 1 & X_{i=1} \\ 1 & X_{i=2} \\ \vdots & \vdots \\ 1 & X_{i=n} \end{bmatrix} \begin{bmatrix} b_0 \\ b_1 \end{bmatrix} = \begin{bmatrix} 1 & 1 & \dots & 1 \\ X_{i=1} & X_{i=2} & \dots & X_{i=n} \end{bmatrix} \begin{bmatrix} Y_{i=1} \\ Y_{i=2} \\ \vdots \\ Y_{i=n} \end{bmatrix}$$

$$\begin{bmatrix} n & \sum X_{i i} \\ \sum X_{i i} & \sum X_{i i}^2 \end{bmatrix} \begin{bmatrix} b_0 \\ b_1 \end{bmatrix} = \begin{bmatrix} \sum Y_{i i} \\ \sum X_{i i} Y_{i i} \end{bmatrix}$$

$$b_0 n + b_1 \sum X_{i i} = \sum Y_{i i}$$

$$b_0 \sum X_{i i} + b_1 \sum X_{i i}^2 = \sum X_{i i} Y_{i i}$$



$$-b_0 \sum_i X_i - b_1 \sum_i X_i (\sum_i X_i / n) = - \sum_i Y_i (\sum_i X_i / n)$$

$$b_0 \sum_i X_i + b_1 \sum_i X_i^2 = \sum_i X_i Y_i$$

$$b_1 \left( \sum_i X_i^2 - \frac{(\sum_i X_i)^2}{n} \right) = \sum_i X_i Y_i - \frac{\sum_i X_i \sum_i Y_i}{n}$$

$$\hat{b}_1 = \frac{\sum xy}{\sum x^2}$$

Then

$$b_0 = \left( \sum_i Y_i - b_1 \sum_i X_i \right) / n$$

$$= \frac{\sum_i Y_i}{n} - b_1 \frac{\sum_i X_i}{n} \quad \text{and}$$

$$\hat{b}_0 = \bar{Y} - b_1 \bar{X}$$

### 3.5 Practical session 1

#### 3.5.1 Assignment/Tutorial

##### Purpose:

- (1) To provide practice in using the APPLES, the IBM PC, and the IBM mainframe.
- (2) To provide practice in running BASIC, FORTRAN and APL programs, and in using the MINITAB statistical package.
- (3) To provide some review in simple regression analysis.

##### Approach:

Given a small set of observations on variables X and Y:

X: 1 2 3 4  
Y: 2 4 6 7

use the above-mentioned hardware/software to

- (1) plot Y versus X
- (2) calculate the least squares regression of Y on X.

Procedure:

(1) On the APPLE, use DOS 3.3 to run the programs REGRESSION and PLOT. These programs require that the X,Y data are in a sequential text file, which can be created using the program CREATE TEXT. All 3 of these BASIC programs are from Orloci and Kenkel (1984).

(2) On the APPLE, use CP/M to run the program LINREG. You can choose the option of entering data from the screen, or the option of reading data from a file. If you choose the 2nd option, then you first have to run program CREATE, which will create a data file called 'DATA'.

(3) On the computer terminal to the IBM mainframe (including IBM PCs used as terminals), run the program LINREG. When you are in BASIC mode, enter 'LOAD LINREG'. Then enter 'RUN', and select the option to enter data from the keyboard. Enter the data as '1,2', '2,4', '3,6', and '4,7'. The t-value for 2 error df is 4.3. When the program offers more statistics, respond with a 'Y'. (N.B.: All letter responses must be capital letters.) After you have run the program, you can enter 'LIST' and if you have some knowledge of BASIC you can examine the program to see how it is designed.

(4) On the computer terminal to the IBM mainframe, run the FORTRAN programs PLOT and REGR. But first, after you have logged on and done your 'DEF STOR 1500K', enter 'TYPE REGR DATA'. Enter 'COPY REGR DATA A PLOT DATA A' to produce an identical data set for the PLOT program. Note the last line which indicates "end of file". Now enter 'TYPE PLOT FORTRAN'. It is rather long and tedious - a utility program after all - but you should look at the first screenfull to see how the data are read in from the file 'PLOT DATA'. Then enter 'TYPE REGR FORTRAN', and do the same. Now run the program 'PLOT FORTRAN' by

entering 'FORTVS PLOT'. (The name 'FORTVS' calls an 'EXEC' file which has been set up to do the compiling, loading, running, and to identify the correct input and output files. If you are interested, you can look at this 'EXEC' file by entering 'TYPE FORTVS EXEC'.) The information that appears on your screen is diagnostics of the progress of what 'FORTVS EXEC' is doing. When you see ';R-----', then the program has run. Your output is in file 'PLOT OUTPUT'. To see it, enter 'TYPE PLOT OUTPUT'. Some parameters of the program run are shown, and then the plot, which may be split between screenfuls. To cure this, go into the editor (type 'XEDIT PLOT OUTPUT') and then delete all the lines except the plot itself. This can be done by putting 'DD' on the dashed "prefix" lines: put 'DD' on the first line of the file and put 'DD' on the last line before the plot itself. Then depress ENTER and all the lines before the plot will disappear. Enter 'FILE', and then clear the screen and enter 'TYPE PLOT OUTPUT' again.

Now run 'REGR FORTRAN' in the same manner. When it has run, enter 'TYPE REGR OUTPUT'.

If you want hard copy of your data files and/or your output files, they can be printed out using 'PRINT' or 'LPRINT'.

In running FORTRAN programs you automatically create files. The compiler creates a machine language file 'fn TEXT', and the run creates a file 'fn LISTING' with run-time diagnostics in it (which you should look at if the program didn't work). And of course there is 'fn OUTPUT' created as well. Before logging off you should always erase such files if you have no use for them. If you don't, then in a few days you will have your disk area filled with "junk files".

(5) On the IBM PC, run the BASIC program LINREG. (This is the same BASIC program you ran on the IBM mainframe.)

The program is on the "MS DOS program" disk.

Again, choose the option to enter data from the keyboard.

The FORTRAN programs PLOT and REGR can also be run on the IBM PC.

(6) On the computer terminal to the IBM mainframe, run the APL

"programs" (called functions in APL) SCATTERPLOT and GLM.  
Once in APL mode, proceed as follows.

)LOAD UNESCO	This loads the workspace.
)FNS	This causes the functions in this workspace to be listed.
)VARS	This causes the variables in this workspace to be listed. (There aren't any initially.)
X← 1 2 3 4	This creates a variable X which contains the vector of numbers '1 2 3 4'.
X	This displays the contents of X on the screen.
Y← 2 4 6 7	These two commands do the same for variable Y.
Y	

(now clear the screen)

X SCATTERPLOT Y	This "runs" the SCATTERPLOT function.
34 20	is your response (as suggested).
N	is your response - you do not want to force plot axes to have the same scale. Follow instructions.
X← 4 1 X	These two commands change variable X from a vector to a 4-by-1 matrix, and then display it on the screen.
X	
Y← 4 1 Y	Same for variable Y.
Y	
D← X, '2' Y	These two commands catenate X and

D                                    Y into a new variable D, and  
    display it on the screen.

(now clear the screen)

1 2 GLM D                            This "runs" the GLM function. The  
    '1 2' says that the independent (X)  
    variable is in column 1 and the  
    dependent (Y) variable is in column 2.  
    The variable D contains the data. You  
    can have more than 1 X-variable  
    (multiple regression). The first  
    column at the bottom left is the  
    predicted Y values. The second is  
    the Y residuals.

)OFF HOLD                            Leaves APL mode but keeps you  
    logged on to the system.

You can also do this exercise on the IBM PC which has APL  
 implemented on it.

(7) On the IBM mainframe, do the same plot and regression  
 analysis in MINITAB, as follows.

READ INTO C1-C2  
 1 2  
 2 4                                    Reads the data into C1 and C2 and  
 3 6                                    then prints the contents of C1 and C2.  
 4 7  
 PRINT C1-C2

WIDTH 55, HEIGHT 16                Changes plot dimensions to fit the  
 (clear the screen)                    screen, then after clearing the  
 PLOT C2 VS C1                        screen, plots Y versus X.

PLOT C2 FROM 0 TO 10                The same plot, but with you (rather

VS C1 FROM 0 TO 5	than MINITAB) controlling the scales on the axes.
REGRESS C2 ON 1 PRED. IN C1	Does the regression of Y on X.
BRIEF 1	Limits the regression output, repeats
REGR C2 1 C1, C3, C4	the regression storing Y residuals in C3
PRINT C3-C4	and Y predicted in C4,
PLOT C4 VS C1	prints the contents of C3 and C4
PLOT C3 VS C4	plots the fitted line, and
	plots Y residuals versus Y predicted.
STOP	Leaves MINITAB but stays on the system.

### 3.6 PRACTICAL SESSION 2

#### 3.6.1 Assignment/Tutorial

##### 3.6.1.1 MINITAB

- A. READ in the data set from file 'REGRI DATA' into C1 and C2. There are 100 observations by 2 variables (Y and X respectively).
- B. Do a Model I regression analysis by matrix algebra:
1. PLOT Y versus X to check that a linear regression model looks sensible.
  2. SET 100 values of 1 into C3, and then COPY C3 and C2 into M1. The X matrix is now in M1.
  3. TRANSPOSE M1 and put it into M2. The X' matrix is now in M2.
  4. MULTIPLY M2 by M1 and put the product X'X into M3.
  5. COPY C1 into M4. The Y matrix is now in M4.
  6. MULTIPLY M2 by M4 and put the product X'Y into M5.
  7. INVERT M3 and put X'X inverse into M6.

8. MULTIPLY M6 by M5 and put the product,  $(X'X \text{ inverse}) \cdot (X'Y)$ , into M7. The matrix of regression coefficients, B, is now in M7.
9. MULTIPLY M1 by M7 and put the product XB into M8. The predicted Y values (Y-hat) are now in M8.
10. COPY M8 into C4, and then LET  $K1 = \text{SUM}((C1 - C4) \cdot (C1 - C4))$ . The Error SS is now in K1.
11. LET  $K2 = \text{SUM}((C1 - \text{AVER}(C1)) \cdot (C1 - \text{AVER}(C1)))$ . The total SS is now in K2.
12. LET  $K3 = K2 - K1$  puts the Regression SS into K3. LET  $K4 = K3 / K2$  puts r-squared into K4. LET  $K5 = K1 / 98$  puts the Error MS into K5. LET  $K6 = K3 / K5$  puts F into K6. You have your ANOVA table.

C. Now do the same analysis using the REGR command:

1. Do BRIEF 1. Then do REGR C1 1 C2, and then BRIEF 6 followed by REGR C1 1 C2. Compare these outputs with each other and with the results from the matrix algebra solution.
2. Do BRIEF 1 and then REGR C1 1 C2, C5, C6. This time you have put the standardized residuals,  $(Y - \hat{Y}) / \text{SQRT}(\text{Error MS})$ , and the predicted Y values, Y-hat, into C5 and C6 respectively. Compare them with your BRIEF 6 output. (To convert the standardized residuals back to the "raw" residuals you would just multiply C5 by  $\text{SQRT}(\text{Error MS}) = \text{SQRT}(K5)$ .)
3. Do HISTOGRAM of C5 to see if the residual errors  $e = Y - \hat{Y}$  appear to be normally distributed. Another way to do it is by an arithmetic probability plot, by doing NSCORES C5, C7 and then PLOT C5 C7. If the residual errors are approximately normal, you should see a fairly straight line.
4. Now see whether the residual errors are independent of the other effects in the model, as they should be. PLOT C5 versus C6. You should see a normally distributed scatter centred on  $e = C5 = 0$ . There should

be no pattern, or relationships, apparent in the plot.

5. Do RLINE C1,C2 which produces estimates of slope and intercept by an iterative procedure quite different from the least-squares estimate and more robust to outliers. Compare these estimates with the least-squares estimates.

### 3.6.1.2 SAS

We will run SAS in batch mode (although it can be run in interactive mode). First, we must create a file containing the SAS job commands. Start by entering 'XEDIT RUNSAS SAS', and then go into INPUT mode within the editor. Enter the following lines:

DATA REGSAS;	Names the SAS data set to be created.
INPUT Y X;	Names the variables and their input order.
CARDS;	Says the data follow, on "card images".
PROC PRINT;	Causes the data just read in to be printed.
PROC PLOT; PLOT Y*X;	Produces a plot of Y versus X.
PROC GLM; MODEL Y=X;	Produces a regression analysis of Y on X.

Note that all SAS statements end with a ';'. You can have several statements to a line, or a statement can continue over several lines, as long as you remember to end each statement with a semi-colon.

Now, what about the data? A SAS job can read data from a data file, but it is easier to just "pull" the data into the SAS job file we have just created. Get out of INPUT mode. Move the "active line" (the brightly lit up line) up or down until the 'CARDS;' line is the active line. (Use the commands 'UP' or 'DOWN' to move the active line - for example if the bottom line is the active line then 'UP 3' should do it.) Now enter 'GET REGRI DATA', and all the data should be inserted into the



file after the 'CARDS;' line. Now enter 'FILE' to leave the editor, and then run your SAS job by entering 'SAS RUNSAS'. When you get the response ';R---', the job has run. Enter 'TYPE RUNSAS SASLOG' to see a record of the run. To see the output, enter 'TYPE RUNSAS LISTING'. Notice that the output is intended for printing on 132-character-wide paper, not for display on an 80-character-wide screen. To print out hard copy on the local printer, enter 'LPRINT B08 fn ft'. I would suggest that you print out both 'RUNSAS SAS' (the job command file) and 'RUNSAS LISTING' (the output file).

### 3.6.2 Job Listings and Outputs.

FILE: REGR    MINITAB A    VM/SP - CONVERSATIONAL MONITOR SYSTEM

PAGE 001

```
READ INTO C1-C2
1 2
2 4
3 6
4 7
PRINT C1-C2
PLOT C2 VS C1
PLOT C2 FROM 0 TO 10 VS C1 FROM 0 TO 5
REGRESS C2 ON 1 PRED, IN C1
BRIEF 1
REGR C2 1 C1, C3, C4
PRINT C3-C4
PLOT C4 VS C1
PLOT C3 VS C4
STOP
```

MINITAB Job Listing for regression run on small data set.

MINITAB output from regression run on small data set.  
FILE: REGR OUTPUT A VM/SP - CONVERSATIONAL MONITOR SYSTEM

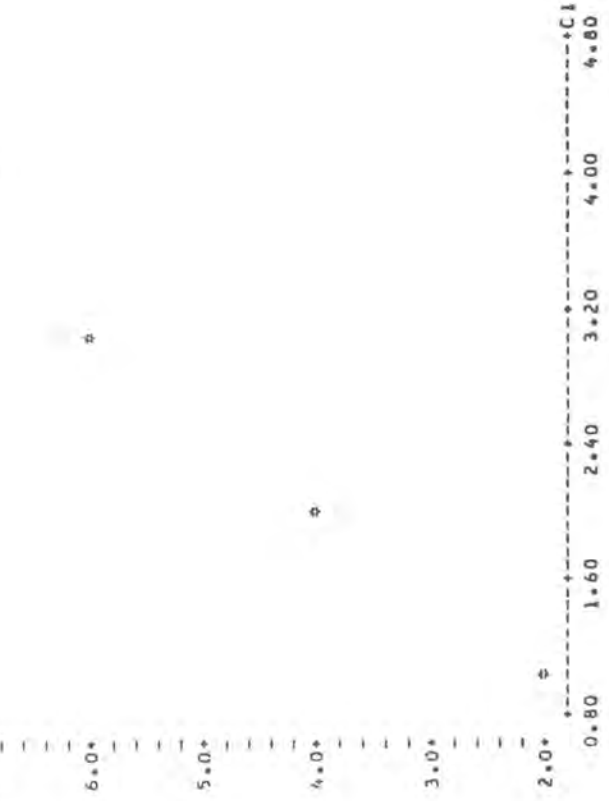
MINITAB RELEASE 81.1 \*\*\* COPYRIGHT - PENN STATE UNIV. 1981  
MAY 22 1985 \*\*\* NATIONAL UNIVERSITY OF SINGAPORE - LOCAL VERSION 02/12/1982  
STORAGE AVAILABLE 4800

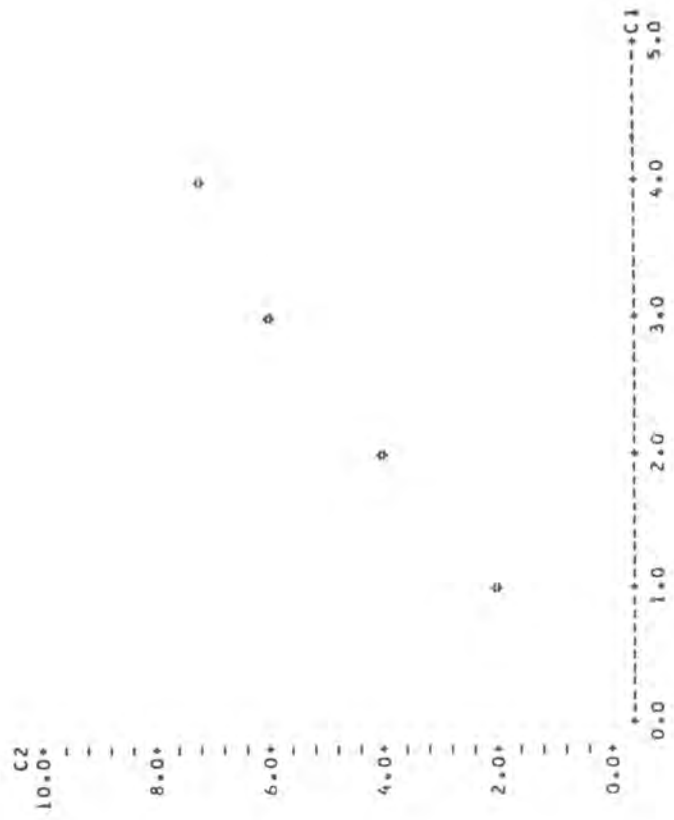
---  
---  
---  
---

COLUMN COUNT ROW	C1	C2
1	4	4
2	1*	2*
3	2*	4*
4	3*	6*
	4*	7*

---

C2
7.0*
6.0*
5.0*
4.0*
3.0*
2.0*





THE REGRESSION EQUATION IS  
 $Y = 0.500 + 1.70 X1$

COLUMN	COEFFICIENT	ST. DEV. OF COEF.	T-RATIO = COEF/S.D.
X1	1.7000	0.5000	1.05
C1	0.1732	0.1732	9.81

THE ST. DEV. OF Y ABOUT REGRESSION LINE IS  
 $S = 0.3873$   
 WITH (4 - 2) = 2 DEGREES OF FREEDOM

R-SQUARED = 98.0 PERCENT  
 R-SQUARED = 96.7 PERCENT, ADJUSTED FOR D.F.

ANALYSIS OF VARIANCE

SOURCE	DF	SS	MS=SS/DF
DUE TO			

REGRESSION 1 14.4500  
 RESIDUAL 2 0.3000  
 TOTAL 3 14.7500

DURBIN-WATSON STATISTIC = 2.23

ST. DEV. OF COEF. T-RATIO =  
 0.4743 1.05  
 0.1732 9.61

S = 0.3873

R-SQUARED = 98.0 PERCENT

ROW	COLUMN	C3	C4
1	-0.94281	2.20000	
2	0.30861	3.90000	
3	1.23443	5.60000	
4	-1.41422	7.30000	

C4  
 7.6+

6.2+

4.8+

3.4+

2.0+

-----C1  
 0.80 1.60 2.40 3.20 4.00 4.80

LINREG on small data set on PC.

Ok  
 run  
 ? 4  
 HOW MANY X-Y PAIRS DO YOU WANT TO ENTER?  
 ? 0  
 IF YOU WANT NO TRANSFORMATION OF X INPUT 0  
 IF YOU WANT A LOG(X) TRANSFORMATION INPUT 1  
 IF YOU WANT A LOG(X+1) TRANSFORMATION INPUT 2  
 ? 0  
 IF YOU WANT NO TRANSFORMATION OF Y INPUT 0  
 IF YOU WANT A LOG(Y) TRANSFORMATION INPUT 1  
 IF YOU WANT A LOG(Y+1) TRANSFORMATION INPUT 2  
 ? 0

ENTER THE DATA AS X-Y PAIRS.

? 1,2  
 ? 2,4  
 ? 3,6  
 ? 4,7

WHAT IS THE T-VALUE FOR THE 95% CONFIDENCE LIMITS  
 WITH 2 DEGREES OF FREEDOM?  
 ? 4.303

WHAT IS THE T-VALUE FOR THE 95% CONFIDENCE LIMITS  
 WITH 2 DEGREES OF FREEDOM?  
 ? 4.303

THE REGRESSION STATISTICS ARE AS FOLLOWS:

THE EQUATION OF THE LINE IS:  $Y = .5 + 1.7 X$

WHERE THE SLOPE IS: 1.7  
 AND THE Y-INTERCEPT IS: .5  
 THE STANDARD ERROR OF THE REGRESSION IS: (+ OR -) 5.960285  
 THE STANDARD ERROR OF THE SLOPE IS: (+ OR -) 2.665521  
 THE 95% C.L. FOR THE SLOPE ARE: -9.769736 13.16974  
 THE STANDARD ERROR OF THE INTERCEPT IS: (+ OR -) 7.299829  
 THE 95% C.L. FOR THE INTERCEPT ARE: -30.91116 31.91116  
 THE CORRELATION COEFFICIENT (R) IS: .411032  
 THE COEFFICIENT OF DETERMINATION (R\*\*2) IS: .1690058

DO YOU WANT MORE STATISTICS PRINTED?  
 TYPE Y OR N.  
 ? Y

THE REGRESSION COMPUTATIONS HAVE PRODUCED THE FOLLOWING:

THE MEANS OF X AND Y ARE: 2.5 4.75  
 THE SUM OF X IS: 10  
 THE SUM OF Y IS: 19  
 THE SUM OF X-SQUARED IS: 30  
 THE SUM OF Y-SQUARED IS: 105  
 THE SUM OF X\*Y IS: 56  
 THE SUM OF SQUARES OF X IS: 5  
 THE SUM OF CROSS-PRODUCTS IS: 8.5  
 THE REGRESSION SUM OF SQUARES IS: 14.45  
 THE RESIDUAL SUM OF SQUARES IS: 71.05  
 THE TOTAL SUM OF SQUARES IS: 85.5  
 THE REGRESSION MEAN SQUARE IS: 14.45  
 THE RESIDUAL MEAN SQUARE IS: 35.525

THE F-VALUE IS: .4067558  
 WITH 1 REGRESSION D OF F, AND  
 A RESIDUAL D OF F OF: 2

DO YOU WANT 95% CONFIDENCE LIMITS?  
 TYPE Y OR N.  
 ? Y

THE RESIDUAL MEAN SQUARE IS: 35.525

THE F-VALUE IS: .4067558  
 WITH 1 REGRESSION D OF F, AND  
 A RESIDUAL D OF F OF: 2

DO YOU WANT 95% CONFIDENCE LIMITS?  
 TYPE Y OR N.  
 ? Y

DO YOU WANT TO SPECIFY THE X VALUES?  
 TYPE Y OR N.  
 ? N

THE PREDICTED VALUES AND 95% C.L. OF Y ARE:

GIVEN X VALUE	PREDICTED Y	LOWER Y	UPPER Y	ERROR
1 2.2	-19.25791	23.65791	4.986733	
2 3.9	-10.1475	17.9475	3.264583	
3 5.600001	-8.447499	19.6475	3.264583	
4 7.3	-14.15791	28.75791	4.986733	

Ok

Output of run of REGR, FORTRAN on small data set.  
FILE: REGR OUTPUT A VM/SP - CONVERSATIONAL MONITOR SYSTEM

THE DATA AS READ IN, BEFORE ANY TRANSFORMATION, ARE:

X	Y
1.000	2.000
2.000	4.000
3.000	6.000
4.000	7.000

THE DATA AFTER TRANSFORMATION, IF ANY, ARE:

X	Y
1.000	2.000
2.000	4.000
3.000	6.000
4.000	7.000

X MEAN= 2.50 Y MEAN= 4.75  
 X VARIANCE= 1.67 Y VARIANCE= 4.92 XY COVARIANCE = 2.83

THE REGRESSION LINE IS Y = 0.5000 + 1.7000X

THE ANALYSIS OF VARIANCE TABLE IS:

SOURCE	SUM OF SQUARES	MEAN SQUARE	F-STATISTIC
1	14.45	14.45	96.33
2	0.30	0.15	
3			

R-SQUARED= .97966 PERCENT R-SQUARED=97.97

Y-PREDICTEDS AND Y-RESIDUALS FOLLOW.

Y-PREDICTEDS	Y-RESIDUALS
2.200	0.200
3.900	-0.100
5.600	-0.400
7.300	0.300

FILE: LAB2A MINITAB A1 V4/SP - CONVERSATIONAL MONITOR SYSTEM

NOTE THIS IS A LINEAR REGRESSION ANALYSIS BY MINITAB

READ C1-C2

19.41	10.
23.15	10.
33.62	10.
26.72	10.
24.60	10.
23.50	10.
26.59	10.
29.83	10.
19.37	10.
20.31	10.
42.93	20.
40.16	20.
35.11	20.
41.14	20.
33.52	20.
31.93	20.
39.58	20.
32.69	20.
37.65	20.
31.34	20.
44.54	30.
52.15	30.
47.24	30.
46.35	30.
48.71	30.
43.38	30.
52.15	30.
46.46	30.
44.42	30.
37.19	30.
60.67	40.
61.03	40.
52.24	40.
55.46	40.
59.02	40.
42.93	40.
53.08	40.
53.92	40.
57.55	40.
61.93	40.
58.96	50.
76.33	50.
77.54	50.
63.40	50.
74.75	50.
74.16	50.
66.59	50.
68.73	50.
62.55	50.
53.33	50.
78.70	60.
77.97	60.
75.02	60.



FILE: LAB2A MINITAJ AL V4/SP - CONVERSATIONAL MONITOR SYSTEM

74.51 00.  
 91.95 50.  
 90.25 50.  
 77.24 00.  
 73.85 50.  
 77.75 00.  
 78.00 50.  
 66.50 70.  
 77.50 75.  
 78.95 70.  
 84.13 70.  
 81.24 70.  
 96.95 70.  
 85.62 70.  
 80.95 70.  
 81.47 70.  
 95.85 70.  
 104.56 80.  
 105.35 80.  
 101.84 80.  
 104.96 80.  
 103.10 80.  
 110.26 90.  
 101.77 80.  
 101.04 90.  
 100.47 80.  
 91.80 80.  
 117.65 90.  
 112.62 90.  
 115.45 90.  
 106.53 90.  
 119.42 90.  
 114.64 70.  
 125.96 80.  
 122.10 70.  
 110.82 90.  
 119.41 90.  
 131.05 100.  
 126.28 100.  
 137.22 100.  
 111.13 100.  
 129.10 100.  
 125.25 100.  
 120.95 100.  
 117.01 100.  
 128.53 100.  
 112.11 100.

PRINT C1-C2  
 PLOT C1 VS C2  
 SET C3  
 10011  
 COPY C3 AND C2 INTO M1  
 TRANPOSE M1 PUT T4 M2  
 MULTIPLY M2 BY M1, PUT PRODUCT IN M3  
 PRINT M3

```

COPY C1 INTO M4
NOTE M3 IS THE X'X MATRIX
MULTIPLY M2 BY M4, PUT PRODUCT IN M5
PRINT M5
NOTE M5 IS THE X'Y MATRIX
INVERT M3, PUT IN M6
MULTIPLY M6 BY M5, PUT PRODUCT IN M7
PRINT M7
NOTE M7 IS THE S MATRIX
MULTIPLY M1 BY M7, PUT PRODUCT IN M8
COPY M8 INTO C4
LET K1=SUM((C1-C4)*(C1-C4))
LET K2=SUM((C1-AVER(C1))*(C1-AVER(C1)))
LET K3=K2-K1
LET K4=K3/K2
LET K5=K1/99
LET K6=K3/K5
PRINT K2
NOTE K2 IS THE TOTAL SUM OF SQUARED DEVIATIONS
PRINT K1
NOTE K1 IS THE ERROR SUM OF SQUARED DEVIATIONS
PRINT K3
NOTE K3 IS THE REGRESSION SUM OF SQUARED DEVIATIONS
PRINT K4
NOTE K4 IS ALSO THE REGRESSION MEAN SQUARED DEVIATIONS
PRINT K5
NOTE K5 IS THE ERROR MEAN SQUARED DEVIATIONS
NOTE K6 IS THE F RATIO
PRINT K4
NOTE K4 IS THE COEFFICIENT OF DETERMINATION R-SQUARED
NOTE TO RUN LINEAR REGRESSION USING THE REGR FUNCTION
BRIEF 6
REGRESS C1 1 C2, C5, C6
HIST C5
SORT(R5), K7
MULTIPLY C5 BY K7, C8
PRINT C3
NSCOR=C5, C7
PLOT C5 VS C6
PLOT C5 VS C7
RLINE C1,C2
STOP

```

PAGE 001

FILES LA02A    OUTPUT    A1   V4/50 - CONVERSATIONAL MONITOR SYSTEM

MINITAB RELEASE 11.1    COPYRIGHT - PENN STATE UNIV, 1931  
 APRIL 29, 1975    NATIONAL UNIVERSITY OF SINGAPORE - LOCAL VERSION 32/12/1972  
 EXAMPLE OF A LINEAR REGRESSION ANALYSIS USING MINITAB

COLUMN	C1	C2
COUNT	100	100
ROW		
1	19.410	10.
2	23.150	10.
3	33.520	10.
4	25.720	10.
5	24.500	10.
6	33.500	10.
7	25.590	10.
8	29.330	10.
9	17.370	10.
10	20.310	10.
11	42.930	20.
12	40.160	20.
13	35.110	20.
14	41.140	20.
15	33.520	20.
16	31.710	20.
17	37.580	20.
18	32.520	20.
19	37.550	20.
20	31.340	20.
21	44.530	30.
22	52.150	30.
23	47.240	30.
24	46.050	30.
25	43.710	30.
26	43.330	30.
27	52.150	30.
28	46.460	30.
29	44.420	30.
30	37.170	30.
31	60.570	40.
32	61.030	40.
33	52.240	40.
34	55.490	40.
35	59.020	40.
36	49.730	40.
37	53.080	40.
38	58.220	40.
39	67.550	40.
40	61.930	40.
41	62.960	50.
42	70.330	50.
43	77.540	50.
44	63.400	50.
45	74.750	50.
46	74.160	50.
47	66.590	50.
48	59.730	50.

PAGE 002

FILE: LARZA OUTPUT AL VM/50 - CONVERSATIONAL MONITOR SYSTEM

49	62.550	50.
50	63.300	50.
51	78.700	50.
52	77.770	50.
53	79.320	60.
54	73.510	50.
55	81.150	50.
56	90.250	50.
57	77.240	50.
58	73.350	60.
59	79.750	50.
60	73.000	50.
61	85.300	70.
62	97.500	70.
63	93.950	70.
64	84.130	70.
65	91.240	70.
66	90.950	70.
67	95.020	70.
68	99.330	70.
69	91.470	70.
70	95.350	70.
71	104.560	30.
72	105.350	30.
73	101.840	30.
74	104.760	30.
75	103.100	30.
76	110.250	30.
77	101.770	30.
78	101.840	30.
79	100.470	30.
80	91.300	20.
81	117.550	20.
82	112.620	30.
83	115.450	30.
84	106.530	30.
85	117.420	30.
86	114.590	20.
87	125.960	30.
88	122.100	30.
89	110.420	30.
90	119.410	30.
91	131.950	100.
92	126.290	100.
93	132.220	100.
94	111.130	100.
95	127.100	100.
96	125.250	100.
97	120.950	100.
98	117.010	100.
99	123.540	100.
100	112.110	100.

-- PLOT OF Y AGAINST X



K2 IS THE TOTAL SUM OF SQUARED DEVIATIONS

-- K2 105827.

-- K1 IS THE RESIDUAL SUM OF SQUARED DEVIATIONS

-- K1 2595.04

-- K3 IS THE REGRESSION SUM OF SQUARED DEVIATIONS

-- K3 103231.

-- K3 IS ALSO THE REGRESSION MEAN SQUARED DEVIATIONS

-- K3 103211.

-- K5 IS THE ERROR MEAN SQUARED DEVIATIONS

-- K5 25.4702

-- K6 IS THE F RATIO

-- K6 3896.75

-- K4 IS THE COEFFICIENT OF DETERMINATION, R-SQUARED

-- K4 0.975669

THE REGRESSION EQUATION IS

Y = 13.5 + 1.112 X1

COLUMN	COEFFICIENT	ST. DEV. OF COEF.	T-RATIO = COEF/S.D.
INTERCEPT	13.519	1.112	12.15
SLOPE	1.11859	0.01792	62.43

THE ST. DEV. OF Y ABOUT REGRESSION LINE IS

S = 5.147

WITH ( 100 - 2 ) = 98 DEGREES OF FREEDOM

R-SQUARED = 97.5 PERCENT

K-SQUARED = 97.5 PERCENT, ADJUSTED FOR D.F.

ANALYSIS OF VARIANCE

DF	SS	MS=SS/DF
REGRESSION 1	101227.5	101227.5
RESIDUAL 98	2595.0	26.5
TOTAL 99	105322.5	

ROW	X1	Y	PRED. Y VALUE	ST. DEV. PRED. Y	RESIDUAL	ST. RES.
1	10	19.410	24.704	0.957	-5.294	-1.08

(X-PRIME X)INVERSE

0	0.0466645	1
1	-0.0006606	0.0000121

HISTOGRAM OF C5, THE RESIDUALS

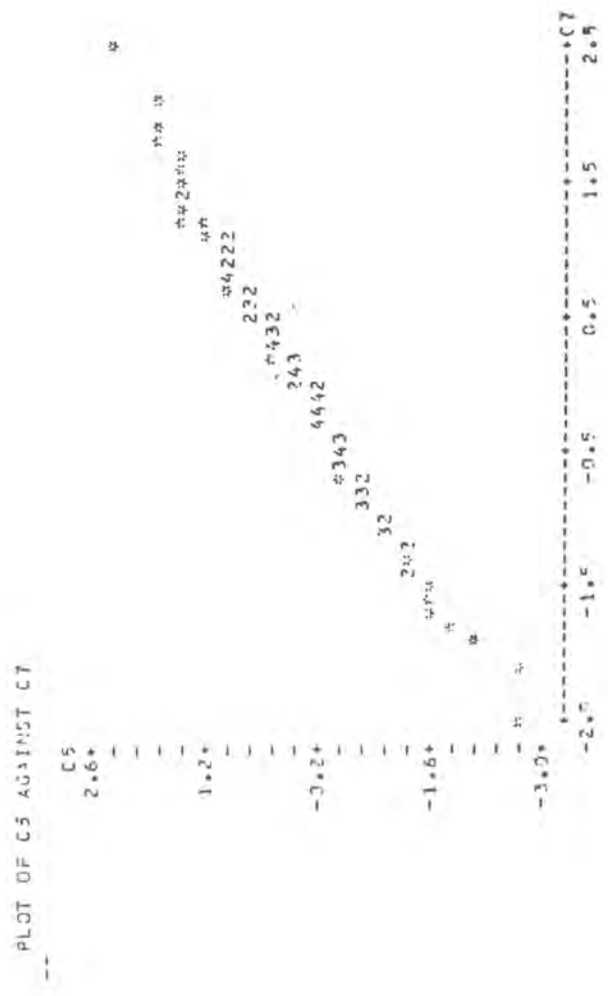
MIDDLE OF INTERVAL	NUMBER OF OBSERVATIONS	
-3.0	1	*
-2.5	1	*
-2.0	2	**
-1.5	6	***
-1.0	11	****
-0.5	18	*****
0.0	22	*****
0.5	15	****
1.0	12	***
1.5	8	**
2.0	3	*
2.5	1	*

ANSWER = 5.1469

COLUMN COUNT	C3	100
-5.3076	-1.5912	0.0745
-4.9198	5.2173	-5.4283
-0.7804	5.3173	-2.3076
1.7827	-4.6072	-2.5177
1.0471	-3.7233	5.1199
2.4242	2.7355	-6.0600
-5.2145	0.5629	9.3424
0.1340	-5.0736	5.3304
-0.9330	-5.1721	-1.9431
1.2233	9.5659	-3.1137
-3.0512	5.7176	7.1769
3.3255	-1.8537	-0.1512
-1.1752	1.9726	0.0700
-2.5574	-1.3046	3.3031
3.2956	0.5057	11.0154
5.0901	0.9195	3.2648
-7.5051	-0.4795	3.2603
		2.0521
		-4.4716
		-4.0097
		5.1199
		-0.6039
		-2.7592
		3.6323
		4.7373
		-3.5769
		-3.9183
		-7.7334
		4.0570
		7.3198
		-1.5906
		3.0095
		-1.44993
		-13.5019
		2.0521
		-4.4716
		-4.0097
		5.1199
		-0.6039
		-2.7592
		3.6323
		4.7373
		-3.5769
		-3.9183
		-7.7334
		4.0570
		7.3198
		-1.5906
		3.0095
		-1.44993
		-13.5019
		2.0521
		-4.4716
		-4.0097
		5.1199
		-0.6039
		-2.7592
		3.6323
		4.7373
		-3.5769
		-3.9183
		-7.7334
		4.0570
		7.3198
		-1.5906
		3.0095
		-1.44993
		-13.5019
		2.0521
		-4.4716
		-4.0097
		5.1199
		-0.6039
		-2.7592
		3.6323
		4.7373
		-3.5769
		-3.9183
		-7.7334
		4.0570
		7.3198
		-1.5906
		3.0095
		-1.44993
		-13.5019

PLOT OF RESIDUALS AGAINST PREDICTED Y

C5



ESTIMATES OF INTERCEPT AND SLOPE BY ITERATIVE METHOD



PAGE 001

FILE: RJNSAS SAS AL V4/SP - CONVERSATIONAL MONITOR SYSTEM

TITLE RUNNING LINEAR REGRESSION USING PROC GLM ON SAS1

DATA REGSAS;  
 INPUT Y X;  
 CAPDS;  
 17.41 10.  
 23.15 10.  
 33.62 10.  
 26.72 10.  
 24.60 10.  
 23.50 10.  
 25.59 10.  
 29.83 10.  
 19.37 10.  
 20.31 10.  
 42.93 20.  
 40.16 20.  
 35.11 20.  
 41.14 20.  
 31.93 20.  
 39.58 20.  
 32.69 20.  
 37.65 20.  
 31.34 20.  
 44.59 30.  
 52.15 30.  
 47.24 30.  
 45.05 30.  
 43.71 30.  
 43.33 30.  
 52.15 30.  
 46.45 30.  
 44.42 30.  
 37.19 30.  
 50.67 40.  
 51.03 40.  
 52.24 40.  
 55.43 40.  
 59.02 40.  
 49.93 40.  
 53.05 40.  
 58.02 40.  
 57.55 40.  
 51.53 40.  
 49.86 50.  
 75.93 50.  
 77.54 50.  
 53.40 50.  
 74.75 50.  
 74.15 50.  
 65.59 50.  
 55.73 50.  
 52.55 50.  
 63.30 50.  
 78.71 50.

77.97 60\*  
79.02 60\*  
79.51 60\*  
81.95 60\*  
90.25 60\*  
77.24 50\*  
73.85 60\*  
79.75 60\*  
78.00 60\*  
86.30 70\*  
77.50 70\*  
98.95 70\*  
84.13 70\*  
91.24 70\*  
95.95 70\*  
95.62 70\*  
89.93 70\*  
91.47 70\*  
95.85 70\*  
104.56 80\*  
105.35 80\*  
101.34 80\*  
104.96 80\*  
103.13 80\*  
110.26 80\*  
101.77 80\*  
101.83 80\*  
100.67 80\*  
91.80 80\*  
117.65 90\*  
112.62 90\*  
115.45 90\*  
136.53 90\*  
119.42 90\*  
114.67 90\*  
125.96 90\*  
122.10 90\*  
115.82 90\*  
119.41 90\*  
131.95 100\*  
124.25 100\*  
132.22 100\*  
111.13 100\*  
129.10 100\*  
125.25 100\*  
120.95 100\*  
119.01 100\*  
128.58 100\*  
112.11 100\*

PROC PRINT;  
PROC PLOT; PLOT Y=X;  
PROC CLM; MODEL Y=X;

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RUNNING LINEAR REGRESSION USING PROC GLM ON SAS

Q93 Y X

1	19.41	10
2	21.15	10
3	33.52	10
4	25.72	10
5	24.50	10
6	23.50	10
7	23.59	10
8	29.33	10
9	19.37	10
10	20.31	10
11	42.73	20
12	40.16	20
13	35.11	20
14	41.14	20
15	33.52	20
16	31.93	20
17	39.53	20
18	32.69	20
19	27.55	20
20	31.34	20
21	44.58	30
22	52.15	30
23	47.24	30
24	46.05	30
25	43.71	30
26	43.39	30
27	52.15	30
28	45.46	30
29	44.42	30
30	37.19	30
31	50.57	40
32	51.03	40
33	52.24	40
34	55.42	40
35	57.02	40
36	49.93	40
37	53.08	40
38	53.32	40
39	67.53	40
40	61.33	40
41	69.95	50
42	75.93	50
43	77.54	50
44	63.40	50
45	74.75	50
46	74.15	50
47	65.39	50
48	63.73	50
49	62.55	50
50	63.30	50
51	73.70	60
52	77.97	60
53	79.02	60
54	78.51	60
55	81.45	60
56	90.25	60

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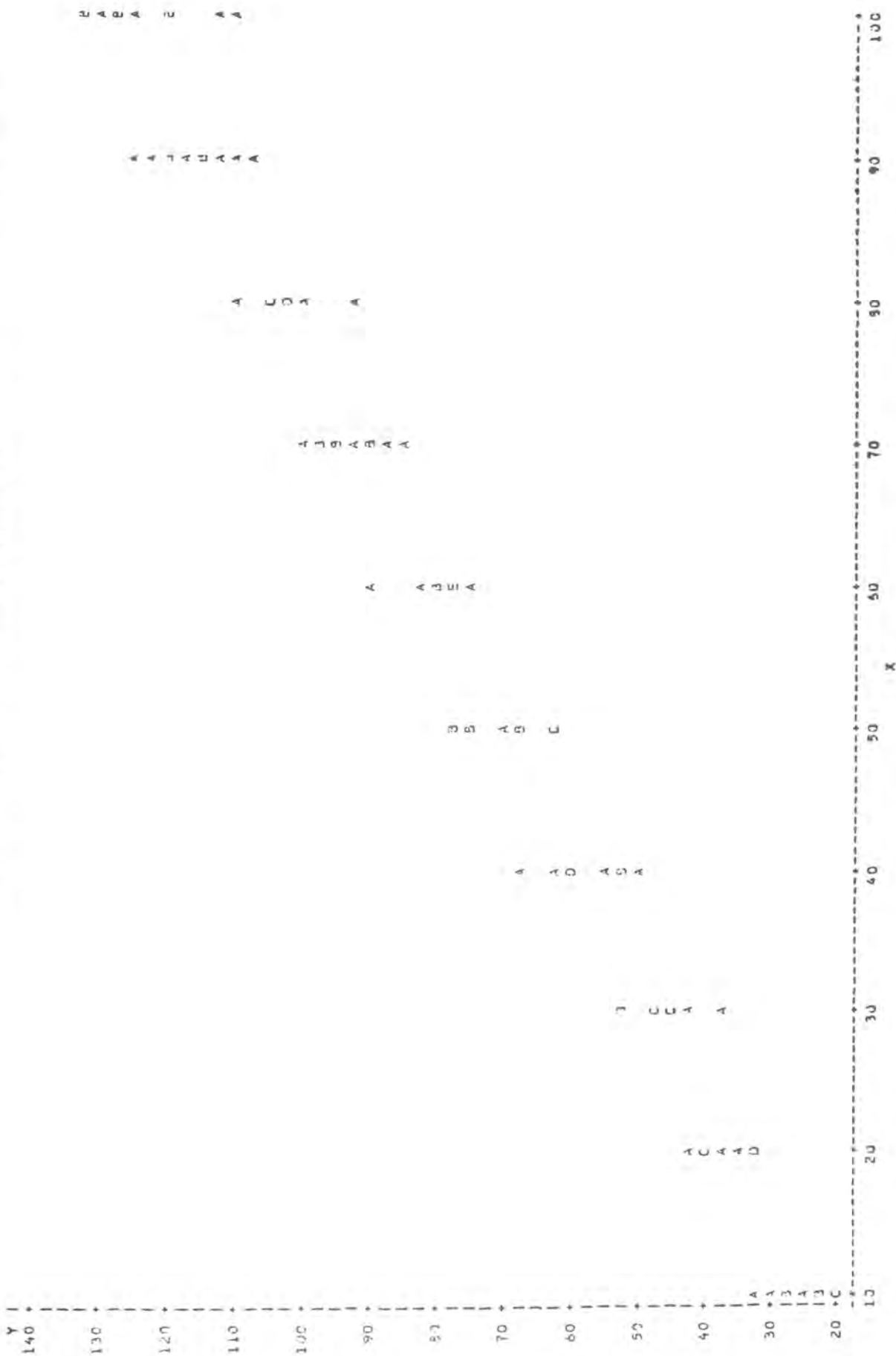
RUNNING LINEAR REGRESSION USING PROC GLM ON SAS

035 Y X

57	77.24	60
58	73.35	50
59	77.75	60
60	71.00	60
61	85.37	70
62	87.50	70
63	93.95	70
64	84.13	70
65	71.24	70
66	95.75	70
67	95.62	70
68	77.73	70
69	91.47	70
70	75.35	70
71	104.55	80
72	103.35	80
73	101.24	80
74	106.95	80
75	121.10	80
76	110.26	80
77	101.77	80
78	101.83	80
79	103.47	80
80	91.30	80
81	117.65	90
82	112.62	90
83	115.45	90
84	106.53	90
85	113.42	90
86	114.69	90
87	125.96	90
88	122.10	90
89	110.82	90
90	119.41	90
91	131.95	100
92	126.24	100
93	132.22	100
94	111.13	100
95	129.10	100
96	125.25	100
97	120.95	100
98	119.01	100
99	123.53	100
100	112.11	100

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RUNNING LINEAR REGRESSION USING PROC GLM ON SAS  
PLOT OF Y\*X LEGEND: A = 1 OBS, B = 2 OBS, ETC.



RUNNING LINEAR REGRESSION USING PROC GLM JM SAS  
GENERAL LINEAR MODELS PROCEDURE

DEPENDENT VARIABLE: Y

SOURCE	DF	SUM OF SQUARES	MEAN SQUARE	F VALUE	PR > F	R-SQUARE	C.V.
MODEL	1	103231.67391212	103231.67391212	3395.96	0.0001	0.975469	6.3528
ERROR	78	2596.3528735	25.47013467		ROOT MSE		Y MEAN
CORRECTED TOTAL	79	105827.73220000			5.14697609		75.00000000
SOURCE	TYPE I SS	F VALUE	PR > F	DF	TYPE III SS	F VALUE	PR > F
X	1	103231.67391212	0.0001	1	103231.67391212	3496.96	0.0001
PARAMETER	ESTIMATE	T FOR H0:	STANDARD ERROR OF ESTIMATE				
INTERCEPT	13.5163333	12.16	1.1113244				
X	1.11351212	62.43	0.1791912				

#### 4. NONLINEAR RELATIONSHIPS: NONLINEAR MODELS

##### 4.1 Some common bivariate relationships

In words	Differential form	Integrated form	Linear form	Examples
1. The rate of change of $y$ with respect to $x$ is a constant.	$\frac{dy}{dx} = b$	$y = a + bx$	same	-very few in biology
2. The percentage rate of change of $y$ with respect to $x$ is a constant. or The rate of change of $y$ with respect to $x$ is proportional to $y$ .	$\frac{dy}{y} = b \frac{dx}{x}$	$y = y_0 e^{bx}$ ("exponential")	$\log y = a + bx$ ("logarithmic")	-Light intensity $y$ at depth $x$ in a homogeneous lake. -The amount of a compound $y$ , over time $x$ , if $y_1 \rightarrow y_2$ at a constant rate. -Population size over time $x$ in an unlimited environment.
3. The percentage rate of change of $y$ is proportional to the percentage rate of change of $x$ .	$\frac{dy}{y} = b \frac{dx}{x}$	$y = Ax^b$ ("power law" or "allometric") pp	$\log y = a + b(\log x)$	-Body weight $y$ related to length $x$ for a growing animal (if shape & density do not change then $b=3$ ).
4. The rate of change of $y$ with respect to $x$ is proportional to the amount by which $y$ is less than $K$ .	$\frac{dy}{dx} = b(K-y)$	$y = K(1 - e^{-bx})$ ("Von Bertalanffy" or "monomolecular")	$\log\left(\frac{K-y}{K}\right) = a - bx$	-Growth in size $y$ with age $x$ for many animals ( $a=0$ if $y=0$ at $x=0$ ). -The amount of a compound $y_2$ over time $x$ , if $y_1 \rightarrow y_2$ at a constant rate.
5. The percentage rate of change of $y$ with respect to $x$ is proportional to the percent of $K$ not filled by $y$ .	$\frac{dy}{y} = \frac{b(K-y)}{K}$	$y = \frac{K}{1 + e^{-b(x-x_0)}}$ ("logistic")	$\log(K-y) = a - bx$	-Population size $y$ over time $x$ in a limited environment.
6. The percentage rate of change in the amount by which $y$ exceeds $a$ , with respect to $x$ , is proportional to the amount by which $y$ exceeds $a$ .	$\frac{dy}{y-a} = b(y-a)$	$y = a - \frac{1}{bx}$	$y = a - \frac{1}{b} (x^{-1})$	-If a rate $x$ (of mortality, say) is proportional to a stimulus $y$ (a toxicant dose, say), then $1/x = x^{-1} =$ expected time until the event occurs (time-to-death, say).

## 4.2 Some common bivariate relationships: assignment/tutorial

Refer to section 4.1. We will take each of the six models in turn.

### 4.2.1 $y = a + bx$ :

This is the model assumed by classical regression analysis, but it rarely describes relationships between variables in biology.

### 4.2.2 $y = y_0 e^{bx}$ :

Exponential growth or exponential decline of  $y$  for each unit increase in  $x$ .

Worked example: A human population grows as follows:

N	:	1000	1035	1066	1109	1147
t (yr)	:	0	1	2	3	4

The simplest sensible null hypothesis is that the population is growing at a constant % rate, which is described by the model  $N = N_0 e^{bt}$  or in linear form  $\log N = a + bt$  where  $a = \log N_0$ . If we calculate the regression of  $\log N$  on  $t$ , we find that  $\hat{b} = 0.0343$  with .95 cl of 0.0318 to 0.0369, and  $\hat{a} = 6.91$ . The  $r^2$  value, which measures the fraction of the variation in  $\log N$  that is related to time  $t$ , is 0.998. Thus our model is  $\log N = 6.91 + 0.0343t$ , or  $N = 1002e^{0.0343t}$ . Over one year the population increases by a factor  $e^{0.0343}$  which is  $100(1.0349 - 1) = 3.49\%/yr$ . The lower .95 cl. on  $\hat{b}$  is 0.0318, which as a % yr value is 3.23. The upper .95 cl on  $\hat{b}$  is 0.0369, which as a % yr value is 3.76. Thus the .95 cl on the % rate of increase of the population are 3.23 to 3.76%/yr.

Assigned problem: First run the above worked example in MINITAB, verify that you get the same results, and then PLOT  $N$  versus  $t$  as well as  $\log N$  versus  $t$ . Verify that  $\log N$  versus  $t$  is an apparently linear scatter. (Use LOGE to transform to logs, and



use EXPO to back-transform. See p. 47 of MINITAB manual.) Now do the following problem using MINITAB:

In 1938, Hatton scraped a rock clean of barnacles just before the annual larval set at St. Malo, France, and then at 6-month intervals he counted the number of barnacles left on the rock. The data follow (with  $t = 0$  at time of set):

N(no./cm <sup>2</sup> ) :	15.0	8.4	4.8	1.8
t (month) :	0	6	12	18

You should be able to :

- generate a plot of N versus t and of log N versus t
- determine the regression model, both as  
 $\log N = a + bt$  and as  $N = N_0 e^{bt}$
- put .95 cl on both  $\hat{b}$  and on the %/mo, mortality
- convert the estimate of %/mo. mortality to %/yr.  
 mortality (not by just multiplying by 12!)

4.2.3  $y = Ax^b$ : A power law relationship between y and x.

Worked example: A not-so-quick biologist, not knowing the relationship between the diameter of a circle and the area of a circle, decide to determine it empirically. So he used a compass to draw many circles of various sizes, and then he measured their diameters roughly with a ruler, and he measured their areas by laying them over graph paper to count little squares. (I told you he wasn't too bright.) The data follow:

Ar(cm <sup>2</sup> ):	.5	3.4	103	28	253	.07	60	10	158	66	85	144
D (cm) :	.8	2.1	11	5.5	18	.3	8.7	3.6	14	9.2	10.4	13.6

Our biologist can at least figure out that the area of something ought to be proportional to the square of a linear measurement on the same thing, so  $\text{Area} = AD^b$  where b ought to be equal to 2. Then,  $\log \text{Area} = a + b \log D$ , where  $a = \log A$ . If we calculate the regression of log Area on log D, we estimate  $\hat{b} = 2.008$ , with  $t = -0.1783$ . So our model is  $\log \text{Area} = -0.2377 + 2.008 \log D$ , or

Area = .788  $D^{2.01}$ . An exponent of  $b=2$  is certainly within our 0.95cl's, and the .95 cl's on A of 0.743 to 0.837 include  $\pi/(2)^2 = 0.785$ . The  $r^2$  is 0.99957.

Assigned problem: First run the above worked example in MINITAB, verify the results, and PLOT Area versus D as well as log Area versus log D. Now do the following problem using MINITAB:

Specimens of the unionid clam Anodonta grandis were collected from the Winnipeg River, in Canada, and length and volume were measured for each. The data follow:

v (ml):	11	12	18	24	27	30	36	40	43	47	54	61	76	73
L (mm):	48	53	60	62	67	70	73	74	77	79	83	86	93	94

You should be able to:

- plot V versus L and log V versus log L
- hypothesize what b should be, in a model of the form  $V = AL^b$ , assuming that shape does not change with growth in size.
- put .95 cl on  $\hat{b}$ , on  $a = \log A$ , and on A.
- say whether your hypothesized value of b appears to be correct.
- say in words what is the meaning of A in this model.
- do the following : Do the regression as 'REGR Ci ON 1 PRED. IN Cj, ST. RESIDS. IN Ck, PRED. Y IN Cm'. (See p. 66 of MINITAB manual.) Now do "PLOT Ck VERSUS Cm" to produce a plot of residuals ( $\hat{y} - y_{obs}$ ) versus predicted values ( $\hat{y}$ ). If the model is adequate, these should be patternless. Now do 'MPLOT Ci VERSUS Cj AND Cm VERSUS Cj' to produce a plot of the data (as log V versus log L) with the predicted values from the model  $\log V = a + b \log L$  also shown on the plot.

4.2.4  $y = K(1 - e^{-bx})$ : Growth to an asymptote with no inflection.

Worked example: A compound A is being converted to compound B at a constant % rate, and we know that there will be 100g when it is all converted. The data collected throughout the conversion

are:

B(g):	0	51	75	87	94	97
t(hr.):	0	1	2	3	4	5

In linear form the model is  $\log \frac{(K - B)}{K} = a - bt$ , with  $K = 100$ , or  $\log(1 - 0.01B) = a - bt$ , where  $a = 0$  if  $B = 0$  at  $t = 0$ . If we calculate the regression of  $\log(1 - 0.01B)$  on  $t$ , we estimate  $\hat{b} = -0.6996$  and  $\hat{a} = 0.00576$ , with .95 cl of  $-0.72$  to  $-0.68$  and  $-0.057$  to  $0.068$  respectively. The .95 cl on  $a$  include  $\hat{a} = 0$ . Notice that  $K$  is within a log term in the linear model which prevents us from solving for  $\hat{K}$  directly if it is unknown. You could find  $\hat{K}$  by trial and error - just search for the value of  $K$  that gives a minimum residual errors from the regression of  $\log(K-B)/K$  on  $t$ . Alternatively you can solve for  $K$  directly by using the Walford Plot technique, to be described later.

Assigned problem: Run the above worked example in MINITAB, verify the results, and plot  $B$  versus  $t$  as well as  $\log(1 - 0.01B)$  versus  $t$ . A problem based on a situation where we do not know  $K$  will be worked later, in relation to the Walford Plot technique.

4.2.5  $y = K/(1 + e^{-b(x-x_0)})$  ): Growth to an asymptote with an inflection halfway up.

Worked example: A dose-mortality experiment yields the following results, where  $M$  is % dead at time  $t$  :

M (%):	0	0	6	18	55	78	95	97	100
t (hr):	0	1	2	3	4	5	6	7	8

In linear form the model is  $\log \frac{(K - M)}{M} = a - bt$ , with  $K = 100$ . Values of  $M = 0$  or  $N = 100$  can not be used. The parameter  $b$  represents the rate of ascent, the parameter  $a$  "positions" the curve on the  $t$ -axis (allowing calculation of the  $t_{m:50}$  or  $LT_{50}$

estimate). If we calculate the regression of  $\log(100 - M)/M$  on  $t$ , we estimate  $\hat{b} = -1.302$  and  $\hat{a} = 5.158$ . Again  $K$  is within a log term which prevents us from solving for  $\hat{K}$  directly if it is unknown, and again it would have to be found by trial and error or by the Walford Plot technique. Here confidence limits on  $\hat{a}$  or  $\hat{b}$  do us little good. From .95 cl on  $a$  we could only calculate 0.95 cl on  $M$  at  $t = 0$ . The best thing to do is to replicate the experiments, estimate  $LT_{50}$  for each one, and use those as replicate  $LT_{50}$  estimates for statistical tests. For this set of data the  $LT_{50}$  estimate is found by solving  $\log(100 - 50)/50 = \hat{a} + \hat{b}t$ , and it is  $\hat{t}_{m=50} = 4.04$  hr. This is  $\hat{x}_0 = \hat{t}_0$  in the integrated model.

Assigned problem: Run the above worked example in MINITAB, verify the results, and plot  $M$  versus  $t$  as well as  $\log(100 - M/M)$  versus  $t$ . Also tray the following plot, assuming that  $t$  values are in  $Ch$   $M$  values are in  $Ci$ , and the values 0 to 8 in increments of 0.5 are in  $Cj$ . Enter the command "LET  $Ck = 100/(1+EXPO(\hat{b} * (Cj - LT_{50}))$ " and then "MPLLOT  $Ci$  VERSUS  $Ch$  AND  $Ck$  VERSUS  $Cj$ ". You will now have observed  $M$  and  $t$  values, and the fitted curve, on the same plot. A problem based on a situation where we do not know  $K$  will be worked later, using the Walford plot technique.

4.2.6  $y = a - 1/bx$ : A hyperbolic relationship in which  $y$  is asymptotic to  $x = 0$  and  $X$  is asymptotic to  $y = a$ .

Worked example: In a dose-mortality experiment the % dead at different doses is observed at each of a series of times, but even under zero dose (control) conditions the animals can not be held longer than 48 hours without mortality. We would like to estimate the dose which would cause 50% mortality (the  $LD_{50}$ ) over a very long time, as would be the case with chronic exposure in the natural environment. The data from the 48-hour experiment follow:

LD (ppm):	16	13	13	12	11
t (hr) :	3	6	12	24	48

Let our model be  $LD = a + b'/t$ , where  $b' = -1/b$ . Since  $LD_{50}$  becomes equal to  $a$  as  $t$  becomes very large, therefore the parameter  $a$  provides our estimate of  $LD_{50}$  for a very long exposure time. As  $t$  approaches zero the  $LD_{50}$  becomes very large, which implies that the organisms can withstand a very high dose for a very short time.

If we calculate the regression of  $LD_{50}$  on  $1/t$  we estimate  $\hat{b}' = 14.2$  and  $\hat{a} = 11.17$ . The .95 cl on  $\hat{a}$  are 9.93 to 12.40, which are also the .95 cl on  $LD_{50}$  for a very long exposure time.

Assigned problem: Run the above worked example in MINITAB, verify the results, and plot  $LD_{50}$  versus  $t$  as well as  $LD_{50}$  versus  $1/t$ .

#### 4.2.7 Estimation of an $LD_{50}$ :

There are two standard models. One is the logistic, which you used to estimate an  $LT_{50}$ . The only difference here would be that you would calculate the regression of  $\log((100-M)/M)$  on  $D$ , the dose, at each time  $t$ . (Previously we regressed  $\log((100 - M)/M)$  on  $t$ , for each dose  $D$ , to obtain an  $LT_{50}$ ). The other model is the probit or cumulative normal model. It is probably the more commonly used, but it has the disadvantage that an equation cannot be given for this model! What is needed is the integral of the normal distribution, which has no exact integral. Therefore computer programs for probit analysis do a numerical integration of a normal distribution. SAS has a probit analysis procedure of this kind.

Assigned problem: Analyse the following data using the SAS probit procedure (PROC PROBIT):

Dose (ppm):	2	3	4	5	6	7
# animals dead:	6	18	55	78	95	97
# animals exposed:	100	100	100	100	100	100

The SAS job file should be as follows:

```
TITLE -----;
DATA -----;
INPUT DOSE N RES;
CARDS;
  2 100 6
  3 100 18
  4 100 55
  5 100 78
  6 100 95
  7 100 97

PROC PROBIT;
VAR DOSE N RES;
```

Those who have good memories will realize that these data are the same as for the assigned problem with the logistic, except that the variable "time" in hours, is now called "dose", in ppm. So you can compare your LD<sub>50</sub> estimate in this probit analysis with the LT<sub>50</sub> obtained in the logistic model analysis.

4.2.8 Walford plots: The growth model  $y = K (1 - e^{-bx})$  was exemplified in section 4.2.4 by the following data set:

y (mm):	0	51	75	87	94	97
x (yr):	0	1	2	3	4	5

If we are told that  $K = 100$ , then we find  $b = 0.6996$  by fitting the linear model  $\log((K - y)/K) = a - bt$ , where we expect  $a$  to be zero if  $y = 0$  at  $x = 0$ . Let us say that we do not know  $K$  a priori, and that we rewrite the data in the form:

$y_x$	:	0	51	75	87	94
$y_{x+1}$	:	51	75	87	94	97

A plot of  $y_{x+1}$  versus  $y_x$  will form a more or less straight line  $y_{x+1} = a' + b'y_x$ . The parameter  $a'$  is the estimated growth during the first time unit (a year in this case), and the parameter  $b'$  is the fraction of the total growth (to the asymptote) which remains after the first year. For these data we find  $\hat{a}' = 50.6$  and  $\hat{b}' = 0.491$ . The parameters of the  $y = K(1 - e^{-bx})$  model are related to  $a'$  and  $b'$ , as  $b = \log b'$  and  $K = a'/(1 - b')$ . For these data  $b = -0.71$  and  $K = 99.5$ , which are close to the previous values. Confidence limits can be placed on  $a'$  and  $b'$  in the same way as previously. Confidence limits on  $b$  would be easy to calculate, but those on  $K$  would not be because  $a'$  and  $b'$  will not be independent of each other.

Such " $y_{x+1}$  versus  $y_x$ " data arise very frequently. In fact we often do not have observations on  $y$  at known  $x$  values. For example, annual rings in trees, fish scales, clam shells, etc. - can be analyzed in this way. If we measure the length at an annual ring and let that be a  $y_x$  value, then we can measure the length at the next annual ring "outward" and let that be the  $y_{x+1}$  value, and so on. We need not know what the value of  $x$  is, and yet we can derive the growth curve for this plant or animal! Another source of such data is mark-recapture studies, where we catch animals one year, measure their sizes (weight, length, or any other size measure), give them individual marks, and release them. One year later you recapture at least some of them, re-measure them, and proceed as above.

What must you assume? First of all, the time interval must be exactly the same (usually one year) for all animals. Second, if you are using annual rings you must be sure that they really are annual rings. Third, you must be fitting the correct Walford Plot model. If the  $y_{x+1}$  versus  $y_x$  plot is not a straight line,

then  $y = K(1 - e^{-bx})$  is not the appropriate growth model.

model. A logistic model,  $y = K/(1 + e^{-b(x - x_0)})$ , will be appropriate if the Walford plot of  $(y_{x+1})^{-1}$  versus  $(y_x)^{-1}$  is linear. A third growth model, commonly used for fish, is the Gompertz model. If this is appropriate, then a plot of  $\log y_{x+1}$  versus  $\log y_x$  will be linear. In fact there is a whole family of growth models which includes these three, and all have corresponding Walford Plot models.

Assigned problem: Intensive trapping of the Singapore Sling Sloth (SSS for short) was done on the N.U.S. campus. All SSS were weighed, individually marked, and released. A year later, 12 SSS were recaptured and reweighed, yielding the following data:

Animal:	1	2	3	4	5	6	7	8	9	10	11	12
1983 wt.(g):	14	17	25	30	32	40	46	50	52	58	58	70
1984 wt.(g):	53	60	49	57	67	59	67	78	72	73	78	86

Regress 1984 weight on 1983 weight. Plot the data. Estimate  $a'$  and  $b'$ , and from them calculate  $K$  and  $b$  in the Von Bertalanffy model. Plot the curve  $wt. = K(1 - e^{-bx})$  for age  $x$  from 0 to 10 using MINITAB commands.

#### 4.2.9 Ratio variables

##### 4.2.9.1 Introduction

Variables derived as the ratio of two observed variables can cause serious problems in statistical analysis. There is no problem when the denominator is a constant, as in a dose-mortality experiment where the variable "% dead" is used and is calculated as the number which have died divided by the total number at the beginning, times 100. That would amount to a change of scale, as would be the case if no. of organisms per m<sup>2</sup> were recorded as no. per cm<sup>2</sup> by dividing by 10<sup>4</sup>. However where the denominator variable has substantial variance, estimates of the true mean of the ratio are biased and any possible correlation - between the ratio and the variables which go into the ratio - is obscured.



4.2.9.2 Worked example: A student collects a large number of hermit crabs covering a range of sizes, and expels them from their shells by applying heat to the top of the shell. Each "naked" crab is weighed and then given a choice of a range of shell sizes of the same kind (the same gastropod mollusc species) of shell. The question is, "What is the relationship between the weight of the crab and the weight of the shell that it chooses to inhabit and to carry around?"

If the student just derives the variable "ratio of shell weight to crab weight", calculates it for each combination of crab and chosen shell, and then finds the mean, standard error and .95 cl on the mean, there are two problems. The first is that a ratio variable is involved, one where the denominator is a response variable with its own substantial variance. The second is that the variance in the denominator, crab weight, may be correlated with the derived variable, ratio of shell weight to crab weight. That is, small crabs may be able to carry shells that are larger in proportion to their size, compared with large crabs.

The first problem is one of a possibly biased estimate of mean ratio of shell weight to crab weight, and an inflated estimate of precision (variance, standard error and confidence limits). The second problem relates to a well-known principle in biology and in architecture: objects which must stand up above a surface must maintain their weight-to-basal area ratio as they increase in weight. Obviously an object suspended in liquid, such as a whale or a ship, is spared this problem. However for an animal on land, or for a building supported by columns, the load-bearing cross sectional area in contact with the ground increases as the square of a linear dimension (e.g. length or height) whereas the weight to be supported increases as the cube of that same linear dimension. Therefore you cannot design an elephant by describing an orders-of-magnitude larger mouse. The same is true of the architect who must design a larger version of an existing building.

Assume that the student's data are as follows:

$W$ (g):	1.19	2.42	3.49	2.14	1.65	0.89	3.29	4.26	1.38	0.93
$W_c^s$ (g):	1.76	2.59	6.98	3.73	1.99	1.28	4.11	10.11	2.24	2.39
(con't)	4.06	3.34	1.16	3.45	1.67	0.83	2.25	1.31	4.12	1.51
	7.04	5.08	2.29	8.79	2.15	1.72	5.75	1.52	3.87	1.42

If shell weight does increase at the same % rate as crab weight then the ratio of shell weight to crab weight will stay the same. For example, if a 4 g crab carries a 6 g shell then doubling the weight of both (a 100 % increase) would result in a 8 g crab carrying a 12 g shell. The ratio of 1.5 remains the same. Therefore we have the process "the % rate of change of one variable is proportional to the % rate of change of another variable", which leads to the power law model, and to the log-log regression model. In this case,

$$\frac{d W_s}{W_s} = b \frac{d W_c}{W_c}$$

$$W_s = A W_c^b$$

If the  $W_s/W_c$  ratio remains the same, then the % rate of change of  $W_s$  is equal to the % rate of change of  $W_c$ , and  $b$  should be equal to 1. If, on the other hand, our "load-bearing cross sectional area" model is applicable then  $b$  should be equal to 2/3.

In linear form our model is

$$\log W_s = a + b \log W_c,$$

where  $a = \log A$ . If we regress  $\log W_s$  on  $\log W_c$ , we estimate  $\hat{b} = 0.742$ , with .95 cl of 0.524 to 0.959, and  $\hat{a} = -0.170$ , with 0.95 cl of -0.453 to 0.114. The estimate of  $A$  is  $\hat{A} = 0.844$ , with 0.95 cl of 0.636 to 1.12.

Therefore we would conclude that  $b$  is significantly different from 1 but not from  $2/3$ , so that the  $W_s/W_c$  ratio does change (decreases) as  $W_c$  increases. It changes in a manner that is compatible with the "load-bearing cross sectional area" model. We would conclude that  $a$  is not significantly different from 0, and that  $A$  is not significantly different from 1. Note that  $\hat{A}$  is the estimate of the ratio of  $W_s$  to  $W_c$  at a crab weight of 1 g, and in a case where  $b$  was equal to 1 it would be an estimate of the ratio  $W_s/W_c$  at all values of  $W_c$ .

Now let us reformulate the model in terms of the ratio  $W_s/W_c$ .

If  $W_s = 0.844 W_c^{0.742}$ , then

$$\frac{W_s}{W_c} = 0.844 W_c^{-0.258}$$

A plot  $W_s/W_c$  versus  $W_c$  is attached.

Was this analysis valid? The truth is that I simulated these data under the model:

$$\frac{W_s}{W_c} = W_c^{-1/3}, \text{ which corresponds to}$$

$$W_s = W_c^{2/3}.$$

#### 4.2.10. Job Listings and Outputs.

FILE: LAB34    INITIAL AT VV/SP - CONVERSATIONAL MONITOR SYSTEM

PAGE 001

```
READ C1-C2
12.0 0
8.4 6
4.4 12
1.4 12
LOGE C1, PUT IN C3
NOTE C3 IS THE LOGE TRANSFORMATION OF N
PRINT C1-C3
PLOT C1 V5 C2
PLOT C3 V5 C2
RESOR C3 1 C2
EXPJ 2.732, PUT IN K1
EXPJ -0.1153, PUT IN K2
LET K3=100*(K2-1)
NOTE K3 IS THE MORTALITY RATE IN % PER MONTH
LET K4=-0.1153-0.0507
NOTE K4 IS THE LOWER 95% CONFIDENCE LIMIT FOR B-HAT
LET K5=-0.1153+0.0507
NOTE K5 IS THE UPPER 95% CONFIDENCE LIMIT FOR B-HAT
PRINT K3-K5
EXPJ K4, PUT IN K6
EXPJ K5, PUT IN K7
LET K8=100*(K6-1)
NOTE K8 IS THE LOWER 95% CONFIDENCE LIMIT FOR MORTALITY RATE(%/MTH)
LET K9=100*(K7-1)
NOTE K9 IS THE UPPER 95% CONFIDENCE LIMIT FOR MORTALITY RATE(%/MTH)
PRINT K8-K9
LET K10=EXPJ(12*(1-0.1153))
LET K11=100*(K10-1)
NOTE K11 IS THE MORTALITY RATE IN % PER YEAR
PRINT K10-K11
STOP
```

FILE: BARNACLE OUTPUT AI V4/SP - CONVERSATIONAL MONITOR SYSTEM

NAME: KAM SUAN PHENG

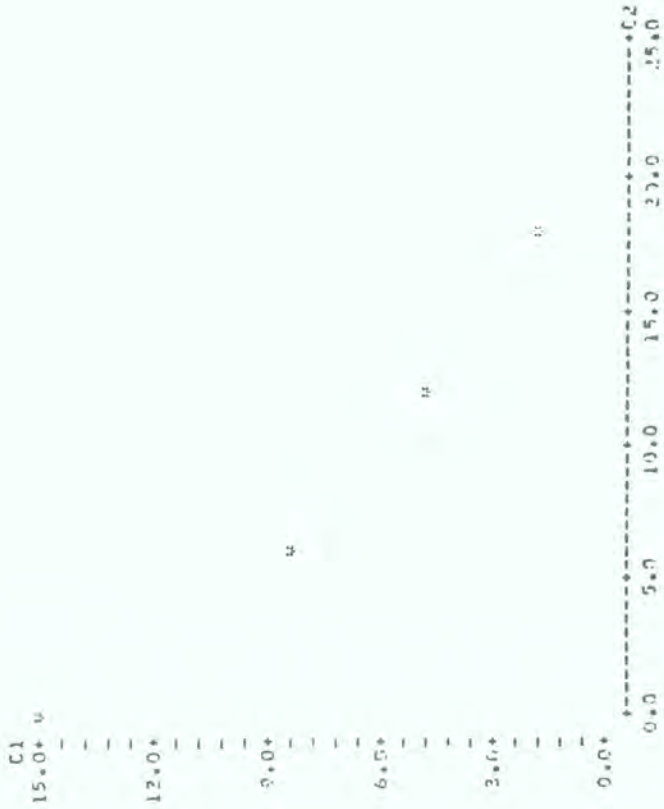
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LINEAR REGRESSION ANALYSIS ON THE BARNACLE PROBLEM

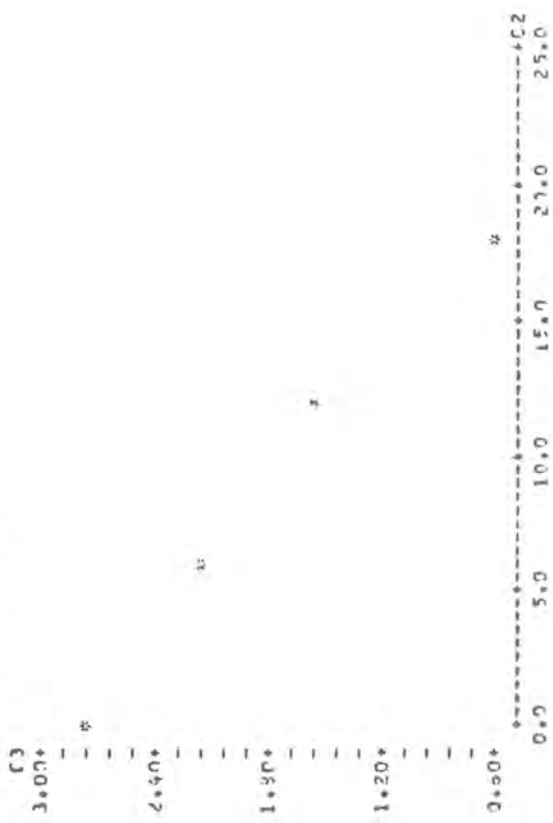
COEFF	C1	C2	C3
COJNT	4	4	4
ROV			
1	15.0000	0.	2.70805
2	8.4000	6.	2.12873
3	4.8000	12.	1.56962
4	1.8000	18.	0.58779

--

PLOT OF N (NO. OF BARNACLES PER SQ.CM) VERSUS T (MONTH)



PLOT OF LOG(M) VERSUS T



LINEAR REGRESSION ANALYSIS (WITH DATA TRANSFORMATION)  
 MODEL IS:  $LOG(M) = A + BT$  (1)  
 THE EXPERIMENTAL FORM IS:  $Y = A EXP(BT)$  (2)  
 WHERE  $LOG(M) = A$  (3)  
 THE REGRESSION EQUATION IS  
 $LOG(M) = 2.79 - 0.115T$

COLUMN	COEFFICIENT	ST. DEV. OF COEF.	T-RATIO = COEF/ST. DEV.
INTERCEPT	2.7952	0.1322	21.07
SLOPE	-0.11536	0.01173	-9.77

THE ST. DEV. OF Y ABOUT REGRESSION LINE IS  
 $S = 0.1580$   
 WITH ( 4-2 ) = 2 DEGREES OF FREEDOM  
 R-SQUARED = 99.0 PERCENT  
 R-SQUARED = 96.0 PERCENT ADJUSTED FOR D.F.

ANALYSIS OF VARIANCE

SOURCE	DF	SS	MS=SS/DF
REGRESSION	1	2.39459	2.39459
RESIDUAL	2	0.34994	0.17497
TOTAL	3	2.74453	

FILE: BARNACLE OUTPUT A1 VM/SP - CONVERSATIONAL MONITOR SYSTEM

CURSIN-WATSON STATISTIC = 2.27

-- K1 = EXP(A) = 16.2173

-- FROM EQUATION (J) ABOVE, M = EXP(A) = K1

THEFORE THE REGRESSION MODEL IN THE EXPONENTIAL FORM IS:

$$Y = 16.22 \exp(-0.1153T)$$

-- K2 = EXP(-0.1153) = 0.8911

-- K3 = -10.8901

K3 IS THE ESTIMATED MORTALITY RATE IN % PER MONTH

K4 = -0.156009

K5 = -0.0540000

K4 AND K5 ARE THE LOWER AND UPPER 95% CONFIDENCE LIMITS FOR B-HAT

-- K6 = EXP(K4) = 0.7639

-- K7 = EXP(K5) = 1.0435

-- K8 = -15.2054

K9 = -6.25575

K8 AND K9 ARE THE LOWER AND UPPER 95% CONFIDENCE LIMITS FOR THE MORTALITY RATE (1/MONTH)

TO CALCULATE THE MORTALITY RATE IN #/YEAR

THE REGRESSION MODEL IS  $\text{LOG}(Y) = A + 12ST$

WHERE  $T = 12t$ ;  $t$  IS TIME IN YEAR

K10 = EXP(12 \* (-0.1153)) = 0.250575

K11 = -74.0325

K11 IS THE MORTALITY RATE = -74.0% PER YEAR

NOTE: MORTALITY STATISTICS DEPT. PENN STATE UNIV. RELEASE 21.1.80  
 STPAGE AVAILABLE 4002

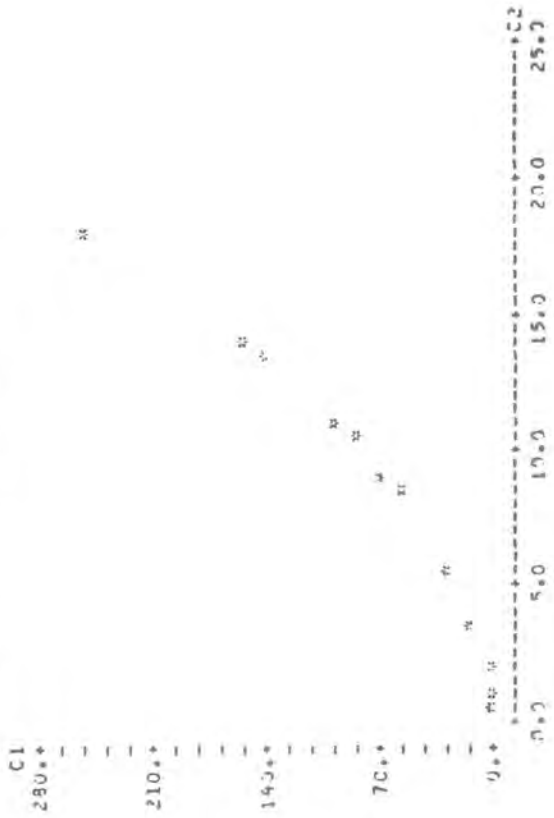
FILE: LAB301 OUTPUT A1 V4/SP - CONVERSATIONAL MONITOR SYSTEM

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LINEAR REGRESSION ANALYSIS OF AREA OF CIRCLE ON DIAMETER  
 MODEL: LOG(A) = A + B\*LOG(D)

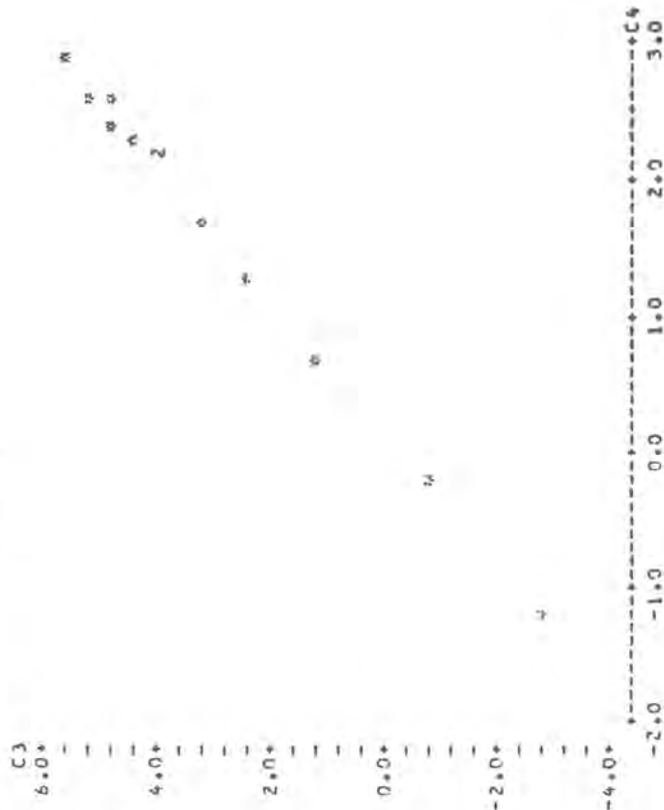
ROW	C1 COUNT	C2 D	C3 LOG(A)	C4 LOG(D)
1	0.500	0.8000	-0.59315	-0.22314
2	3.400	2.1000	1.22377	0.74194
3	103.000	11.0000	4.63473	2.39799
4	28.000	5.5000	3.33220	1.70475
5	253.000	18.0000	5.53339	2.89037
6	0.070	0.3000	-2.65026	-1.20397
7	60.000	9.7000	4.09434	2.16332
8	10.000	3.6000	2.30258	1.28093
9	153.000	14.0000	5.76260	2.63906
10	66.000	9.2000	4.19965	2.21920
11	85.000	10.6000	4.44265	2.34181
12	144.000	13.6000	4.96981	2.61007

--- PLOT OF AREA OF CIRCLE (A) VERSUS DIAMETER (D)





PLOT OF LOG(A) VERSUS LOG(I)



--

THE REGRESSION EQUATION IS  
 $LOG(A) = -0.238 + 2.013 LOG(I)$

COLUMN	COEFFICIENT	ST. DEV. OF COEF.	T-RATIO = COEF/S.D.
INTERCEPT	-0.23767	0.02654	-9.02
SLOPE	2.00324	0.01311	153.19

THE ST. DEV. OF Y ABOUT REGRESSION LINE IS  
 $S = 0.05512$   
 WITH ( 12- 2) = 10 DEGREES OF FREEDOM

R-SQUARED = 100.0 PERCENT  
 N-SQUARED = 100.0 PERCENT, ADJUSTED FOR D.F.

ANALYSIS OF VARIANCE

DF	SS	MS=SS/DF	ST. RES.
REGRESSION 1	71.29090	71.29090	2.778
RESIDUAL 10	0.03330	0.00330	-0.10 X
TOTAL 11	71.32124		

ROW	X1	Y	PRED. Y	ST. DEV.	RESIDUAL	ST. RES.
4	1.70	3.3322	3.1858	0.0150	0.1464	
6	-1.20	-2.6593	-2.6555	0.0404	-0.0037	

R DENOTES AN OBS. WITH A LARGE ST. RES.  
 X DENOTES AN OBS. WHOSE X VALUE GIVES IT LARGE INFLUENCE.  
 DURBIN-WATSON STATISTIC = 1.70

\*\*\* MIJITAD \*\*\* STATISTICS DEPT \* PENN STATE UNIV. \* RELEASE 81.1 \*  
 ST\*PAGE AVAILABLE 4930

EXAMPLE OF PLOT PROCEDURE USING FORTRAN

THE DATA AS READ IN ARE:

X	Y
10.000000	19.410000
20.000000	42.920000
30.000000	54.530000
40.000000	60.560000
50.000000	67.950000
60.000000	73.670000
70.000000	86.800000
80.000000	104.550000
90.000000	117.640000
100.000000	131.940000

HORIZONTAL AXIS IS DIMENSION 1  
 VERTICAL AXIS IS DIMENSION 2

HORIZONTAL AXIS

MINIMUM VALUE= 10.00000  
 MAXIMUM VALUE= 100.00000  
 SCALING UNIT = 1.00000  
 ONE TICK= 0.000

VERTICAL AXIS

MINIMUM VALUE= 19.41000  
 MAXIMUM VALUE= 131.95000  
 SCALING UNIT = 5.52700  
 ONE TICK= 0.000

OVERLAPPING OBJECTS (NOT PLOTTED)

ID NUMBER COORDINATES

EXAMPLE OF LINEAR REGRESSION ANALYSIS (WITH TRANSFORMATION)  
USING FORTRAN  
MODEL: LOG(Y) = A + B\*LOG(X)

THE DATA AS READ IN, BEFORE ANY TRANSFORMATION, ARE:

X	Y
0.500	0.500
1.000	1.000
2.000	2.000
5.000	5.000
10.000	10.000
20.000	20.000
50.000	50.000
100.000	100.000
200.000	200.000
500.000	500.000
1000.000	1000.000
2000.000	2000.000
5000.000	5000.000
10000.000	10000.000

THE DATA AFTER TRANSFORMATION, IF ANY, ARE:

X	Y
-0.223	-0.693
0.742	1.224
2.303	4.635
1.703	3.332
2.890	5.333
-1.204	-2.659
2.167	4.094
1.291	2.303
2.639	5.061
2.219	4.190
2.342	4.443
2.510	4.970

X MEAN= 1.63 Y MEAN= 3.04  
X VARIANCE= 1.51 Y VARIANCE= 6.48 XY COVARIANCE = 3.23

THE REGRESSION LINE IS Y = -0.2077 + 2.0032X

THE ANALYSIS OF VARIANCE TABLE IS:

SOURCE	SUM OF SQUARES	MEAN SQUARE	F-STATISTIC
1	71.27	71.27	11.27
10	0.03	0.00	0.00
11	71.32		

R-SQUARED=0.9998 PERCENT R-SQUARED=99.98

Y-PREDICTEDS AND Y-RESIDUALS FOLLOW.

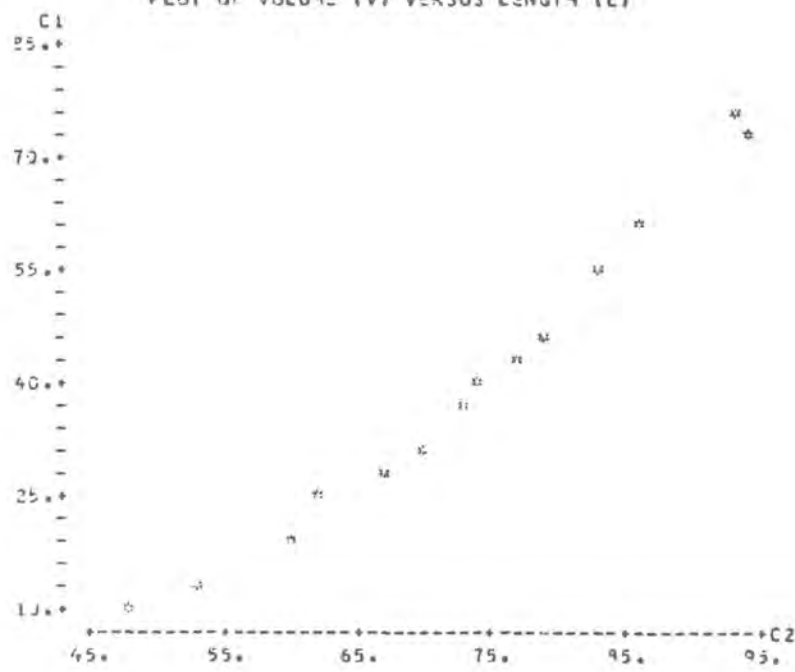
```
READ C1-C2
11 43
12 53
14 60
24 62
27 67
30 70
36 73
40 74
43 77
47 79
54 83
61 86
76 93
73 94
LOGE C1, PUT IN C3
NOTE C3 IS THE LOGE TRANSFORMATION OF V
LOGE C2, PUT IN C4
NOTE C4 IS THE LOGE TRANSFORMATION OF L
PRINT C1-C4
PLOT C1 VS C2
PLOT C3 VS C4
REGR C3 1 C4, STD RESIDUALS IN C5, PREDICTED LOG(V) IN C6
WIDTH 100, 50
MPLT C3 VS C4 AND C6 VS C4
STOP
```

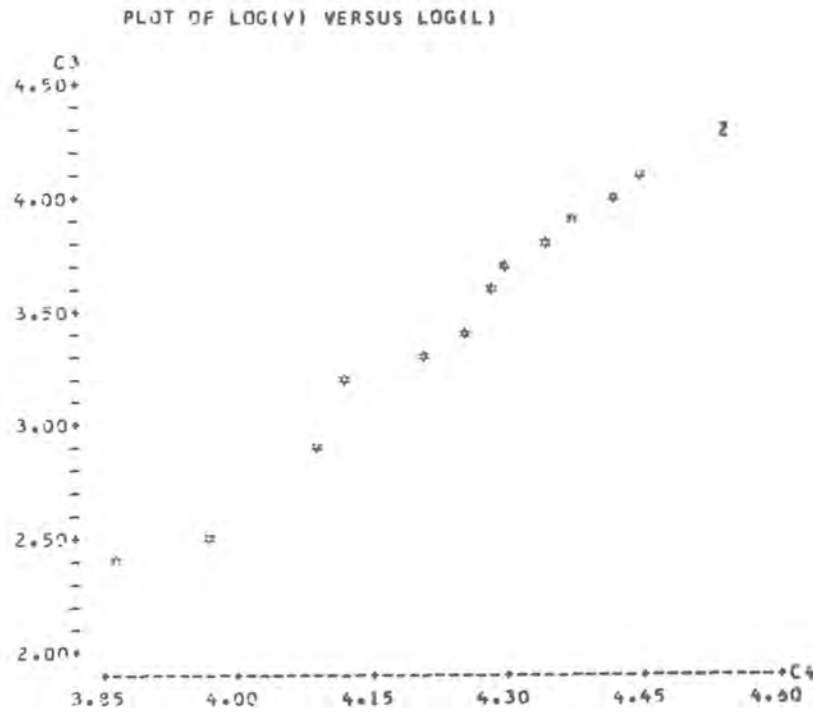
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 MAY 24 1985 \*\*\* NATIONAL UNIVERSITY OF SINGAPORE - LOCAL VERSION 02/12/1982  
 LINEAR REGRESSION OF CLAM DATA  
 MODEL: LOG(V) = A + B\*LOG(L) WHERE V IS VOLUME IN ML  
 L IS LENGTH IN MM

```
--
```

COLUMN	C1	C2	C3	C4
COUNT	14	14	14	14
ROW	V	L	LOG(V)	LOG(L)
1	11.	44.	2.37799	3.87120
2	12.	53.	3.48491	3.97029
3	13.	60.	2.87037	4.09434
4	24.	62.	3.17805	4.12713
5	27.	67.	3.27584	4.20469
6	30.	70.	3.40120	4.24850
7	36.	73.	3.58352	4.29046
8	40.	74.	3.69888	4.30406
9	43.	77.	3.76120	4.34381
10	47.	79.	3.35015	4.36945
11	54.	83.	3.79893	4.41984
12	51.	86.	4.11097	4.45435
13	76.	93.	4.33073	4.53260
14	73.	94.	4.29046	4.54327

PLLOT OF VOLUME (V) VERSUS LENGTH (L)





THE REGRESSION EQUATION IS  
 $LOG(V) = - 2.532 + 3.056 LOG(L)$

COLUMN	COEFFICIENT	ST. DEV. OF COEF.	T-RATIO = COEF/S.D.
INTERCEPT	-2.5322	0.4032	-27.82
SLOPE	3.05662	0.09364	32.64

THE ST. DEV. OF Y ADJUT REGRESSION LINE IS  
 $S = 0.06763$   
 WITH ( 14- 2) = 12 DEGREES OF FREEDOM

R-SQUARED = 98.9 PERCENT  
 R-SQUARED = 79.4 PERCENT, ADJUSTED FOR D.F.

ANALYSIS OF VARIANCE

DUE TO	DF	SS	MS=SS/DF
REGRESSION	1	4.873100	4.873100
RESIDUAL	12	0.054880	0.004573
TOTAL	13	4.927977	

ROW	X1	Y	PRED. Y	ST.DEV.	RESIDUAL	ST.RES.
	C4	C3	VALUE	PRED. Y		
1	3.87	2.3779	2.3706	0.0414	0.0973	1.32 X
2	3.97	2.4849	2.6335	0.0333	-0.1196	-2.024

R DENOTES AN OBS. WITH A LARGE ST. RES.  
 X DENOTES AN OBS. WHOSE X VALUE GIVES IT LARGE INFLUENCE.

DURBIN-WATSON'S STATISTIC = 2.06

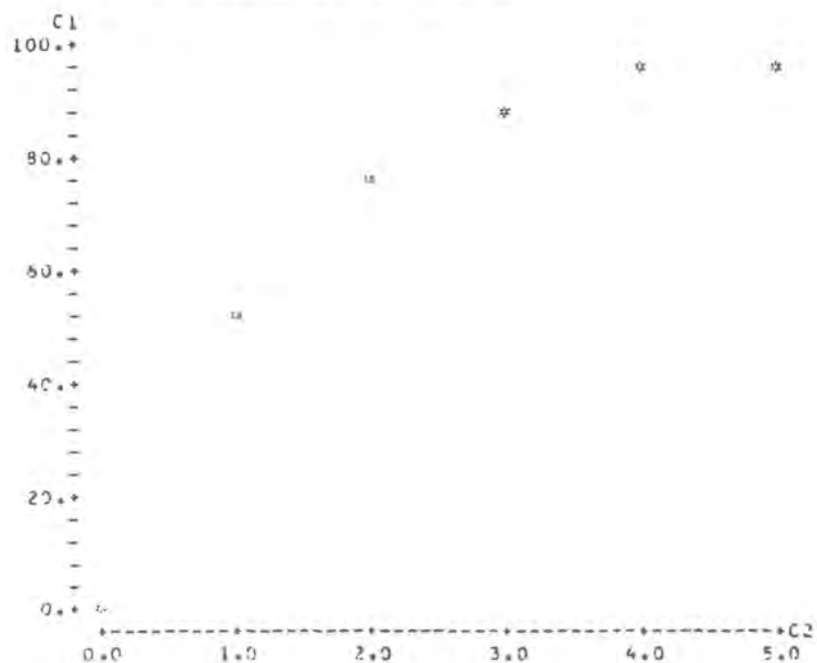


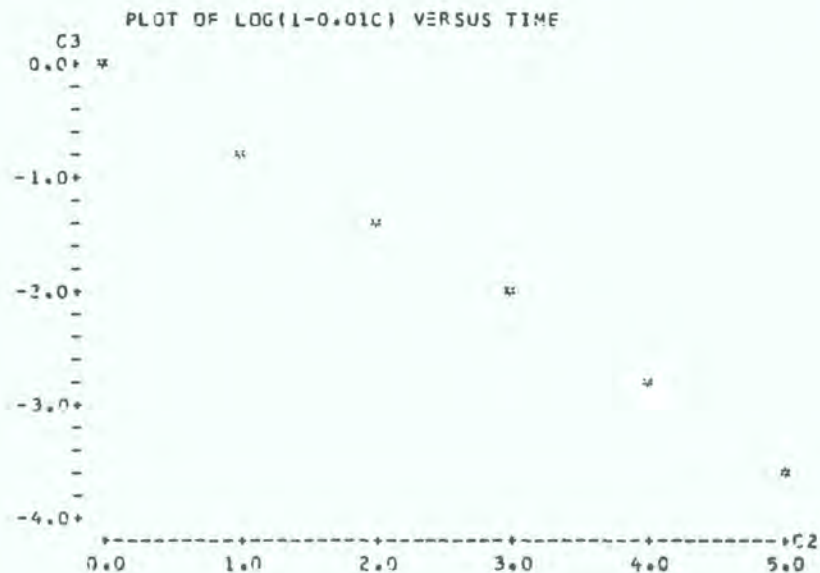
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LINEAR REGRESSION OF GENERATION OF COMPOUND C FROM COMPOUND A ON TIME  
 MODEL:  $Y = K(1 - \exp(-RT))$   $K = 100$   
 OR  $\text{LOG}(1 - 0.01C) = A + RT$

COLU#N	C1	C2	C3
COUNT	6	6	6
ROW	C	T	LOG(1-0.01C)
1	0.	0.	0.0
2	51.	1.	-0.71335
3	75.	2.	-1.39629
4	87.	3.	-2.04022
5	94.	4.	-2.81341
6	97.	5.	-3.50655

PLLOT OF LOG(1-0.01C) VERSUS TIME





--

THE REGRESSION EQUATION IS  
 $\text{LOG}(1-0.01C) = 0.0058 - 0.700 T$

COLUMN	COEFFICIENT	ST. DEV. OF COEF.	T-RATIO = COEF/S.D.
INTERCEPT	0.00576	0.02257	0.25
SLOPE	-0.697524	0.007456	-93.84

THE ST. DEV. OF Y ABOUT REGRESSION LINE IS  
 $S = 0.03119$   
 WITH  $(n - 2) = 4$  DEGREES OF FREEDOM

R-SQUARED = 100.0 PERCENT  
 R-SQUARED = 99.9 PERCENT, ADJUSTED FOR D.F.

ANALYSIS OF VARIANCE

DUE TO	DF	SS	MS=SS/DF
REGRESSION	1	8.565739	8.565739
RESIDUAL	4	0.003191	0.0007973
TOTAL	5	8.568930	

BURDIN-WATSON STATISTIC = 2.17

```

READ C1-C2
  6 2
 18 3
 55 4
 78 5
 95 6
 77 7
LET C3=LOG5((100-C1)/C1)
NOTE C3 IS THE LOG5((100-X)/X); X=C1
PRINT C1-C3
PLOT C1 VS C2
PLOT C3 VS C2
REGR C3 1 C2
SET C4
0 0.5 1.0 1.5 2.0 2.5 3.0 3.5 4.0 4.5 5.0 5.5 6.0 6.5 7.0 7.5 8.0
LET C5=100/(1+EXP(1-1.30244)*(C4-4.0367)))
WIDTH 100, 50
MPLOT C1 VS C2 AND C5 VS C4
STOP

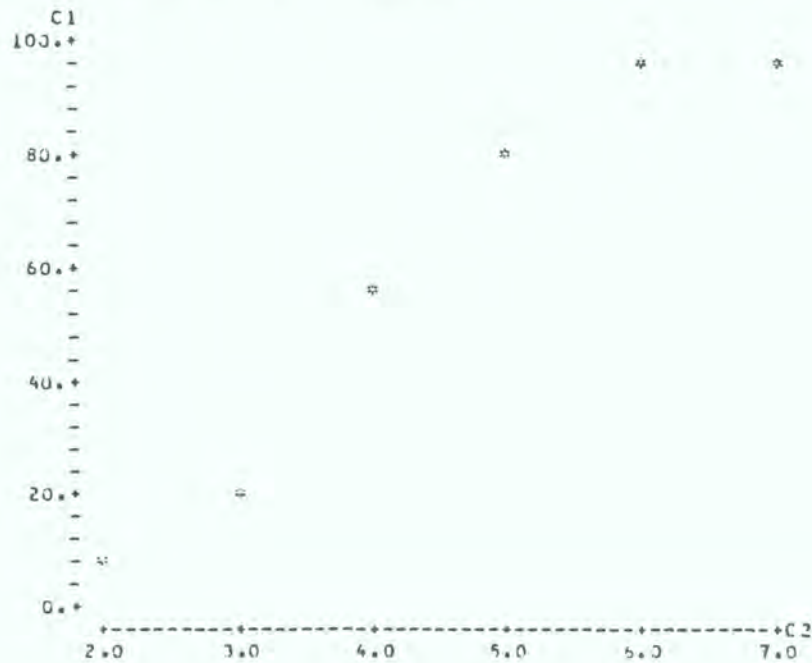
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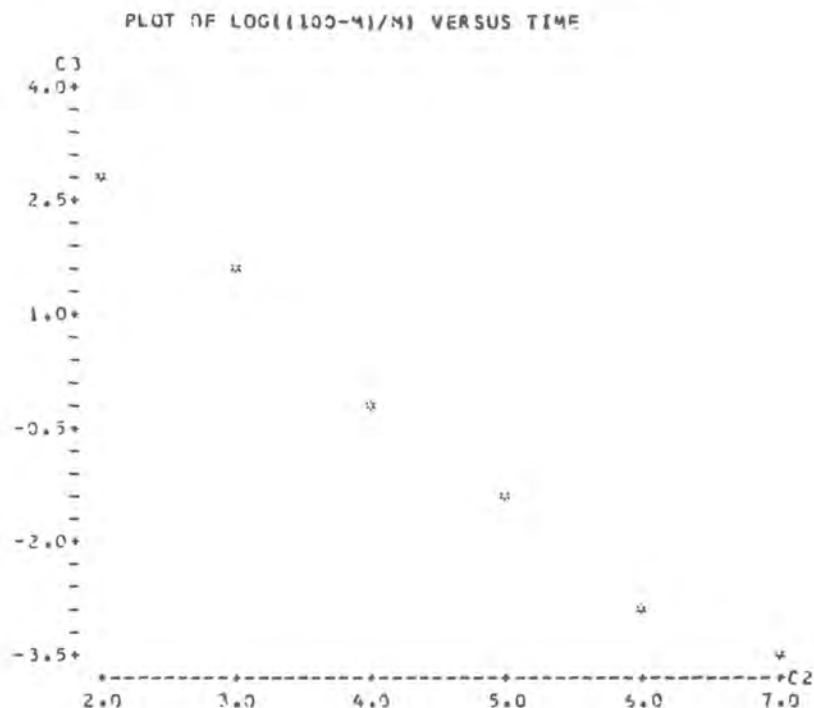
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LINEAR REGRESSION ANALYSIS OF % MORTALITY ON TIME  
 MODEL:  $\text{LOG}((K-M)/M) = A + BT$  K = 100

ROW	C1	C2	C3
COUNT	b	b	b
1	6.	2.	3.75154
2	18.	3.	1.51635
3	55.	4.	-0.20067
4	78.	5.	-1.26567
5	75.	6.	-2.74444
6	97.	7.	-3.47610

---  
 PLOT OF % MORTALITY VERSUS TIME





THE REGRESSION EQUATION IS

$$\text{LOG}((K-M)/M) = 5.26 - 1.30 T$$

COLUMN	COEFFICIENT	ST. DEV. OF COEF.	T-RATIO = COEF/S.D.
INTERCEPT	5.2579	0.3621	14.52
SLOPE	-1.30244	0.07524	-17.31

THE ST. DEV. OF Y ABOUT REGRESSION LINE IS

$$S = 0.3148$$

WITH  $(6 - 2) = 4$  DEGREES OF FREEDOM

R-SQUARED = 98.7 PERCENT

K-SQUARED = 98.4 PERCENT, ADJUSTED FOR D.F.

ANALYSIS OF VARIANCE

SOURCE	DF	SS	MS=SS/DF
REGRESSION	1	29.63520	29.63520
RESIDUAL	4	0.39623	0.09907
TOTAL	5	30.03143	

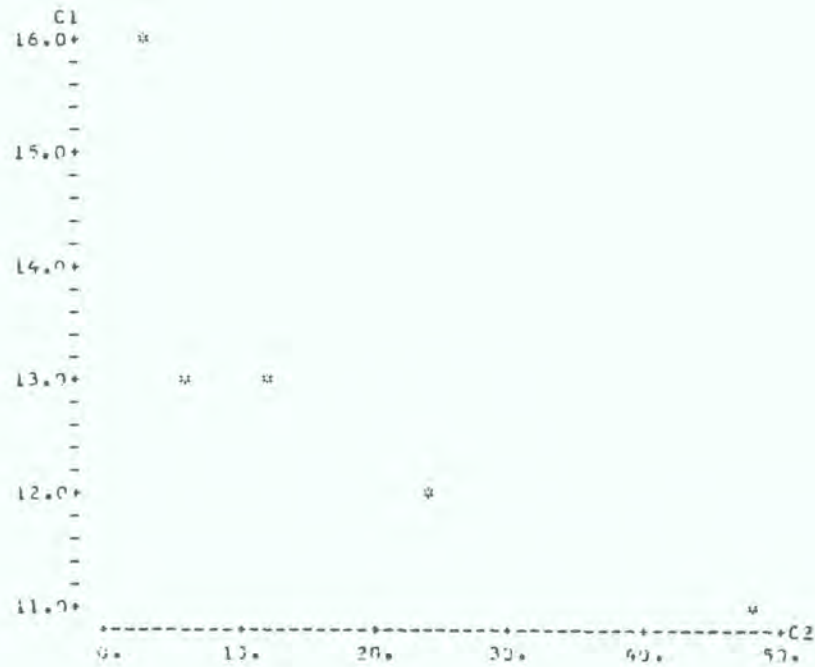
DURBIN-WATSON STATISTIC = 2.44

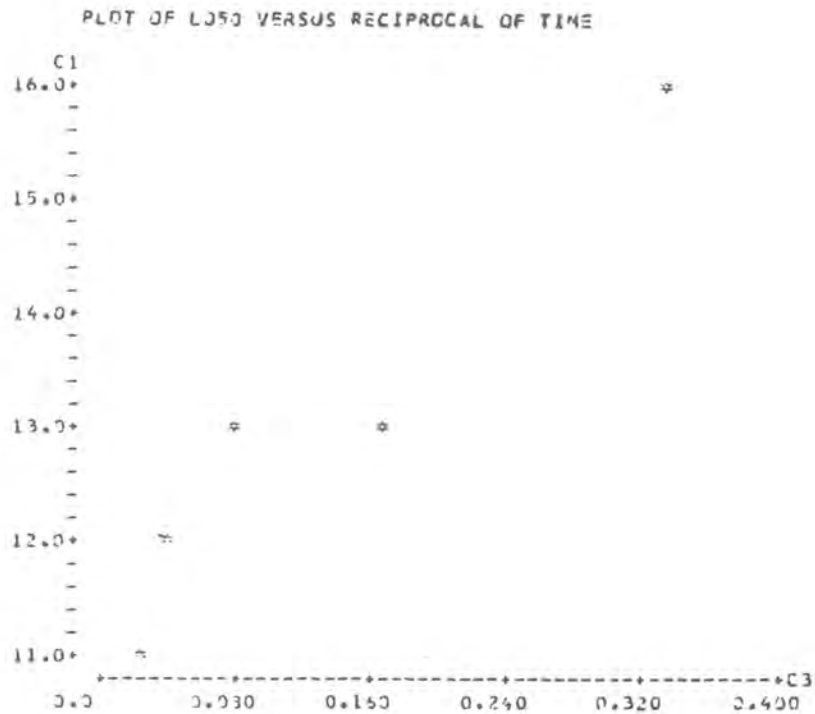
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LINEAR REGRESSION OF L050 ON RECIPROCAL OF TIME  
 MODEL: L050 = A + B/T

```
--
COLUMN      C1          C2          C3
COUNT      5          5          5
ROW
 1          15.         7.         0.333333
 2          13.         6.         0.166667
 3          13.         17.        0.063333
 4          12.         24.        0.041667
 5          11.         48.        0.020833
--
```

PLOT OF L050 VERSUS TIME





THE REGRESSION EQUATION IS

$$Y = 11.2 + 14.2 (1/T)$$

COLUMN	COEFFICIENT	ST. DEV. OF COEF.	T-RATIO = COEF/S.D.
INTERCEPT	11.1657	0.3939	28.72
SLOPE	14.194	2.250	6.25

THE ST. DEV. OF Y ABOUT REGRESSION LINE IS

$$S = 0.2742$$

WITH ( 5 - 2 ) = 3 DEGREES OF FREEDOM

R-SQUARED = 92.9 PERCENT

R-SQUARED = 70.6 PERCENT, ADJUSTED FOR D.F.

ANALYSIS OF VARIANCE

DF	SS	MS=SS/DF
REGRESSION 1	13.0107	13.0107
RESIDUAL 3	0.9392	0.3297
TOTAL 4	14.0000	

1 0.333 16.000 15.093 0.528 0.102 0.45 X

DURBIN-WATSON STATISTIC = 2.49

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EXAMPLE OF PROBIT ANALYSIS USING PROC PROBIT IN SAS  
 PROBIT ANALYSIS ON DOSE

16:31 TUESDAY, APRIL 30, 1985 1

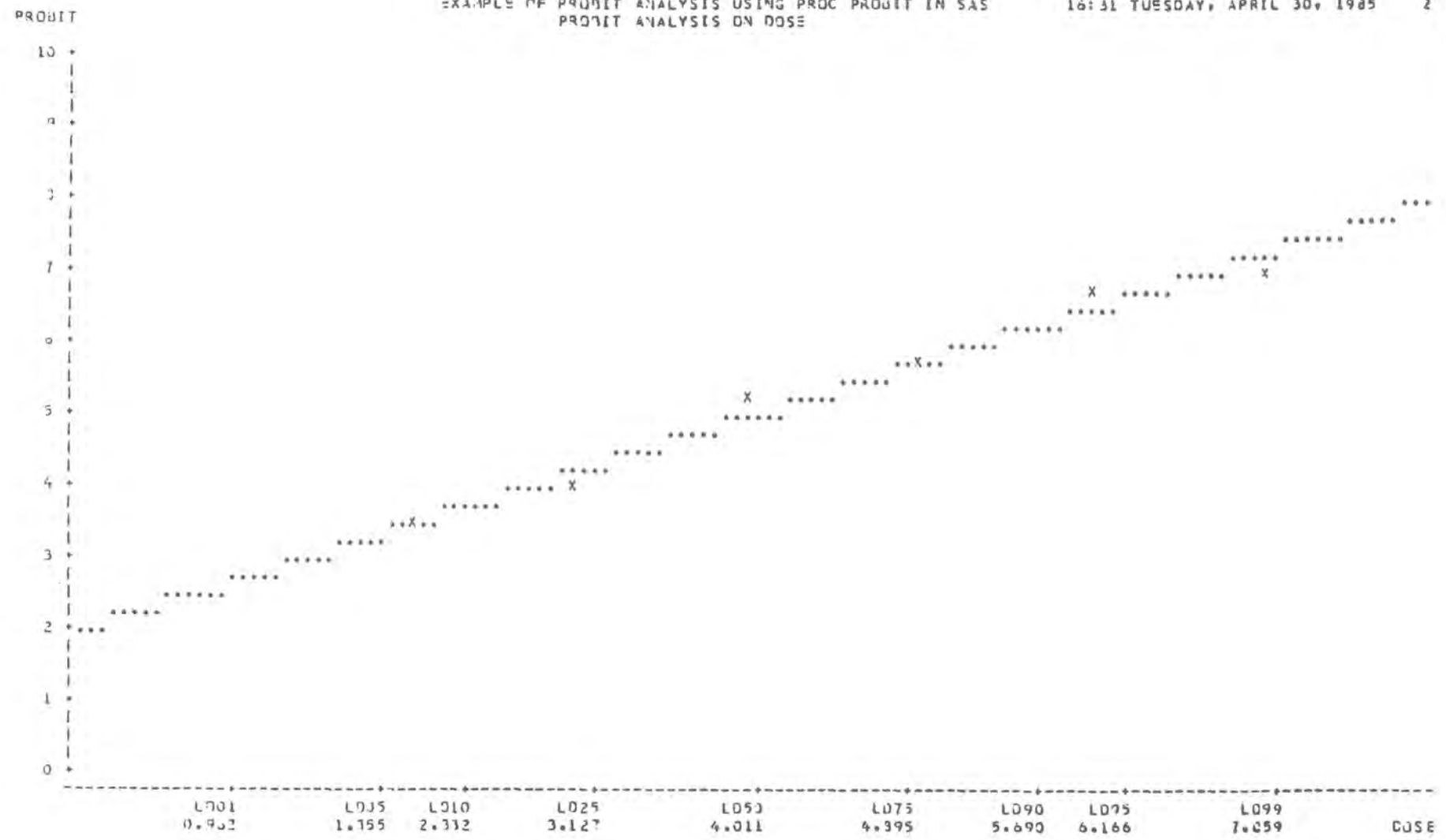
ITERATION	INTERCEPT	SLOPE	MU	SIGMA
0	2.04634315	0.72371477	4.05324025	1.37227965
1	1.94459905	0.75154909	4.01204350	1.31311195
2	1.93915063	0.76314104	4.01045673	1.31037377
3	1.93713975	0.76314476	4.01025433	1.31036858

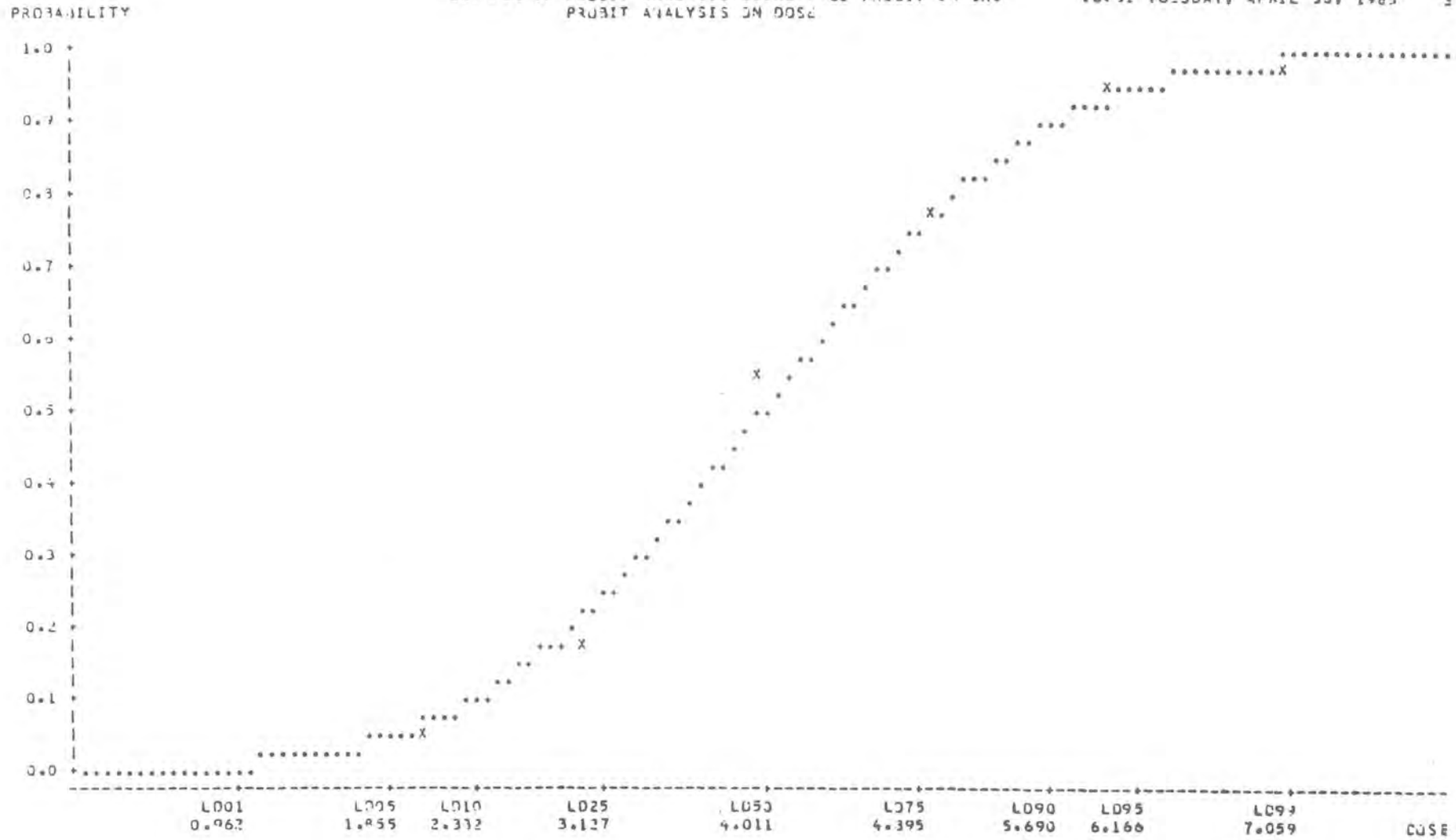
COVARIANCE MATRIX

	INTERCEPT	SLOPE		MU	SIGMA
INTERCEPT	0.04960371	-0.01059361	MU	0.00758349	-0.00069754
SLOPE	-0.01067363	0.00253384	SIGMA	-0.00069754	0.00753283

CHI-SQ = 5.6015 WITH 4 DF PROB > CHI-SQ = 0.2309

NOTE: SINCE THE CHI-SQUARE IS SMALL ( $P > 0.10$ ), FIDUCIAL LIMITS WILL BE COMPUTED USING A T VALUE OF 1.93.





EXAMPLE OF PROBIT ANALYSIS USING PROC PROBIT IN SAS  
 PROBIT ANALYSIS ON DISE

15:31 TUESDAY, APRIL 30, 1985 4

PROBABILITY	DOSE	95 PERCENT CONFIDENCE LIMITS	
		LOWER	UPPER
0.01	0.95743158	0.45225931	1.36262082
0.02	1.71783380	0.45922077	1.46244036
0.03	1.54532179	1.11694124	1.48590538
0.04	1.71591792	1.31050499	2.03924393
0.05	1.39549033	1.45772272	2.16420334
0.06	1.77352337	1.50135134	2.27075232
0.07	2.07702465	1.71335809	2.36433484
0.08	2.16767321	1.32294150	2.44326844
0.09	2.25197157	1.71300435	2.52473095
0.10	2.73154713	2.00535476	2.59523261
0.15	2.55274509	2.36554371	2.08857464
0.20	2.90302032	2.54971401	3.12384613
0.25	3.12702466	2.49145391	3.32772371
0.30	3.32359519	3.10648032	3.51297508
0.35	3.50594330	3.33357645	3.53660425
0.40	3.57137675	3.48830643	3.45375074
0.45	3.34619215	3.65459539	4.01791535
0.50	4.01035483	3.33547125	4.19207782
0.55	4.17351751	4.00362409	4.34897154
0.60	4.34293291	4.17165921	4.52133199
0.65	4.51376866	4.34245073	4.70246589
0.70	4.69301279	4.51953258	4.89620881
0.75	4.89458500	4.70771040	5.10920936
0.80	5.11388835	4.71424937	5.34724756
0.85	5.369376458	5.15197036	5.52494847
0.90	5.60115973	5.44749756	5.98701615
0.91	5.76773809	5.51941432	6.07394816
0.92	5.85201645	5.59527212	6.16855521
0.93	5.94463502	5.67964035	6.27276360
0.94	6.04319129	5.77364099	6.38935234
0.95	6.16521933	5.80061742	5.52255739
0.96	6.30489834	6.00601055	6.67733802
0.97	6.47538767	6.15980932	6.87244043
0.98	6.70202236	6.36374303	7.12965257
0.99	7.05722778	6.63422119	7.53399566

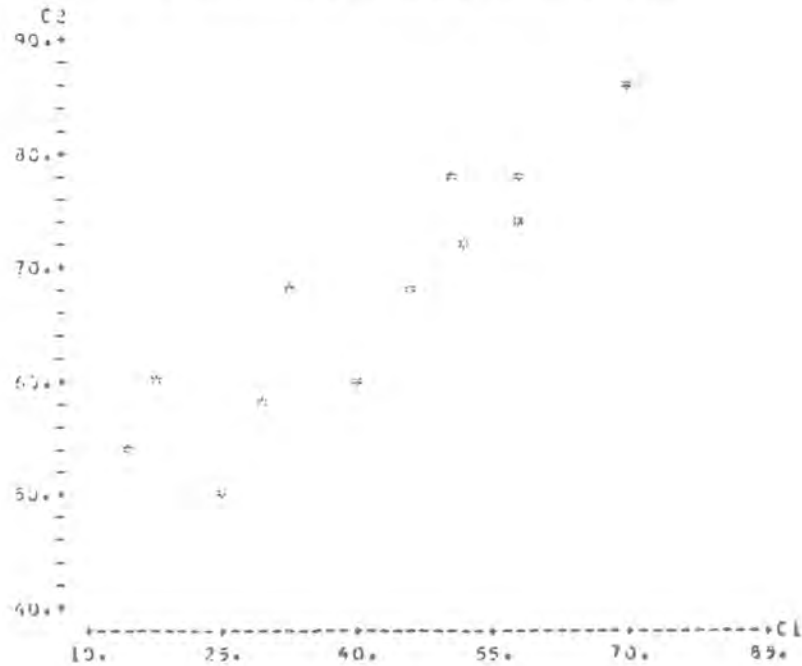
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 STORAGE AVAILABLE 4900

EXAMPLE OF USING WALFORD PLOT FOR OBTAINING THE  
 "VON BERTALANFFY" GROWTH EQUATION FOR SSS

```

--
COLUMN      C1      C2
COUNT      12      12
ROW         W AT 1983  W AT 1984
 1          14.     53.
 2          17.     59.
 3          25.     49.
 4          30.     57.
 5          32.     67.
 6          40.     59.
 7          46.     67.
 8          50.     78.
 9          52.     72.
10          58.     73.
11          58.     78.
12          70.     85.
    
```

--- PLOT OF HEIGHT IN 1984 VERSUS HEIGHT IN 1983



--

THE REGRESSION EQUATION IS

$$W = 43.2 + 0.571 W$$

COLUMN	COEFFICIENT	ST. DEV. OF COEF.	T-RATIO = COEF/S.D.
INTERCEPT	43.135	4.049	10.67
SLOPE	0.57059	0.07132	6.25

THE ST. DEV. OF Y ABOUT REGRESSION LINE IS

$$S = 5.333$$

WITH ( 12- 2) = 10 DEGREES OF FREEDOM

R-SQUARED = 70.6 PERCENT

R-SQUARED = 77.6 PERCENT, ADJUSTED FOR D.F.

ANALYSIS OF VARIANCE

DUE TO	DF	SS	MS=SS/DF
REGRESSION	1	1110.54	1110.54
RESIDUAL	10	234.33	23.44
TOTAL	11	1374.82	

DURBIN-WATSON STATISTIC = 2.51

--

ITERATION 1    Y =    0.554054 X +    43.301804

ITERATION 2    Y =    0.500433 X +    42.356764

SLOPE =    0.6055

LEVEL =    42.2453

--

--

THE REGRESSION EQUATION IS

$$Y = 51.7 + 0.0507 X1 + 0.0064 X2$$

COLUMN	COEFFICIENT	ST. DEV. OF COEF.	T-RATIO = COEF/S.D.
--	51.976	8.609	6.03
X1    C1	0.0507	0.4542	0.11
X2    C3	0.006416	0.005627	1.14

THE ST. DEV. OF Y ABOUT REGRESSION LINE IS

$$S = 5.254$$

WITH ( 12- 3) = 9 DEGREES OF FREEDOM

R-SQUARED = 32.2 PERCENT

R-SQUARED = 73.2 PERCENT, ADJUSTED FOR D.F.

ANALYSIS OF VARIANCE

SOURCE	DF	SS	MS=SS/DF
REGRESSION	2	1146.52	573.26
RESIDUAL	7	247.43	27.50
TOTAL	11	1394.92	

FURTHER ANALYSIS OF VARIANCE

SS EXPLAINED BY EACH VARIABLE (WHEN ENTERED IN THE ORDER GIVEN)

SOURCE	DF	SS
REGRESSION	2	1146.52
C1	1	1110.54
C3	1	35.93

ROW	X1	Y	PRED. Y	ST. DEV.	RESIDUAL	ST. RES.
12	70.3	75.30	75.84	3.45	-0.54	-0.32 X

X DENOTES AN OBS. WHOSE X VALUE GIVES IT LARGE INFLUENCE.

DURBIN-WATSON STATISTIC = 2.72

--  
--

```

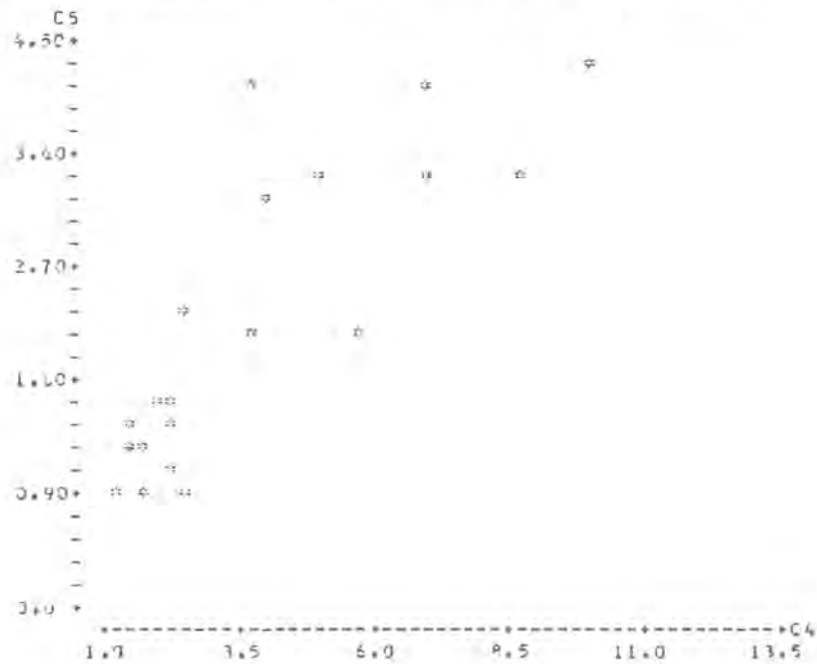
SET C4
1.75 2.59 4.03 3.73 1.99 1.28 4.11 17.11 2.24 2.39 7.04 5.09 2.29
2.79 2.15 1.72 5.75 1.52 3.97 1.42
SET C5
1.17 2.42 3.47 2.14 1.63 0.89 3.29 4.26 1.33 0.73 4.06 3.34 1.15
3.45 1.67 0.83 2.25 1.31 4.12 1.51
LET C6=C5/C4
PLOT C5 VS C4
LOGE C4, C1
LOGE C5, C3
PLOT C3 VS C1
REGR C3 1 C1, C7, C8
PLOT C7 VS C9
*PLOT C3 VS C1 AND C9 VS C1
LET K1=0.7413
LET K2=2.101
LET K3=0.1035
LET K4=K1-K2**K3
LET K5=K1+K2**K3
PRINT K1 K4 K5
LET K1=-0.1675
LET K3=0.1349
LET K4=K1-K2**K3
LET K5=K1+K2**K3
PRINT K1 K4 K5
EXP0 K1, K1
EXP0 K4, K4
EXP0 K5, K5
PRINT K1 K4 K5
GENE 1, 0.2, 13, C9
LET C10=C.9441*(C7**-(0.7413-1))
*PLOT C6 VS C4 AND C10 VS C9
STOP
    
```



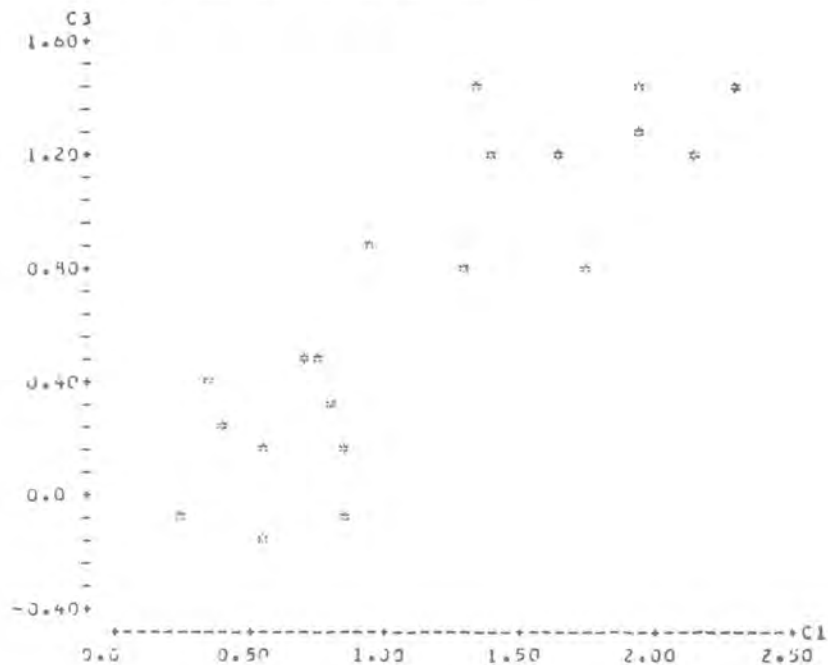
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EXAMPLE OF LINEAR REGRESSION ANALYSIS FOR RATIO VARIABLES  
THE HERMIT CRAB PROBLEM

---  
---  
PLOT OF WEIGHT OF SHELL (WS) VERSUS WEIGHT OF CRAB (WC)



PLOT OF LOG(WS) VERSUS LOG(WC)



THE REGRESSION EQUATION IS  
 $\text{LOG(WS)} = -0.171 + 0.743 \text{ LOG(WC)}$

COLUMN	COEFFICIENT	ST. DEV. OF COEF.	T-RATIO = COEF/S.D.
INTERCEPT	-0.1707	0.1357	-1.26
SLOPE	0.7425	0.1039	7.15

THE ST. DEV. OF Y ABOUT REGRESSION LINE IS  
 $S = 0.2879$   
 WITH  $(20 - 2) = 18$  DEGREES OF FREEDOM

R-SQUARED = 14.0 PERCENT  
 R-SQUARED = 72.5 PERCENT, ADJUSTED FOR D.F.

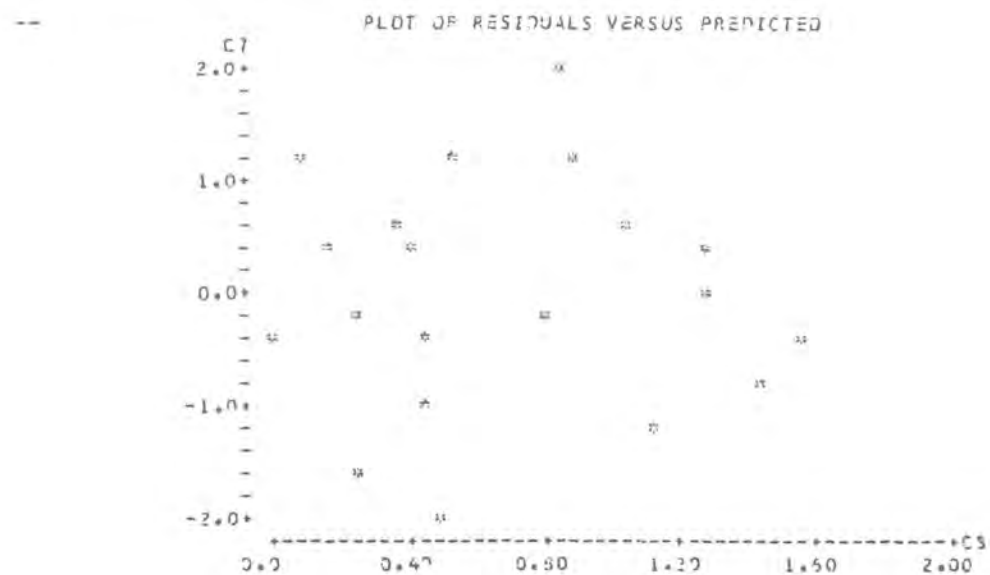
ANALYSIS OF VARIANCE

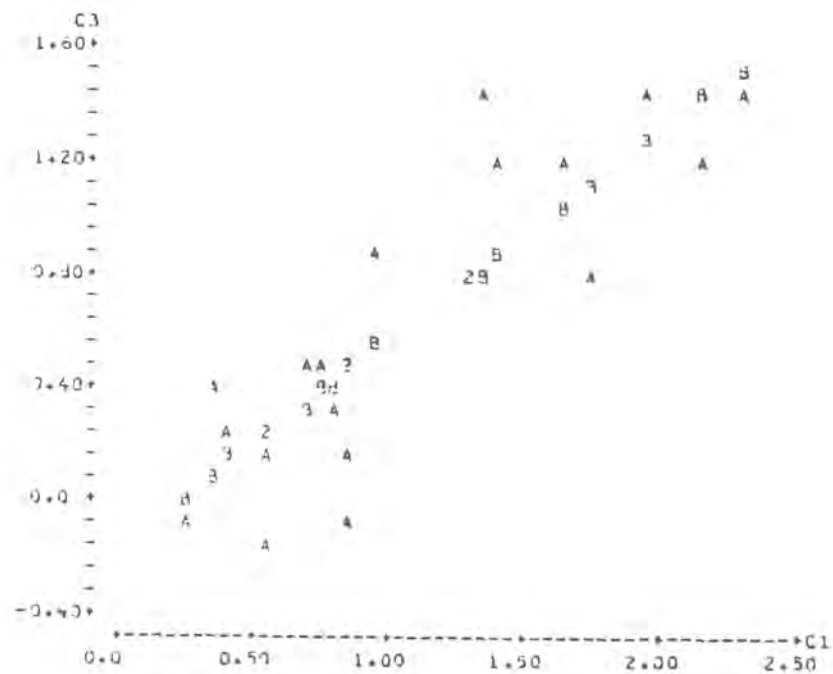
SOURCE	DF	SS	MS=SS/DF
REGRESSION	1	4.29798	4.29798
RESIDUAL	18	1.51276	0.08404
TOTAL	19	5.81073	

ROW	X1 C1	Y C3	PRED. Y VALUE	ST.DEV. PRED. Y	RESIDUAL	ST.RES. X
9	2.31	1.4493	1.5472	0.1377	-0.0979	-0.38 X
19	1.35	1.4159	0.9342	0.0684	0.5417	2.06R

R DENOTES AN OBS. WITH A LARGE ST. RES.  
 X DENOTES AN OBS. WHOSE X VALUE GIVES IT LARGE INFLUENCE.

DURBIN-WATSON STATISTIC = 1.63





--  
 K1 0.741900  
 K4 0.524347  
 K5 0.954253  
 --

--  
 K1 -0.160500  
 K4 -0.452725  
 K5 0.113925  
 --

--  
 ANSWER = 0.8441  
 --  
 ANSWER = 0.5359  
 --  
 ANSWER = 1.1207  
 --

--  
 K1 0.844087  
 K4 0.535765  
 K5 1.12057  
 --  
 --

## 5. ANALYSIS OF COVARIANCE

### 5.1 Introduction:

It originated as a procedure for improving an analysis of variance (to test for differences among different groups, or treatments) by correcting for an uncontrollable variable which is influencing the variable of interest. Here we will consider it as a method for comparing bivariate relationships among different groups, or treatments. For those interested in a reference, the new (7th) edition of Snedecor and Cochran's "Statistical Methods" has good coverage (Chapter 18).

Let us illustrate the analysis of covariance model by an assigned problem based on Walford Plot data. The size measurements are lengths at consecutive annual winter rings in shells of clams living in the intertidal zone on Hudson Bay, Canada, at two tidal levels:

0m tide level

Lx	:	2.7	3.0	3.2	3.3	3.4	3.6	3.9	4.3	4.3	4.8	5.6	5.7	6.0	6.6
														6.9	7.8
Lx+1	:	5.8	6.3	5.4	5.1	6.1	6.4	5.9	7.9	8.1	7.5	8.9	8.5	9.2	9.0
														8.8	10.2

Lx	:	8.7	9.1	9.1	10.2	12.3
Lx+1	:	11.5	11.2	12.3	12.0	14.3

1.1 m tide level

Lx	:	3.5	3.7	3.8	4.0	4.1	4.2	4.4	4.5	4.7	4.8	4.9	5.0			
													5.1	5.6	6.0	6.5
Lx+1	:	7.9	8.8	9.2	8.6	7.9	9.5	8.8	8.6	9.6	8.3	8.2	9.3			
													10.6	10.0	10.9	11.0

Lx	:	7.3	7.7	9.0	9.4	11.2	12.1
Lx+1	:	11.5	11.1	11.8	12.3	14.1	13.9

We wish to plot the data, check the plot for linearity, estimate the Walford Plot regression models and the corresponding Von Bertalanffy models, plot the model, and test whether the slopes and intercepts of the Walford Plot models differ between the two tidal levels.

## 5.2 Assigned problem:

1. Use the SET command to put Lx values into C1, Lx+1 values into C2, and a 0/1 code into C3 to represent tidal level.

2. Use the command "PLOT C2 C1, C3" to produce a Walford Plot of the Lx+1 versus Lx data, with points labeled by tidal level. Check that the trend, for each tidal level, is linear.

3. MINITAB does not have an "analysis of covariance" procedure, but we can do the necessary calculations using the REGR command. For 2 groups it is especially easy. What we want to do is regress Lx+1 on 3 predictor variables: Lx in C1, the tidal level code in C3, and the product of Lx times the tidal level code which we can put into C4 by 'MULT C1 BY C3, C4". Now do "REGR IN C1 C3 C4, ST. RESID. IN C5, PRED. Y IN C6."

4. If the regression coefficient for C4 is significant (check the tvalue) then the slopes of the Walford Plot models differ between tidal levels, and a test of intercepts has no meaning. Growth rate decreases with age faster at one tidal level than at the other. Asymptotic sizes probably, though not necessarily, differ.

5. If the slopes were not significantly different (they should be) then you would proceed to test for differences in intercepts, by seeing if the regression coefficient for C3 is significant. If the intercepts differ then you have different initial growth rates but the same relative decrease in growth rate with age. Asymptotic sizes will differ. If intercepts do not differ, then there is no evidence that the growth curves differ between tidal levels.

For a SAS analysis of covariance run you would use PROC GLM as follows:

```

TITLE ;
DATA COLDCLAMS;
INPUT YX YXP1 TL;
CARDS;

2.7    5.8    0
3.0    6.3    0
⋮      ⋮      ⋮
3.5    7.9    1
3.7    8.8    1
⋮      ⋮      ⋮

PROC GLM; CLASS TL;                test of difference
MODEL YXP1 = TL YX YX*TL;          between slopes

PROC GLM; BY TL;                   estimates separate
MODEL YXP1 = YX;                   slopes

PROC GLM; CLASS TL;               estimates a common
MODEL YXP1 = TL YX;               slope and tests for
                                   intercepts

PROC PLOT;                          does Walford Plot,
PLOT YXP1*YX = TL;                 with points coded
                                   by tidal level

```

For an analysis of covariance with more than 2 groups, the SAS package is probably the easiest.

#### 5.4 Covariance analysis as an alternative to ratio variables

##### 5.4.1 Assignment

Question: Do frogs differ in % water content between spring & fall? Since frogs sampled in spring and fall will probably be a mixture of sizes, one must also ask whether % water content varies with size of frog.

water wt.

A typical approach:     Let  $Y_i = \frac{\text{-----}}{\text{total wt.}}$  for frog  $i$ .

Collect  $n_1$  spring frogs and  $n_2$  fall frogs, determine water wt. and total wt, and calculate  $Y_i$  for all  $n_1 + n_2$  frogs. Test against the null hypothesis  $H_0$ : "that the mean  $Y$  is the same for spring frogs and fall frogs", using a t-test or a 1-way ANOVA. This is a bad approach - difficult to test what you want, difficult to interpret, and based on a ratio variable.

The covariance analysis approach: Let  $Y =$  water wt. and  $X =$  dry wt. Deal with the question of possible variation of "% water content" with size of frog first. The null hypothesis  $H_0$  is that as frog size varies the water wt. portion changes at the same % rate as does the dry wt. portion. For example, if a 10 gram frog is 6 grams water and 4 grams dry wt. then a 12 gram frog will be about 7.2 grams water (a 20% increase) and 4.8 grams dry weight (also a 20% increase), and the % water content is 60% in both cases. The model is  $dY/Y = b dX/X$ , and  $b=1$  if  $H_0$  is true.

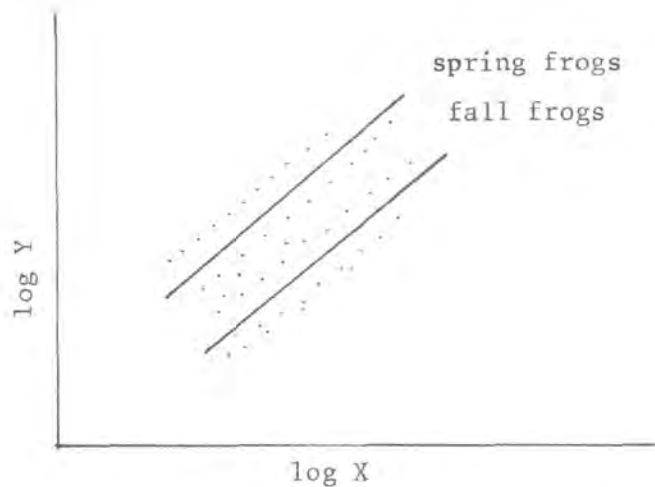
The nonlinear model is  $Y = AX^b$  which is  $Y = AX$  if  $H_0$  is true. Thus  $Y/X =$  constant value  $A$ .

If  $b=1$  then  $Y/X$  is not constant, but varies with size of frog.

If  $b < 1$  then big frogs have lower % water, and if  $b > 1$  they have higher % water. If spring and fall frogs of similar size have the same % water then  $A$  should be the same for both seasons.

The linear model is  $\log Y = a + b \log X$  (where  $a = \log A$ ). In covariance analysis one tests a sequence of hypotheses:





H1 : that the amount of variation about the regression lines is the same for both groups. If this is accepted, then

H2 : that the slopes (b-values) of the regression lines are the same. If this is accepted, then

H3 : that the common slope b is some specified value (e.g.,  $b=1$ )

H4 : that the intercepts (a-values) of the two common-slope regression lines are the same.

In this situation these hypotheses have the following biological interpretations:

H1 : that variation in % water content among frogs of similar size does not differ between spring and fall.

H2 : that any variation (or lack thereof) in % water content with size of frog is similar in spring and fall. If this is accepted, then

H3 : for the common slope  $b=1$ , that % water content does not vary with size of frog.

H4 : that the average % water content of frogs of similar size does not differ between spring and fall.  
Since  $a = \log A$ , different a-values reflect different  $Y/X$  ratios.

### 5.4.2. Job Listing and Output.

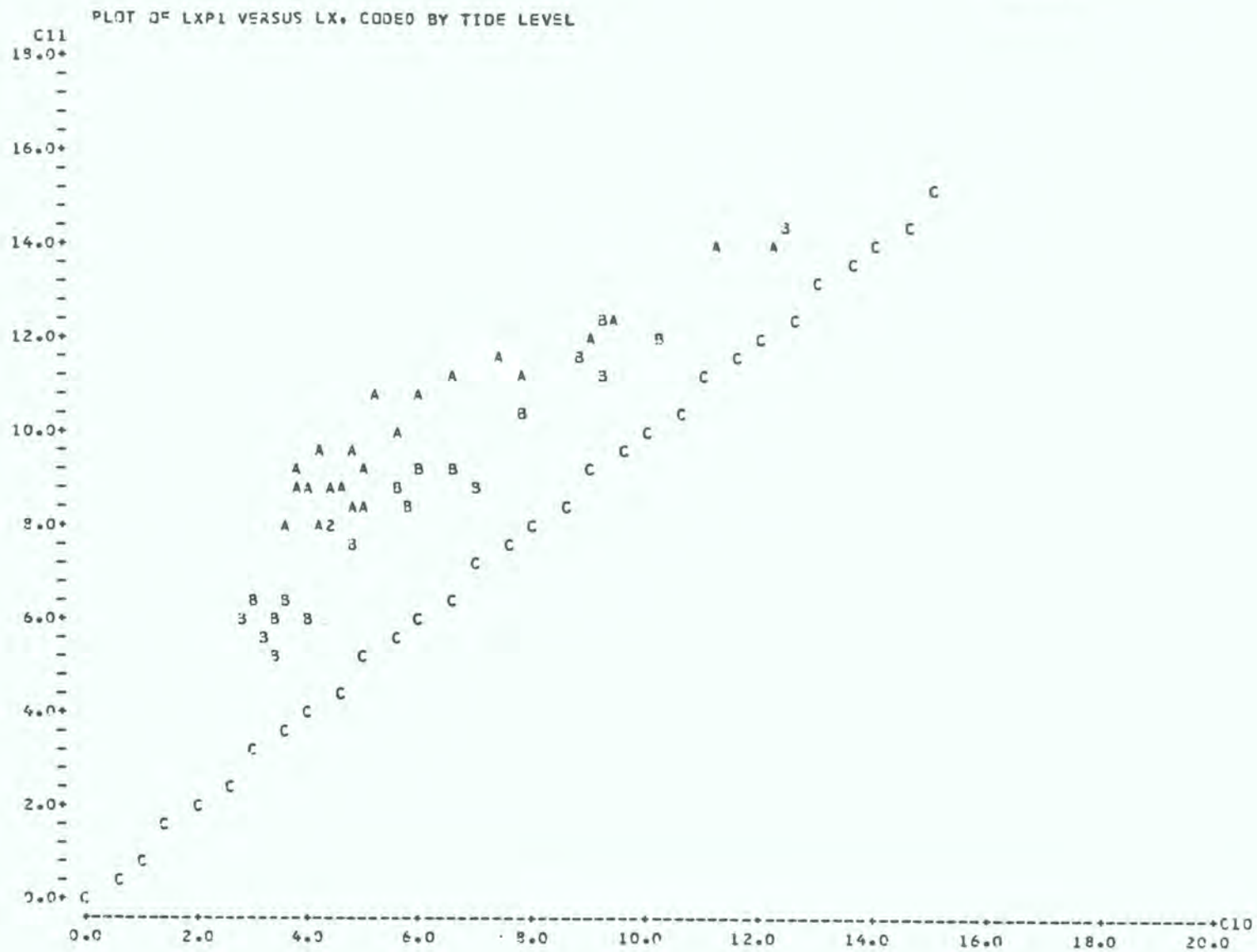
FILE: CCLAMJ MINITAB A1 VM/SP - CONVERSATIONAL MONITOR SYSTEM

PAGE 001

```
SET C1
2.7 3.0 3.2 3.3 3.4 3.5 3.7 4.3 4.3 4.9 5.5 5.7 6.0
5.6 6.9 7.9 9.7 9.1 10.2 12.3
3.5 3.7 3.9 4.0 4.1 4.2 4.4 4.5 4.7 4.8 4.9 5.7 5.1
5.6 6.0 6.5 7.3 7.7 7.9 7.4 11.2 12.1
SET C2
5.8 6.3 5.4 5.1 6.1 5.4 5.7 7.9 8.1 7.5 8.9 8.5 9.2
9.0 8.9 10.2 11.5 11.2 12.3 12.0 14.3
7.9 8.8 9.2 8.6 7.9 7.5 3.9 9.6 7.6 3.3 8.2 9.7 10.6
10.0 10.9 11.0 11.5 11.1 11.8 12.3 14.1 13.7
SET C3
Z1(0) Z2(1)
PRINT C1-C3
PICK 1 Z1 C1, C8
PICK 2 Z3 C1, C10
PICK 1 Z1 C2, C9
PICK 2 Z3 C2, C11
WIDTH 107, 50
GENE 0, 0.5, 15, C12
MPLOT C11 VS C10 AND C7 VS C8 AND C12 VS C12
MULT C1 BY C3, C4
REGR C2 3 C1 C3 C4, C5, C6
REGR C9 1 C8
REGR C11 1 C10
LET K1=LOGE(0.922)
LET K2=3.12/(1-0.922)
LET K3=LOGE(0.674)
LET K4=5.74/(1-0.674)
PRINT K1-K4
LET C7=K2*(1-EXP(K1*C9))
PICK 1 Z1 C6, C11
PICK 2 Z3, C6, C12
GENE 0, 0.5, 15, C14
MPLOT C14 VS C14 AND C11 VS C8 AND C12 VS C10
LET C9=K4*(1-EXP(K3*C10))
JOIN 0 C9 C7, C9
JOIN 0 C10 C3, C10
JOIN 3 C3, C3
LPLOT C7 C10, C3
STOP
```

FILE: CCLAM3 OUTPUT A1 VM/SP - CONVERSATIONAL MONITOR SYSTEM

COLUMN COUNT ROW	C1 43 LXPL	C2 43 LX	C3 43 TL
1	2.7000	5.8000	0.
2	3.0000	6.3000	0.
3	3.2000	5.4000	0.
4	3.3000	5.1000	0.
5	3.4000	5.1000	0.
6	3.6000	6.4000	0.
7	3.9000	5.9000	0.
8	4.3000	7.9000	0.
9	4.3000	8.1000	0.
10	4.8000	7.5000	0.
11	5.6000	8.9000	0.
12	5.7000	8.5000	0.
13	6.0000	8.2000	0.
14	6.6000	9.0000	0.
15	6.9000	8.8000	0.
16	7.9000	10.2000	0.
17	9.7000	11.5000	0.
18	9.1000	11.2000	0.
19	9.1000	12.3000	0.
20	10.2000	12.0000	0.
21	12.3000	14.3000	0.
22	3.5000	7.9000	1.
23	3.7000	8.8000	1.
24	3.3000	9.2000	1.
25	4.0000	3.6000	1.
26	4.1000	7.9000	1.
27	4.2000	9.5000	1.
28	4.4000	3.9000	1.
29	4.5000	3.6000	1.
30	4.7000	9.6000	1.
31	4.8000	9.3000	1.
32	4.9000	8.2000	1.
33	5.0000	9.3000	1.
34	5.1000	10.6000	1.
35	5.5000	10.0000	1.
36	6.0000	10.9000	1.
37	6.5000	11.0000	1.
38	7.3000	11.5000	1.
39	7.7000	11.1000	1.
40	9.0000	11.8000	1.
41	9.4000	12.3000	1.
42	11.2000	14.1000	1.
43	12.1000	13.9000	1.



THE REGRESSION EQUATION IS  
 $Y = 3.12 + 0.922 X1 + 2.81 X2 - 0.228 X3$

	COLUMN	COEFFICIENT	ST. DEV. OF COEF.	T-RATIO = COEF/S.D.
	INTERCEPT	3.1222	0.3218	9.70
X1	C1	0.92235	0.04955	18.62
X2	C3	2.8133	0.4703	5.98
X3	C4	-0.22791	0.07270	-3.13

THE ST. DEV. OF Y ABOUT REGRESSION LINE IS  
 $S = 0.6019$   
 WITH ( 43- 4) = 39 DEGREES OF FREEDOM

R-SQUARED = 93.7 PERCENT  
 R-SQUARED = 93.3 PERCENT, ADJUSTED FOR D.F.

ANALYSIS OF VARIANCE

DUE TO	DF	SS	MS=SS/DF
REGRESSION	3	211.2982	70.4327
RESIDUAL	39	14.1290	0.3623
TOTAL	42	225.4272	

FURTHER ANALYSIS OF VARIANCE  
 SS EXPLAINED BY EACH VARIABLE WHEN ENTERED IN THE ORDER GIVEN

DUE TO	DF	SS
REGRESSION	3	211.2982
C1	1	184.9595
C3	1	22.7785
C4	1	3.5602

ROW	X1	Y	PRED. Y	ST. DEV. PRED. Y	RESIDUAL	ST. RES.
21	12.3	14.3000	14.4572	0.3419	-0.1572	-0.34 X
42	11.2	14.1000	13.7132	0.3061	0.3868	0.75 X
43	12.1	13.9000	14.3382	0.3501	-0.4382	-0.90 X

X DENOTES AN OLS. WHOSE X VALUE GIVES IT LARGE INFLUENCE.

DURBIN-WATSON STATISTIC = 1.31

THE REGRESSION EQUATION (FOR TL=3 ONLY)  
 $Y = 3.12 + 0.922 X1$

	COLUMN	COEFFICIENT	ST. DEV. OF COEF.	T-RATIO = COEF/S.D.
	INTERCEPT	3.1222	0.3137	9.95
X1	C3	0.92236	0.04930	18.70

THE ST. DEV. OF Y ABOUT REGRESSION LINE IS  
 S = 0.5867  
 WITH ( 21- 2) = 19 DEGREES OF FREEDOM

R-SQUARED = 95.0 PERCENT  
 R-SQUARED = 94.8 PERCENT, ADJUSTED FOR D.F.

ANALYSIS OF VARIANCE

DUE TO	DF	SS	MS=SS/DF
REGRESSION	1	125.5396	125.5386
RESIDUAL	19	6.5399	0.3442
TOTAL	20	132.0786	

ROW	X1	Y	PRED. Y	ST.DEV.	RESIDUAL	ST.PES.
	C9	C9	VALUE	PRED. Y		
21	12.3	14.300	14.467	0.333	-0.167	-0.35 X

X DENOTES AN OBS. WHOSE X VALUE GIVES IT LARGE INFLUENCE.

DURBIN-WATSON STATISTIC = 1.89

THE REGRESSION EQUATION (FOR TL=1 ONLY)

Y = 5.94 + 0.694 X1

	COLUMN	COEFFICIENT	ST. DEV. OF COEF.	T-RATIO = COEF/S.D.
--	--	5.9355	0.3510	16.91
X1	C10	0.69444	0.05445	12.75

THE ST. DEV. OF Y ABOUT REGRESSION LINE IS  
 S = 0.0160  
 WITH ( 22- 2) = 20 DEGREES OF FREEDOM

R-SQUARED = 99.1 PERCENT  
 R-SQUARED = 98.5 PERCENT, ADJUSTED FOR D.F.

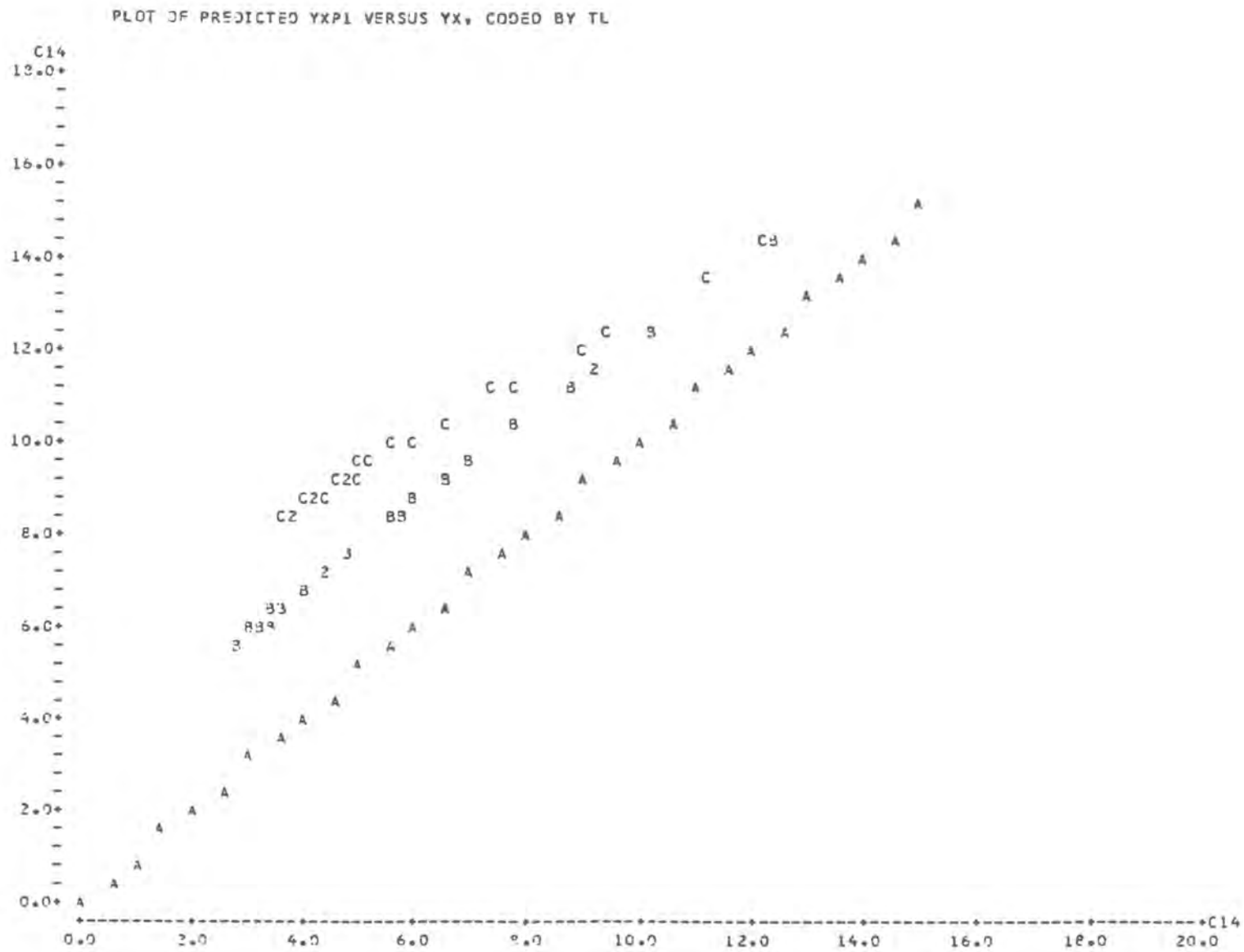
ANALYSIS OF VARIANCE

DUE TO	DF	SS	MS=SS/DF
REGRESSION	1	61.7174	61.7174
RESIDUAL	20	7.5390	0.3775
TOTAL	21	69.3064	

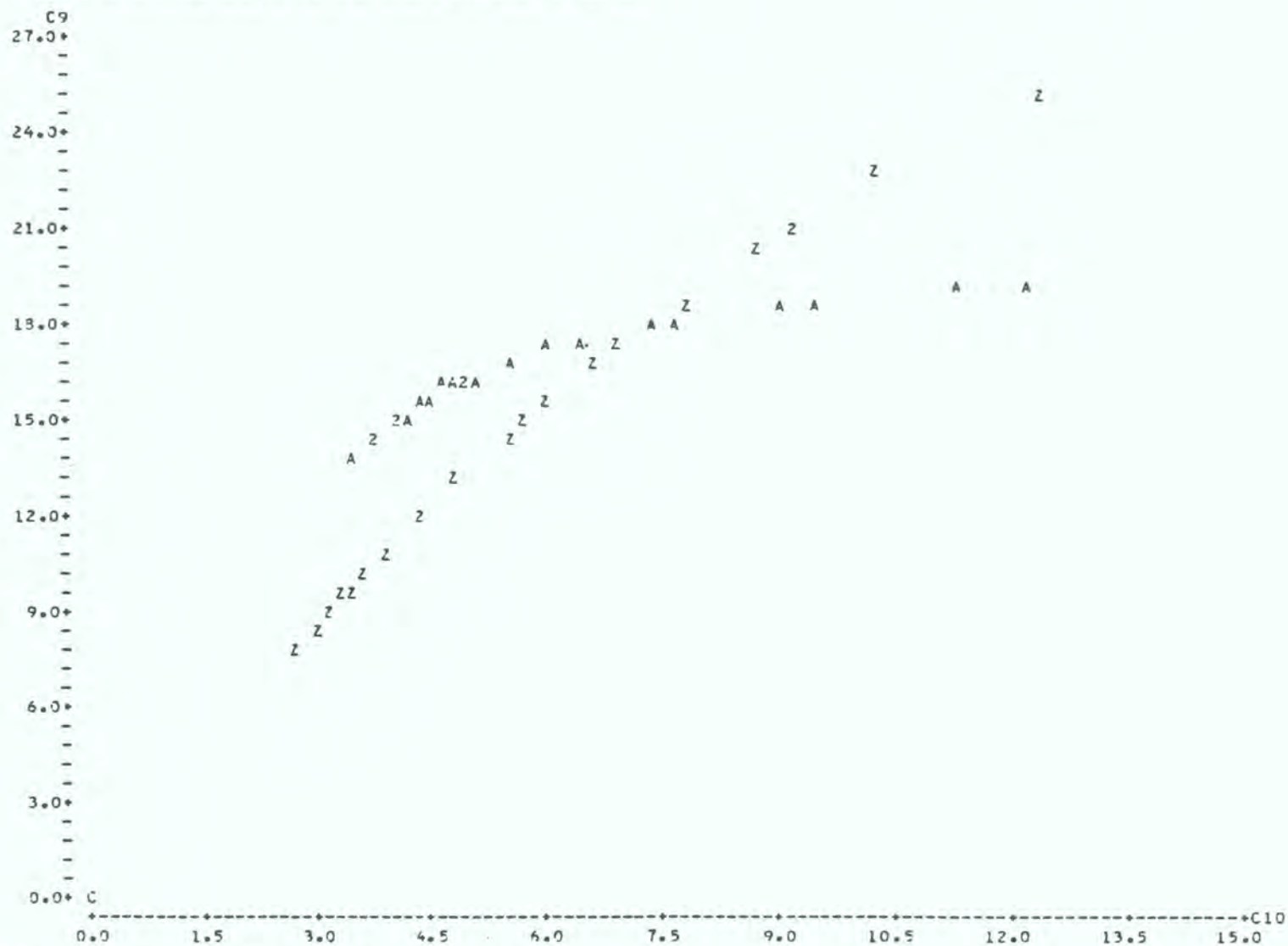
DURBIN-WATSON STATISTIC = 1.72

ESTIMATING THE GROWTH EQUATION (VON BERTALANFFY MODEL)  
 $L = K(1 - \exp(-kX))$

K1	-0.0812101	B FOR TL = 0
K2	40.0000	K FOR TL = 0
K3	-0.365283	B FOR TL = 1
K4	19.4113	K FOR TL = 1



PLOT OF ESTIMATED GROWTH CURVES (VON BERTALANFFY MODEL)





TITLE EXAMPLE OF ANALYSIS OF COVARIANCE USING SAS;  
 DATA COVDCLAM;  
 INPUT YX YXPI TL;  
 CARDS;

27.000	59.100	0.0
30.000	63.000	0.0
32.000	54.000	0.0
37.000	51.000	0.0
34.000	61.000	0.0
36.000	64.000	0.0
39.000	59.000	0.0
43.000	79.000	0.0
43.000	81.000	0.0
49.000	75.000	0.0
56.000	89.000	0.0
57.000	85.000	0.0
60.000	92.000	0.0
65.000	90.000	0.0
69.000	88.000	0.0
79.000	102.000	0.0
87.000	115.000	0.0
91.000	112.000	0.0
91.000	123.000	0.0
102.000	120.000	0.0
123.000	143.000	0.0
35.000	79.000	1.000
37.000	88.000	1.000
38.000	92.000	1.000
40.000	86.000	1.000
41.000	79.000	1.000
42.000	95.000	1.000
44.000	88.000	1.000
45.000	86.000	1.000
47.000	96.000	1.000
48.000	83.000	1.000
49.000	82.000	1.000
50.000	93.000	1.000
51.000	106.000	1.000
56.000	100.000	1.000
60.000	109.000	1.000
65.000	110.000	1.000
73.000	115.000	1.000
77.000	111.000	1.000
90.000	118.000	1.000
94.000	123.000	1.000
112.000	141.000	1.000
121.000	139.000	1.000

PROC GLM; CLASS TL;  
 MODEL YXPI = TL YX YX\*TL;  
 PROC GLM; BY TL;  
 MODEL YXPI = YX;  
 PROC GLM; CLASS TL;  
 MODEL YXPI = TL YX;  
 PROC PLOT;  
 PLOT YXPI\*YX = TL;

EXAMPLE OF ANALYSIS OF COVARIANCE USING SAS  
GENERAL LINEAR MODELS PROCEDURE

13:44 WEDNESDAY, MAY 8, 1985 2

DEPENDENT VARIABLE: YX\*1

SOURCE	DF	SUM OF SQUARES	MEAN SQUARE	F VALUE	PR > F	R-SQUARE	C.V.
MODEL	3	21129.70567092	7043.23522364	194.41	0.0001	0.937323	6.4334
ERROR	33	1412.49398024	36.22317393				YX*1 MEAN
CORRECTED TOTAL	42	22542.60465116				6.01898488	93.55813953

SOURCE	DF	TYPE I SS	F VALUE	PR > F	DF	TYPE III SS	F VALUE	PR > F
TL	1	2404.20421926	65.36	0.0001	1	1296.57011580	35.79	0.0001
YX	1	19369.46698953	507.05	0.0001	1	17915.85492905	494.53	0.0001
YX*TL	1	356.03446313	9.83	0.0033	1	356.03446313	9.83	0.0033

EXAMPLE OF ANALYSIS OF COVARIANCE USING SAS  
 TL=0

13:44 WEDNESDAY, MAY 8, 1985 3

GENERAL LINEAR MODELS PROCEDURE

DEPENDENT VARIABLE: YXP1

SOURCE	DF	SUM OF SQUARES	MEAN SQUARE	F VALUE	PR > F	R-SQUARE	C.V.
MODEL	1	12553.81321554	12553.81321554	364.72	0.0001	0.950434	6.8296
ERROR	19	653.99570327	34.42035933				YXP1 MEAN
CORRECTED TOTAL	20	13207.80891881			5.36692921		85.90476190

SOURCE	DF	TYPE I SS	F VALUE	PR > F	DF	TYPE III SS	F VALUE	PR > F
YX	1	12553.81321554	364.72	0.0001	1	12553.81321554	364.72	0.0001

PARAMETER	ESTIMATE	T FOR H0: PARAMETER=0	PR >  T	STD ERROR OF ESTIMATE
INTERCEPT	31.22213617	7.95	0.0001	3.13652418
YX	0.92235754	19.10	0.0001	0.04823724

GENERAL LINEAR MODEL PROCEDURE

DEPENDENT VARIABLE: YXP1

SOURCE	DF	SUM OF SQUARES	MEAN SQUARE	F VALUE	PR > F	R-SQUARE	C.V.
MODEL	1	6171.63323713	6171.63323713	162.65	0.0001	0.390500	6.1072
ERROR	20	753.97267197	37.69863360				YXP1 MEAN
CORRECTED TOTAL	21	6930.59090909			ROOT MSE 6.15993214		100.86363036

SOURCE	DF	TYPE I SS	F VALUE	PR > F	DF	TYPE III SS	F VALUE	PR > F
YX	1	6171.63323713	162.65	0.0001	1	6171.63323713	162.65	0.0001

PARAMETER	ESTIMATE	T FOR H0: PARAMETER=0	PR >  T	STD ERROR OF ESTIMATE
INTERCEPT	50.35520498	16.91	0.0001	3.50963941
YX	0.59443764	12.75	0.0001	0.25445143

EXAMPLE OF ANALYSIS OF COVARIANCE USING SAS  
GENERAL LINEAR MODELS PROCEDURE

CLASS LEVEL INFORMATION

CLASS	LEVELS	VALUES
TL	2	0 1

NUMBER OF OBSERVATIONS IN DATA SET = 43

EXAMPLE OF ANALYSIS OF COVARIANCE USING SAS  
 GENERAL LINEAR MODELS PROCEDURE

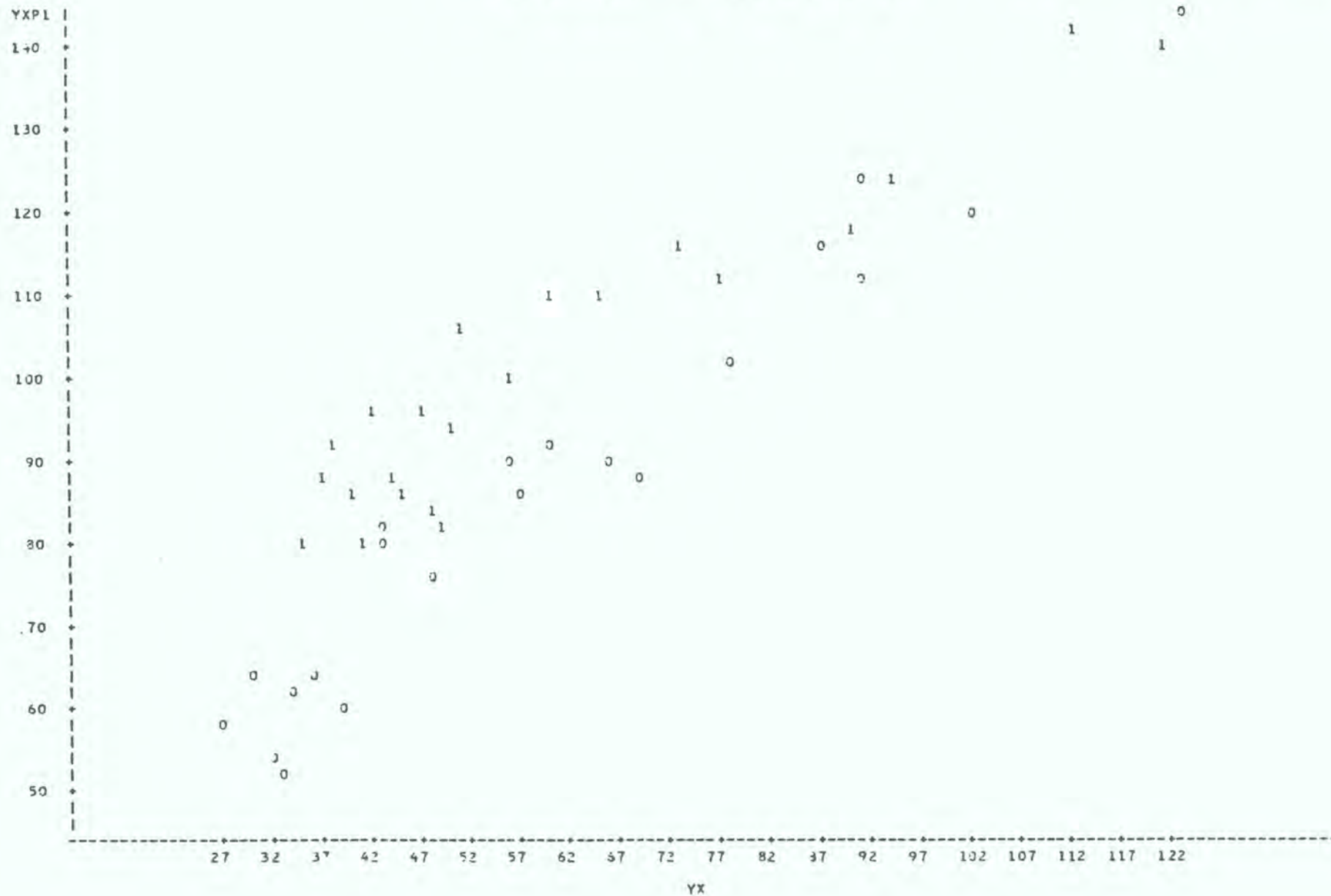
13:44 WEDNESDAY, MAY 8, 1985 6

DEPENDENT VARIABLE: YXP1

DF	SUM OF SQUARES	MEAN SQUARE	F VALUE	PR > F	R-SQUARE	C.V.	
3	20773.67129779	10336.83560390	234.37	0.0001	0.921529	7.1079	
40	1763.93344337	44.22333603				YXP1 MEAN	
42	22542.60465116					93.55813953	
						ROOT MSE	
						6.65005286	
DF	TYPE I SS	F VALUE	PR > F	DF	TYPE III SS	F VALUE	PR > F
1	2404.20421326	54.37	0.0001	1	2277.97274936	51.51	0.0001
1	19367.46593753	415.33	0.0001	1	18369.46698953	415.33	0.0001

EXAMPLE OF ANALYSIS OF COVARIANCE USING SAS  
 PLOT OF YXP1\*YX SYMBOL IS VALUE OF TL

13:44 WEDNESDAY, MAY 8, 1935 7



EXAMPLE OF ANALYSIS OF COVARIANCE USING SAS  
GENERAL LINEAR MODELS PROCEDURE

17:39 WEDNESDAY, MAY 8, 1985 1

CLASS LEVEL INFORMATION

CLASS	LEVELS	VALUES
TL	2	0 1

NUMBER OF OBSERVATIONS IN DATA SET = 22



EXAMPLE OF ANALYSIS OF COVARIANCE USING SAS  
 GENERAL LINEAR MODELS PROCEDURE

17:39 WEDNESDAY, MAY 8, 1985 2

DEPENDENT VARIABLE: YXP1

SOURCE	DF	SUM OF SQUARES	MEAN SQUARE	F VALUE	PR > F	R-SQUARE	C.V.
MODEL	3	12457.39328909	4152.63276270	82.86	0.0001	0.932477	7.5312
ERROR	18	902.10171191	50.11676177		ROOT MSE		YXP1 MEAN
CORRECTED TOTAL	21	13360.00000000			7.07931930		94.00000000

SOURCE	DF	TYPE I SS	F VALUE	PR > F	DF	TYPE III SS	F VALUE	PR > F
TL	1	716.54545455	18.29	0.0005	1	622.75395700	12.43	0.0024
YX	1	11385.47459721	227.20	0.0001	1	10307.39191266	205.67	0.0001
YX*TL	1	154.37123633	3.09	0.0958	1	154.87823633	3.09	0.0958

EXAMPLE OF ANALYSIS OF COVARIANCE USING SAS  
TL=0

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GENERAL LINEAR MODELS PROCEDURE

DEPENDENT VARIABLE: YXP1

SOURCE	DF	SUM OF SQUARES	MEAN SQUARE	F VALUE	PR > F	R-SQUARE	C.V.	
MODEL	1	8276.61566523	8276.61566523	189.97	0.0001	0.954767	7.5396	
ERROR	9	392.11160750	43.56795639				YXP1 MEAN	
CORRECTED TOTAL	10	8553.72727273			0.00060273		87.54545455	
SOURCE	DF	TYPE I SS	F VALUE	PR > F	DF	TYPE III SS	F VALUE	PR > F
YX	1	8276.61566523	189.97	0.0001	1	8276.61566523	189.97	0.0001
PARAMETER	ESTIMATE	T FOR H0: PARAMETER=0	PR >  T	STD ERROR OF ESTIMATE				
INTERCEPT	30.03149402	6.50	0.0001	4.62311759				
YX	0.95711583	13.78	0.0001	0.06944192				

GENERAL LINEAR MODELS PROCEDURE

DEPENDENT VARIABLE: YXPI

SOURCE	DF	SUM OF SQUARES	MEAN SQUARE	F VALUE	PR > F	R-SQUARE	C.V.
MODEL	1	3264.73716832	3264.73716832	57.61	0.0001	0.864894	7.4936
ERROR	9	509.99710441	56.66556716				YXPI MEAN
CORRECTED TOTAL	10	3774.72727273			7.52765350		100.45454545

SOURCE	DF	TYPE I SS	F VALUE	PR > F	DF	TYPE III SS	F VALUE	PR > F
YX	1	3264.73716832	57.61	0.0001	1	3264.73716832	57.61	0.0001

PARAMETER	ESTIMATE	T FOR H0: PARAMETER=0	PR >  T	STD ERROR OF ESTIMATE
INTERCEPT	56.92925043	9.23	0.0001	6.16710065
YX	0.74809101	7.59	0.0001	0.09355756

EXAMPLE OF ANALYSIS OF COVARIANCE USING SAS  
 GENERAL LINEAR MODELS PROCEDURE

17:39 WEDNESDAY, MAY 8, 1985 6

DEPENDENT VARIABLE: YXP1

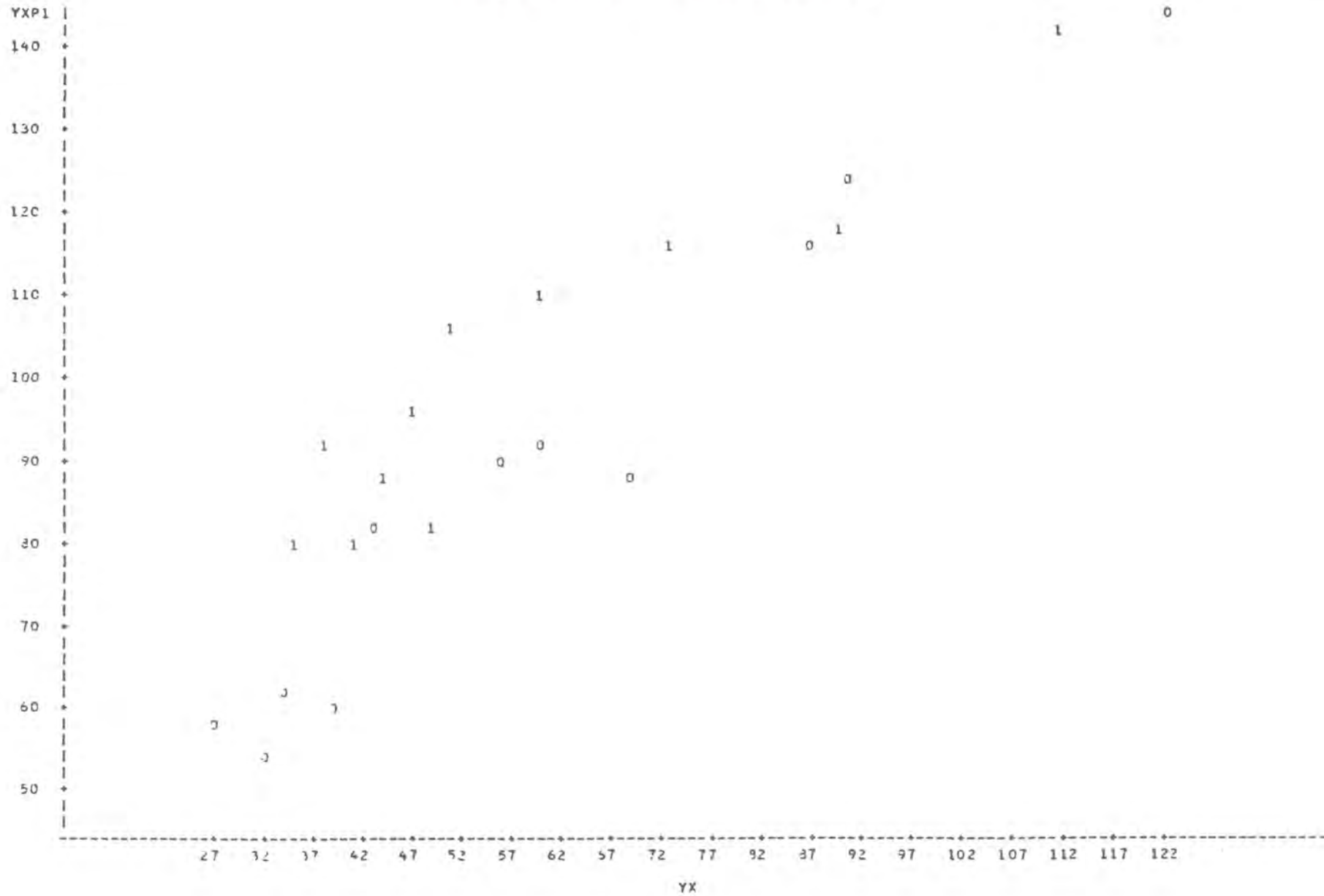
SOURCE	DF	SUM OF SQUARES	MEAN SQUARE	F VALUE	PR > F	R-SQUARE	C.V.
MODEL	2	12303.02005176	6151.51002588	110.58	0.0001	0.920885	7.9347
ERROR	19	1056.97994924	55.63052357				YXP1 MEAN
CORRECTED TOTAL	21	13360.00000000			7.45858724		94.00000000

SOURCE	DF	TYPE I SS	F VALUE	PR > F	DF	TYPE III SS	F VALUE	PR > F
TL	1	916.54545455	16.43	0.0007	1	1167.55551974	20.99	0.0002
YX	1	11386.4749721	204.68	0.0001	1	11386.4749721	204.68	0.0001

EXAMPLE OF ANALYSIS OF COVARIANCE USING SAS  
 PLOT OF YXP1\*YX SYMBOL IS VALUE OF TL

17:39 WEDNESDAY, MAY 8, 1985 7



COLUMN COUNT ROW	DRYWT 30	WATER 30	SEASON (F=1,S=2) 30
1	2.1300	1.6500	1.
2	3.0400	2.6300	1.
3	4.6900	4.9700	1.
4	2.3200	2.1000	1.
5	1.5600	1.3000	1.
6	4.4200	3.9900	1.
7	8.3900	8.6000	1.
8	5.5600	5.8300	1.
9	4.4600	3.7100	1.
10	11.4900	10.7100	1.
11	6.7800	7.5700	1.
12	3.1200	3.0500	1.
13	9.5500	9.4400	1.
14	8.7100	9.6100	1.
15	1.3100	1.0300	1.
16	2.8300	2.2200	2.
17	2.9800	2.0400	2.
18	3.1000	2.3600	2.
19	2.0400	1.3100	2.
20	1.5900	1.0500	2.
21	1.9300	1.5000	2.
22	1.5500	1.0700	2.
23	4.7500	3.5800	2.
24	8.4100	7.0800	2.
25	3.1200	2.1300	2.
26	2.2400	1.7700	2.
27	18.4300	16.0300	2.
28	11.2400	8.9300	2.
29	2.6600	1.5900	2.
30	1.2900	0.9200	2.

HERE IS THE "STATS ON A RATIO VARIABLE" APPROACH:

LET 'PCTWAT'=100 ('WATER'/'WATER'+'DRYWT')

COLUMN COUNT ROW	DRYWT 30	WATER 30	SEASON 30	PCTWAT 30
1	2.1300	1.6500	1.	43.6508
2	3.0400	2.6300	1.	46.3845
3	4.6900	4.9700	1.	52.5370
4	2.3200	2.1000	1.	47.5113
5	1.5600	1.3000	1.	45.4545
6	4.4200	3.9900	1.	47.4435
7	8.3900	8.6000	1.	50.6180
8	5.5600	5.8300	1.	51.1853
9	4.4600	3.7100	1.	45.4100
10	11.4900	10.7100	1.	48.2432
11	6.7800	7.5700	1.	52.7526

12	3.1200	3.0500	1.	49.4327
13	9.5500	9.4400	1.	49.7104
14	8.7100	9.6100	1.	52.4563
15	1.3100	1.0300	1.	44.0171
16	2.8100	2.2200	2.	43.9604
17	2.9900	2.0400	2.	40.6375
18	3.1700	2.3600	2.	43.2234
19	2.0400	1.3100	2.	39.1045
20	1.5400	1.0500	2.	39.7727
21	1.9300	1.5000	2.	43.7318
22	1.5500	1.0700	2.	40.8397
23	4.7500	3.5800	2.	42.9772
24	8.4100	7.0900	2.	45.7069
25	3.1200	2.1300	2.	40.5714
26	2.2400	1.7700	2.	44.1397
27	19.4300	16.0300	2.	46.5177
28	11.2400	8.9300	2.	44.2737
29	2.6600	1.5900	2.	37.4118
30	1.2900	0.9200	2.	41.6290

ONEWAY ANALYSIS OF VARIANCE, ON 'PCTWAT', BETWEEN SEASONS

ANALYSIS OF VARIANCE

DUE TO	DF	SS	MS=SS/DF	F-RATIO
SEASON	1	284.04	284.04	35.39
ERROR	29	224.73	8.03	
TOTAL	30	508.76		

SEASON	N	MEAN	ST. DEV.
FALL	15	48.45	3.08
SPRING	15	42.30	2.56

INDIVIDUAL 95 PERCENT C. I. FOR LEVEL MEANS  
(BASED ON POOLED STANDARD DEVIATION)

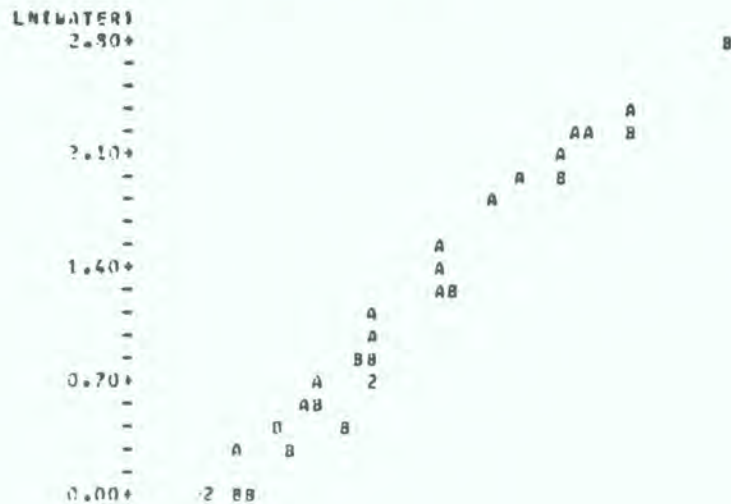


NOW HERE IS THE "LOG-LOG ANALYSIS OF COVARIANCE" APPROACH:

- \* LET 'DRYWT' = LOGE('DRYWT')
- \* LET 'WATER' = LOGE('WATER')

COLUMN	LN(DRYWT)	LN(WATER)	SEASON
COUNT	30	30	30
ROW			

1	0.75612	0.50078	1.
2	1.11196	0.96698	1.
3	1.50185	1.60342	1.
4	0.94157	0.74194	1.
5	0.44469	0.26236	1.
6	1.48614	1.38377	1.
7	2.12706	2.15176	1.
8	1.71567	1.76302	1.
9	1.49515	1.31103	1.
10	2.44148	2.37118	1.
11	1.71579	2.02419	1.
12	1.13733	1.11514	1.
13	2.25656	2.24496	1.
14	2.16447	2.26290	1.
15	0.27003	0.02956	1.
16	1.04723	0.79751	2.
17	1.09192	0.71295	2.
18	1.13147	0.85966	2.
19	0.71295	0.27003	2.
20	0.46575	0.04379	2.
21	0.65752	0.40547	2.
22	0.43325	0.06766	2.
23	1.55314	1.27536	2.
24	2.12962	1.95727	2.
25	1.13783	0.75612	2.
26	0.90649	0.57098	2.
27	2.91593	2.77446	2.
28	2.41749	2.18942	2.
29	0.97333	0.46573	2.
30	0.25464	-0.08338	2.







\* LET \*DRWT.SN\*='LN(DRYWT)'+\*SEASON\*

COLUMN COUNT ROW	LN(DRYWT) 30	LN(WATER) 30	SEASON 30	DRWT.SN 30
1	0.75612	0.50078	1.	0.75612
2	1.11186	0.96693	1.	1.11186
3	1.50185	1.60342	1.	1.50185
4	0.84157	0.74194	1.	0.84157
5	0.44469	0.26236	1.	0.44469
6	1.48614	1.38379	1.	1.48614
7	2.12704	2.15176	1.	2.12704
8	1.71560	1.76302	1.	1.71560
9	1.49515	1.31103	1.	1.49515
10	2.44148	2.37118	1.	2.44148
11	1.91399	2.02419	1.	1.91398
12	1.13783	1.11514	1.	1.13783
13	2.25654	2.24496	1.	2.25654
14	2.16447	2.26280	1.	2.16447
15	0.27003	0.02956	1.	0.27003
16	1.04028	0.79751	2.	2.08055
17	1.09192	0.71295	2.	2.18385
18	1.13140	0.85966	2.	2.26280
19	0.71795	0.27003	2.	1.42590
20	0.46373	0.04879	2.	0.92747
21	0.65752	0.40547	2.	1.31504
22	0.43825	0.06766	2.	0.87651
23	1.55814	1.27536	2.	3.11629
24	2.12942	1.95727	2.	4.25884
25	1.13783	0.75612	2.	2.27567
26	0.80648	0.57098	2.	1.61295
27	2.91398	2.77446	2.	5.82796
28	2.41948	2.18942	2.	4.83896
29	0.97933	0.46373	2.	1.75665
30	0.25464	-0.08339	2.	0.50928

-- REGRESS \*LN(WATER)\* ON 3 PREDICTORS \*LN(DRYWT)\* \*SEASON\* \*DRWT.SN\*

	COLUMN	COEFFICIENT	ST. DEV. OF COEF.	F-RATIO = COEF/S.D.
	--	-0.0811	0.1180	
X1	LN(DRYWT)	1.16970	0.07586	
X2	SEASON	-0.16729	0.06958	
X3	DRWT.SN	-0.04067	0.04625	-0.38 (P>0.05) NS

CONCLUSION: NO SLOPE DIFFERENCE BETWEEN SEASONS,  
WHICH MEANS THAT IF PERCENT WATER VARIES WITH  
SIZE OF FROG (I.E., WITH DRY WT.) THEN IT

DOES SO IN THE SAME MANNER IN BOTH SEASONS.

REGRESS LN(WATER) ON 2 PREDICTORS LN(DRYWT) \* SEASON \*

COLUMN	COEFFICIENT	ST. DEV. OF COEF.	T-RATIO = COEF/S.D.
X1 LN(DRYWT)	0.00609	0.06366	
X2 SEASON	1.10608	0.02284	4.64 (P<0.01) **
	-0.22140	0.03239	-6.33 (P<0.01) **

R-SQUARED = 97.0 PERCENT

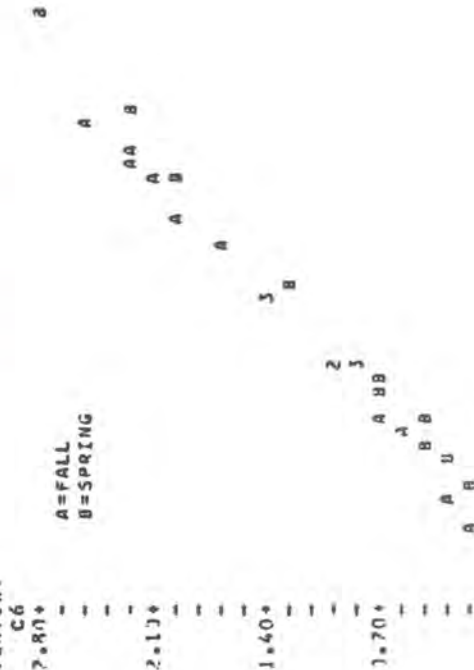
CONCLUSIONS: (1) FOR A GIVEN SIZE (I.E., DRY WEIGHT) FROG WATER CONTENT DIFFERS BETWEEN SEASONS. (2) SINCE B>1, PERCENT WATER CONTENT INCREASES WITH SIZE (I.E., DRY WEIGHT) OF FROG.

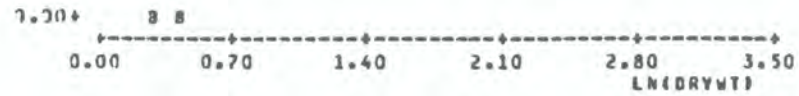
FALL: LN(WATER) = 0.00609 + 1.10608\*LN(DRYWT) - 0.2214\*(1)  
 = -0.21531 + 1.10608\*LN(DRYWT)  
 WATER = 0.8063\*(DRYWT)-1.10608

SPRING: LN(WATER) = 0.00609 + 1.10608\*LN(DRYWT) - 0.2214\*(2)  
 = -0.43671 + 1.10608\*LN(DRYWT)  
 WATER = 0.6462\*(DRYWT)-1.10608

ALINE LN(WATER), LN(DRYWT) / ROBUST REGRESSION ESTIMATES SLOPE VERY CLOSE TO SAME VALUE AS MODEL 1 REGRESSION.

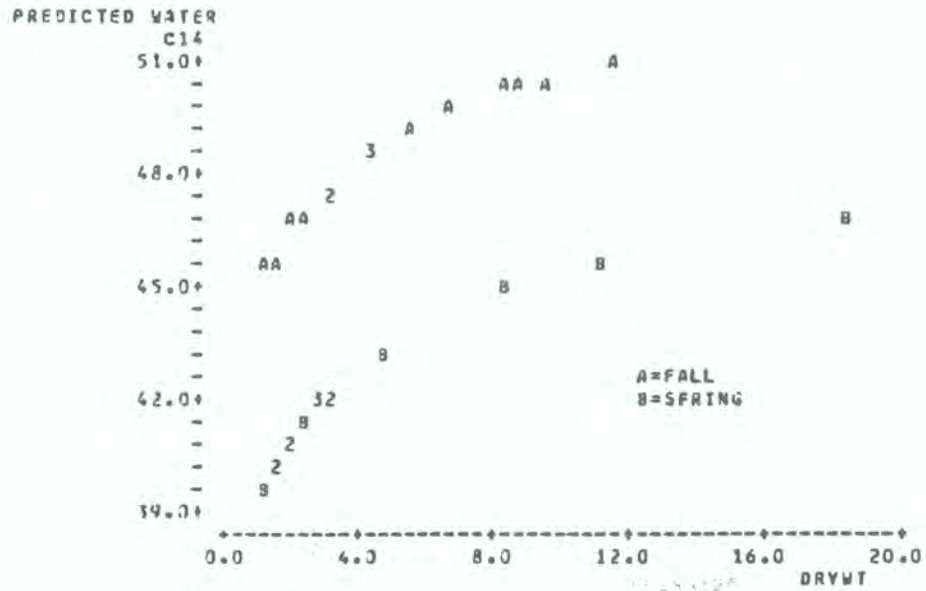
PREDICTED LN(WATER)

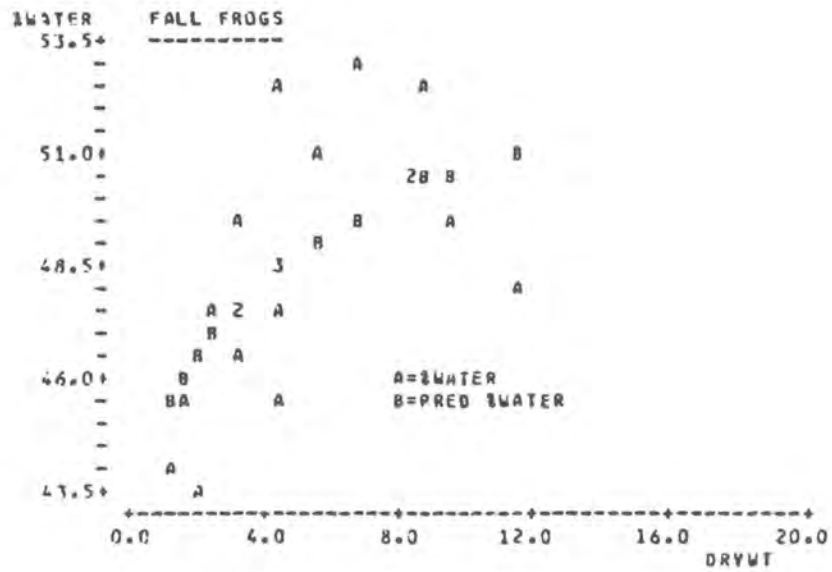
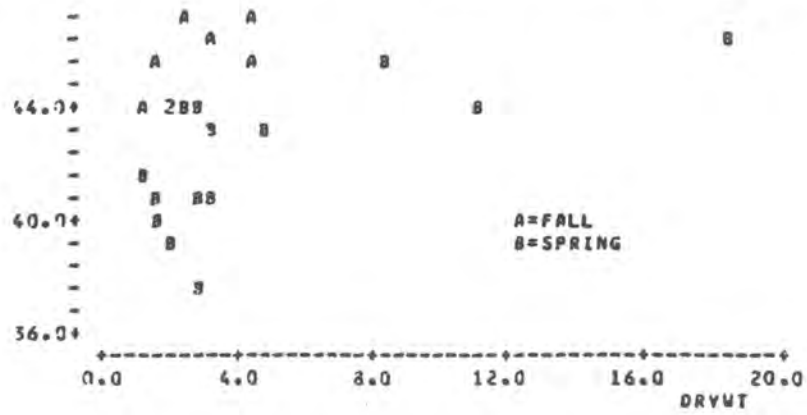


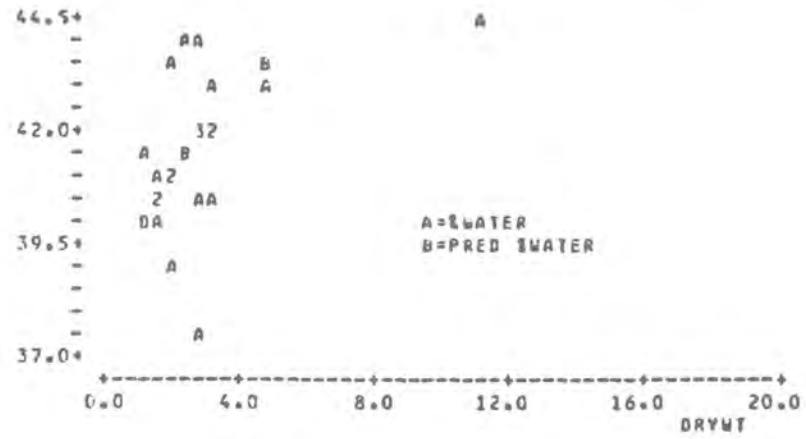


FALL:  
 $PRED. \%WATER = 100 \cdot 0.8063 \cdot (DRYWT + 1.10608) / (DRYWT + PRED. \%WATER)$   
 $= 100 \cdot 0.8063 \cdot (DRYWT + 1.10608) / (DRYWT + 0.8063 \cdot (DRYWT + 1.10608))$

SPRING:  
 $PRED. \%WATER = 100 \cdot 0.6462 \cdot (DRYWT + 1.10608) / (DRYWT + PRED. \%WATER)$   
 $= 100 \cdot 0.6462 \cdot (DRYWT + 1.10608) / (DRYWT + 0.6462 \cdot (DRYWT + 1.10608))$







## 6. INTRODUCTION TO MULTIVARIATE ANALYSIS

### 6.1 A priori structure in a data set:

In general a data set has  $n$  observations (usually  $n$  samples) on  $p$  variables. Typically there are  $n$  rows and  $p$  columns, so the data matrix can be represented by

$$X = \begin{matrix} & & P \\ n & \left[ \begin{array}{c} \\ \\ \\ \end{array} \right] \end{matrix}$$

Quite often the data matrix has a a priori structure. That is, we perceive the rows and/or the columns to fall into groups which existed conceptually before we examined the collected data, and preferably before we collected the data. In fact, this a priori structure usually represents the design of the data analysis which will be applied.

The figure 6.2 shows the various types of a priori structure of a data matrix. The dashed horizontal or vertical lines represent partitions of rows or columns into groups of rows or of columns. Each example suggests a category of types of data analysis, and the univariate cases should be familiar to you. (Figure taken from Green (1979) with permission).

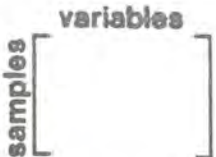
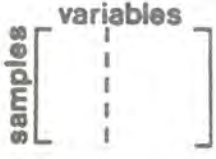
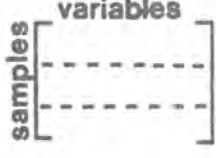
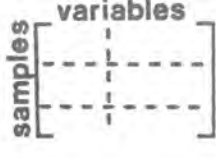
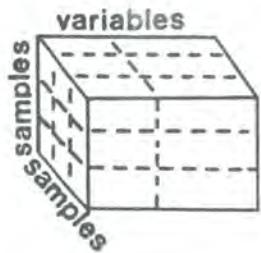
PARTITIONING	SYMBOLIC	MULTIVARIATE MODELS	UNIVARIATE MODELS
NONE		principal components analysis factor analysis cluster analysis	a priori restriction to one factor or component
VARIABLES		canonical correlation analysis	multiple regression and correlation
SAMPLES		MV analysis of variance and discriminant analysis	analysis of variance
VARIABLES AND SAMPLES		MV analysis of covariance and discriminant analysis with covariance	analysis of covariance
MULTI- DIMENSIONAL		factorial MV analysis of variance and covariance designs	factorial analysis of variance and covariance

FIGURE 6.2. Data matrices and statistical models

### 6.3 Types of data matrices and statistical analyses

#### 6.3.1 No a priori structure:

We have simply collected an observational data set,  $n$  observations on  $p$  variables. The  $p$  variables are not divided into "predicted" and "predictor" types, and the  $n$  observations are not divided into a priori groups (such as different treatments, locations, times). If  $p = 1$  we have the univariate case, and we would usually summarize the data graphically or do summary statistics appropriate for one column of data. These would be sample statistics such as  $\bar{x}$ ,  $s^2$ ,  $s$ ,  $SE$  and  $.95$   $ci$  on  $x$ . If  $p > 1$  we can of course do this for each variable, but besides looking at the pattern of variation of each variable we can also look at the pattern of covariation between each pair of variables. We speak of the "mean vector" and the "deviation cross-products matrix" and the "variance - covariance matrix".

For  $p = 3$  these are:

$$\bar{X} = [\bar{x}_1 \quad \bar{x}_2 \quad \bar{x}_3]$$

$$W = \begin{bmatrix} \Sigma(x_1 - \bar{x}_1)^2 & \Sigma(x_1 - \bar{x}_1)(x_2 - \bar{x}_2) & \Sigma(x_1 - \bar{x}_1)(x_3 - \bar{x}_3) \\ \Sigma(x_2 - \bar{x}_2)(x_1 - \bar{x}_1) & \Sigma(x_2 - \bar{x}_2)^2 & \Sigma(x_2 - \bar{x}_2)(x_3 - \bar{x}_3) \\ \Sigma(x_3 - \bar{x}_3)(x_1 - \bar{x}_1) & \Sigma(x_3 - \bar{x}_3)(x_2 - \bar{x}_2) & \Sigma(x_3 - \bar{x}_3)^2 \end{bmatrix}$$

$$D = W/(n-1) = \begin{bmatrix} s_1^2 & s_{12} & s_{13} \\ s_{21} & s_2^2 & s_{23} \\ s_{31} & s_{32} & s_3^2 \end{bmatrix} \quad \begin{bmatrix} S_1^2 & S_{12} & S_{13} \\ S_2 & S_{23} \\ S_3^2 \end{bmatrix}$$

If the data were standardized  $[(x - \bar{x}_j)/s_j]$  we would have the



correlation matrix

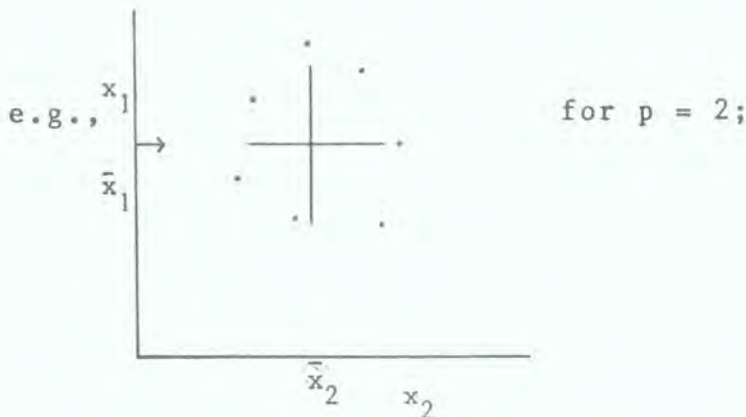
$$R = \begin{bmatrix} 1 & r_{12} & r_{13} \\ r_{12} & 1 & r_{23} \\ r_{31} & r_{32} & 1 \end{bmatrix} = \begin{bmatrix} 1 & r_{12} & r_{13} \\ & 1 & r_{23} \\ & & 1 \end{bmatrix} = \begin{bmatrix} r_{12} & r_{13} \\ & r_{23} \\ & & 1 \end{bmatrix}$$

In the univariate case, these are all quite trivial and boring:

$$\bar{x} = [\bar{x}] , D = [S^2] \text{ and } R = [1]$$

In the multivariate case we will:

(1) locate the mean vector of the data in a p-dimensional space,



(2) test the D or the R matrix for "structure" that is, against the null hypothesis that

$$R \text{ estimates } \rho = \begin{bmatrix} 1 & 0 & 0 \\ 0 & 1 & 0 \\ 0 & 0 & 1 \end{bmatrix} = \text{an identity matrix.}$$

(all true correlations are zero)

A confidence bound on a set of standardized observations, which yielded an R matrix that was an identity matrix, would be a sphere (for  $p = 3$ , a circle for  $p = 2$ , a spheroid for  $p > 3$ ).

Thus the usual test against  $H_0 =$  "identity matrix" is called a test of sphericity, or the "sphericity test". If "sphericity" is rejected, in favor of an elliptical shape then one is saying that at least some correlations appear to be non-zero, which is saying that the data "have structure";

(3) describe any structure, given that the  $H_0$  : "no structure" has been rejected. Some of the methods consist of identifying which variables are correlated. Other methods consist of finding, a posteriori, the partitions - of variables (vertical) or of observations (horizontal) - which "best" describe the structure in the data. Of course one can not turn around and test the "significance" of such partitions. It would make as much sense to separate a group of people into those shorter than the median height and those taller, and then test whether height differed between the groups. The only valid test, of "structure" versus "no structure", has already been done. Cluster analysis is a category of methods for partitioning observations into sets. If the  $n$ -by- $p$  data matrix is partitioned so that the  $n$  observations are grouped into  $g < n$  sets of observations, then we have a cluster analysis solution which has reduced the data matrix from  $n$ -by- $p$  to  $g$ -by- $p$ . If it is a "best" cluster analysis solution then it has in some sense done so in a manner that has retained the maximum possible information about the structure in the original unreduced data matrix. Analysis by ordination similarly reduces the data matrix so as to retain maximum possible information in the reduced description, but it does so by partitioning the variables into "best" sets. Thus the original  $p$ -dimensional space is reduced to  $k < p$  - dimensional space. Again, tests of the "significance" of the partitioning are not very meaningful.

### 6.3.2 A priori structure is partitioning of variables into groups:

Let us deal with the simplest and most common case where one

partition separates the variables into two sets, one on the left which is the criterion set and one on the right which is the predictor set. Here the a priori structure tells us that the data were collected in order to predict the left-hand variables from the right-hand variables. If a linear additive model is applicable (perhaps after transformation of the original variables) then we have the general linear model, or if neither set is clearly the predictor or the criterion variable set, then we have canonical correlation analysis. In the univariate case (one variable in the left-hand set), we have multiple regression and multiple correlation analysis, respectively. If, in the univariate case, the right-hand set also contains only one variable, then we have simple linear regression and correlation, respectively.

### 6.3.3 A priori structure is partitioning of observations into groups:

Observations can be partitioned into any number of sets. They may be treatments, locations, times, or combinations of those, but this a priori structure tells us that the data were collected in order to predict values on the  $p$  variables from knowledge of the group membership of an observation (or vice versa). Of course a test of whether there is any predictive power (whether the groups in fact differ on the variables) is the first step, and in the univariate case (one column in the data matrix) that is the only possible step (an ANOVA). When  $p > 1$ , and the test for group differences on the variables is significant, we would usually proceed to describe the group differences in terms of the relative contributions by the different variables to the group differences. The test would be a MANOVA (an acronym with obvious meaning). The descriptive analysis goes by various names: discriminant analysis, multiple discriminant analysis, and canonical analysis.

### 6.3.4 A priori structure which is a combination, or a multiple, of the above:

If partitioning is both vertical and horizontal then there is a predictor set of variables used to predict a criterion set of variables, but the observations on all these variables fall into different a priori groups. This is univariate (one criterion variable) or multivariate (>1 criterion variable) analysis of covariance. If groups, or treatment levels, are defined for more than one factor then we have a factorial UV or MV analysis of variance or covariance. Any univariate linear additive model - any regression, ANOVA or covariance design in existence - is just a special case of multivariate model.

#### 6.4 Example of some basic calculations for multivariate analysis

##### 6.4.1 Description

You have learned the necessary calculations and how to do them in MINITAB and in APL; matrix addition, multiplication, inversion, and transposition, and finding roots and vectors.

MINITAB does not have a command to calculate a W matrix or a D matrix, but the MINITAB job file (section ) shows how to do it.

Enter these 3-variable data into C1 - C3:

$$\begin{bmatrix} 4.5 & 2.9 & 3.0 \\ 4.9 & 4.1 & 3.1 \\ 4.2 & 3.5 & 3.3 \\ 4.1 & 3.8 & 2.9 \\ 4.7 & 3.6 & 3.6 \\ 4.4 & 3.7 & 3.5 \end{bmatrix}$$

Obviously there are n=6 observations on p=3 variables. Calculate the mean vector by doing

```
AVER C1, K1
AVER C2, K2
```

AVER C3, K3

Now use the MINITAB job on the attached sheet to calculate W and D. You should find that

$$D = \begin{bmatrix} 0.0907 & 0.0280 & 0.0213 \\ 0.0280 & 0.1600 & 0.0100 \\ 0.0213 & 0.0100 & 0.0787 \end{bmatrix}$$

Repeat this job run, but this time change lines 7-9, of the job file so the data are standardized on each variable. (Change to: LET Ci=(Ci-AVER(Ci))/STAN(Ci). Now D will be the R matrix. Is it the same as you obtain by doing "CORR C1-C3, M1"? Finally, with M1 containing the R matrix, do

```
EIGEN  M1, C4, M2
PRINT  C4
PRINT  M2
```

C4 contains the eigenvalues, which sum to 3 as did the diagonal of R. The columns of M2 contain the eigenvector coefficients associated with the eigenvalue above it. You have just done a principal components analysis!

#### 6.4.2 Program for calculating W and D : MINITAB

```
1  NRAND 50, 10, 2, C1
2  NRAND 50, 12, 3, C2
3  NRAND 50, 15, 4, C3
4  SET C4
5  3(49)
6  LET C4=1/C4
7  LET C1=C1-AVER(C1)
8  LET C2=C2-AVER(C2)
9  LET C3=C3-AVER(3)
10 COPY C1-C3 INTO M1
11 TRAN M1, M2
12 MULT M2 M1, M3
```

```

13 PRINT M3
14 DIAG C4, M4
15 MULT M4 M3, M4
16 PRINT M4
17 STOP

```

## 6.5 Some MINITAB examples for multivariate analysis

### 6.5.1 Example of eigenanalysis of non-symmetric matrix

```

PRINT M1
      MATRIX M1                4 ROWS BY          4 COLUMNS

      [ 1.23981   1.11477   0.28177   0.32404
        1.70016   1.52869   0.38640   0.44436
       -0.52596  -0.47290  -0.11954  -0.13747
        9.21021   8.28137   2.09321   2.40721 ] = W-1 A

* MULT M1 BY M1,M2
* MULT M2 BY M2,M2
* MULT M2 BY M2,M2          "Powering" the matrix
* MULT M2 BY M2,M2
* MULT M2 BY M1,M3
* COPY M2 TO C11-C14
* COPY M3 TO C15-C18
* LET K3=(SUM(C15))/(SUM(C11))
* PRINT K3
      K3          5.05618          = the first root
* LET C10=C15/C11
* PRINT C10

COLUMN      C10
COUNT      4
      5.05618      5.05618      5.05618      5.05618 - check

* LET C10=C11/10000
* PRINT C10
COLUMN      C10

```

```
COUNT          4
      4473961.      6135180.  -1897976.      33235921.
```

```
*LET C10=C10/SQRT(SUM(C10*C10))
```

```
*PRINT C10
```

```
COLUMN      C10
```

```
COUNT          4
```

```
[0.131028      0.179680  -0.055586      0.973374]
```

= the vector associated with the 1st root

### 6.5.2 Example of calculation of determinant of a matrix

```
*PRINT M1
```

```
MATRIX M1
```

```
3 ROWS BY
```

```
3 COLUMNS
```

```
[ 1.00000      0.70000      0.80000
  0.70000      1.00000      0.60000
  0.80000      0.60000      1.00000 ]
```

```
* EIGEN M1,C1,M2
```

```
* LET C2=LOGE(C1)
```

```
* LET K1=EXPO(SUM(C2))
```

```
* PRINT K1
```

```
K1      0.182000      - the determinant
```

### 6.5.3 Test of sphericity on a correlation matrix.

```
PRINT K1-K3
```

```
K1      0.182000      = determinant of the matrix
```

```
K2      49.0000      = number samples less one
```

```
K3      3.00000      = number variables
```

```
* LET K4=-((K2-(2*K3+5))/6)*LOGE(K1)
```

```
* LET K5=K3*(K3-1)/2
```

```
* PRINT K4-K5
```

```
K4      80.3601 = X2
```

```
K5      3.00000 = df
```

## 7. ORDINATION AND CLUSTER ANALYSIS

### 7.1 Tutorial/assignment

The data set to be used will be 'SEDABC DATA'. These are sediment samples obtained by grabs from 10-20m depth (below mean low water) at 3 locations 1 km apart in the lower Bay of Fundy on the Atlantic coast of Canada (where these samples were taken, the tidal range is about 18m). There are 60 samples (n=60), 20 from each of the 3 locations A, B and C. There are 4 variables: % sand, % silt-clay, % gravel, and organic content as % of total dry weight. The first 3 variables add to 100%. The 5th column of data contains "location codes": 1=A, 2=B, and 3=C.

To begin with, we will ignore the fact that we know that the samples come from 3 locations. We will treat the data as "unpartitioned" for purposes of analysis, and we will apply a principal components analysis, a cluster analysis, and the "variable subset section" FORTRAN program 'RSLCTIBM FORTRAN' (based on an algorithm originally proposed by L. Orloci). Each of these methods somehow "look for" partitions of the data in order to describe the structure in the data. After these analyses we will "remember" that the samples come from different locations and we will see whether the structure that has been described is related to the locations.

We will use MINITAB, SAS, APL, and a FORTRAN program. The Orloci & Kenkel Apple DOS 3.3 BASIC programs also include programs analagous to those we will use. We will not use them as part of this tutorial/assignment. However some of you may wish to try them if you are going to be limited to BASIC programs "back home".

#### 7.1.1 MINITAB

Run the MINITAB example of doing a PCA (a handout you have already been given). Include the "sphericity test" insert it just after you do the "EIGEN---" command. (Choose your own C, K, and M numbers so they do not conflict with the PCA analysis!) If you have problems with the sphericity test because one of the



roots is zero, then drop the 4th root which is zero and do the sphericity test using only the 3 non-zero roots. Run interactively first, then as a batch job, and then print out the 'fn MINITAB' and 'fn OUTPUT'.

### 7.1.2 APL

- a. Now go into APL. If you have not yet done so, read the descriptions of functions MATFORM, COVAR, GEIG, and ISOTROPY (by entering each name with "DES" appended). If that is unclear, enter "DESCRIBEFNS".
- b. The workspace UNESCO also contains the variable SEDABC. Enter 'SEDABC' and you will see the same data set as in the file 'SEDABC DATA'. The function MATFORM was used to enter the data and shape them into this 60-by-5 matrix.
- c. Run COVAR using the option to create a covariance matrix (enter '0 COVAR SEDABC [;1 2 3 4]'). Do you understand the bracketed part? If not, just enter 'SEDABC [;1 2 3 4]' and compare the response with the response you get when you enter 'SEDABC'. Rename the covariance matrix from M to MC (enter 'MC←M').  
Run COVAR again, this time with the option to create a correlation matrix (enter '1 COVAR SEDABC [;1 2 3 4]'). Rename it to MR (enter 'MR←M').
- d. Now run GEIG on the covariance matrix by entering 'GEIG MC'. Write down this root (which is for PC I) and its associated vector, then follow the instructions and continue by entering 'GEIG N'. Write down the root and vector for PC II. Again enter 'GEIG N' and write down the root and vector for PC III. Sum the 3 roots. Enter 'MC' and sum the diagonal elements (the variances) in the covariance matrix. Are they the same? If they are, then the 4th root is zero (as you know it is from the MINITAB analysis), so there is no point in doing 'GEIG N' again.
- e. Repeat (d) on the correlation matrix (stored in MR). Also do the sphericity test on the correlation matrix, by entering '60 ISOTROPY roots', where 60 is the number of samples from which the correlation matrix was calculated and "roots" is a vector containing the roots. Again, you can not include

- a zero root so leave out the 4th root, which is zero.
- f. Compare all your APL results with those you obtained using MINITAB.

### 7.1.3 SAS

- a. Now prepare a file named 'SEDABC SAS', using XEDIT. Your file should look like this:

```
TITLE SAS ANALYSIS ON SEDIMENT DATA;
DATA SEDABC;
INPUT PERSAND PERSLTCL PERGRAV PERORG LOCATION;
CARDS;
```

(the SEDABC data go here - use the 'GET SEDABC DATA' command)

```
PROC PRINT;
PROC PLOT; PLOT PERSAND*PERSLTCL=LOCATION;
PROC PLOT; PLOT PERSAND*PERGRAV=LOCATION;
PROC PLOT; PLOT PERSAND*PERORG=LOCATION;
PROC PLOT; PLOT PERSLTCL*PERGRAV=LOCATION;
PROC PLOT; PLOT PERSLTCL*PERORG=LOCATION;
PROC PLOT; PLOT PERGRAV*PERORG=LOCATION;
PROC PRINCOMP OUT=COVPCS COV; VAR PERSAND PERSLTCL PERGRAV
PERORG;
PROC PRINCOMP DATA=SEDABC OUT=CORPCS; VAR PERSAND PERSLTCL
PERGRAV PERORG;
PROC PLOT DATA=COVPCS; PLOT PRIN1*PRIN2=LOCATION;
PROC PLOT DATA=CORPCS; PLOT PRIN1*PRIN2=LOCATION;
PROC CLUSTER DATA=SEDABC OUTTREE=TREE;
VAR PERSAND PERSLTCL PERGRAV PERORG; ID LOCATION;
PROC PRINT DATA=TREE;
PROC PLOT; PLOT _CCC_ * _NCL_;
```

- b. Run the SAS job, and look at the output (enter 'TY SEDABC LISTING'). Compare the correlations between the variables with the bivariate plots of the variables. Compare the bivariate plots of the variables with the "PC I vs. PC II"

plots. How do the PCA and cluster analysis results relate to the three locations?

#### 7.1.4 FORTTRAN

Now we will run the FORTRAN program 'RSLCTIBM FORTRAN'. This is the algorithm which selects a subset of variables, such that the subset best represents (is most highly correlated with) the whole set.

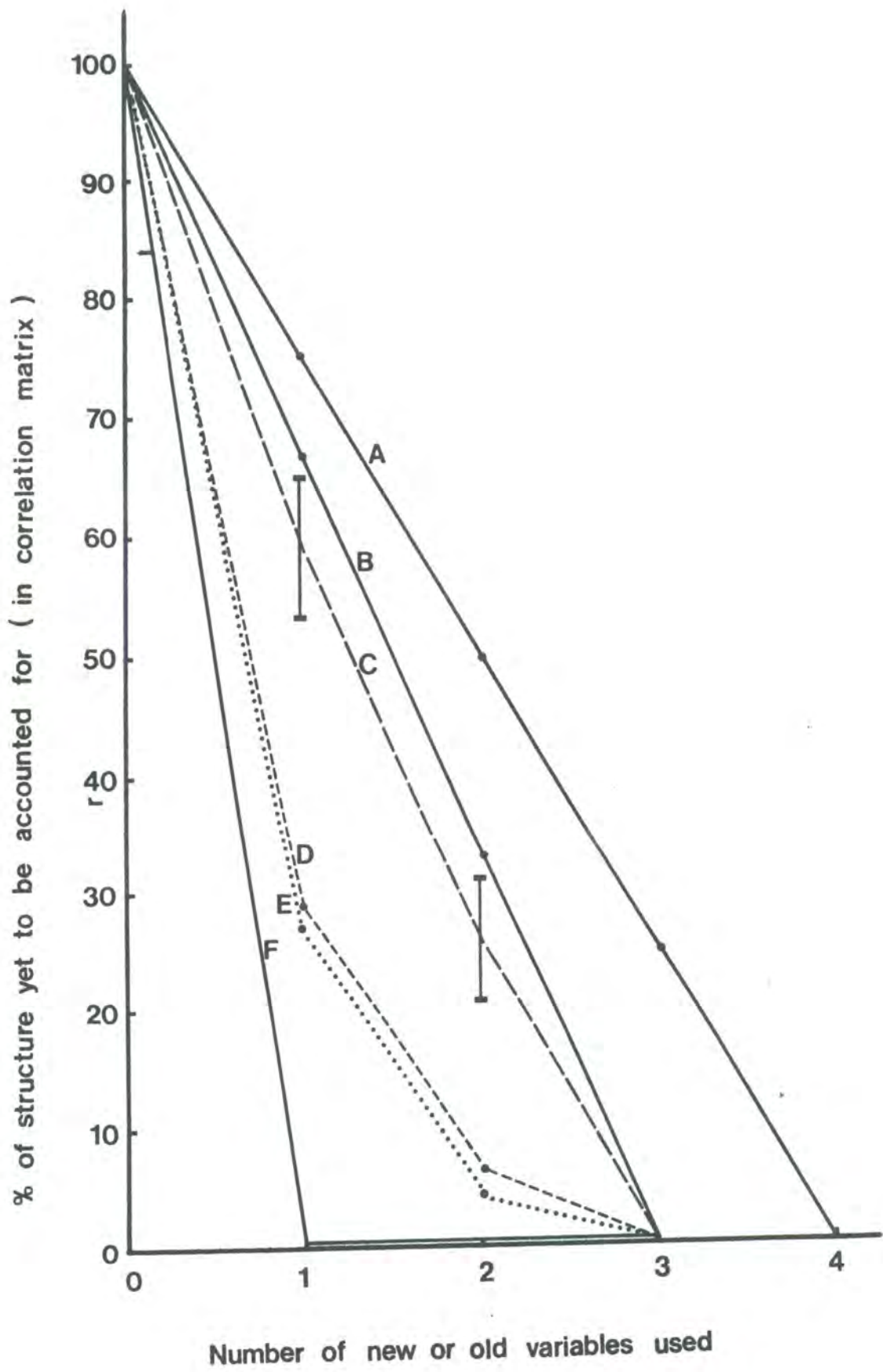
- a. Enter 'TY RSLCTIBM DATA' and observe that the data are the same as 'SEDABC DATA'.
- b. Enter 'TY RSLCTIBM FORTRAN' and read the comment lines that proceed the program itself.
- c. Now run the program, by entering 'FORTVS RSLCTIBM'.  
Wait for the final "R;-----", and then enter 'TY RSLCTIBM OUTPUT'. Read it and try to understand it.
- d. Note the first 2 variables "selected". What percent of the total correlation structure do they account for? What percent of the correlation structure did the first two principal components (PCs) account for (refer to MINITAB, SAS or APL runs)? Are the 2 "best variables" almost as good at accounting for correlation structure as the 2 best linear combinations of all 4 variables (that is one way of saying what PCs are)? Look at the SAS bivariate plots. Does the bivariate plot of the 2 best variables against each other show the most information? Does it show low or high correlation?
- e. Look at the vectors associated with the first 2 PCs. Is the coefficient associated with the "best variable" in the PC I vector relatively large in magnitude? Is the coefficient associated with the "2nd best variable" in the PC II vector relatively large in magnitude?

#### 7.1.5 Overall evaluation.

Now try to evaluate all this. You used 4 analytical approaches to evaluate the correlation structure in this n-60-by-p-4 data matrix, and you ignored the information contained in a 5th "location code" variable. Try to answer the following

questions:

- a. Does "location" appear to be reacted to, or involved in, the correlation structure you described when ignoring the location information? Try to interpret any relationship you see.
- b. You used 4 analytical methods or approaches: (1) the sphericity test of the  $H_0$ : "no nonzero correlations" which is equivalent to  $H_0$ : "no correlation structure"; (2) principal components analysis (PCA) which finds new variables (new axes) which most efficiently display the structure in the data; (3) RSLCT which selects the best of the original variables for display of the structure in the data; and (4) cluster analysis which finds the best groups of samples to describe the structure. Can you see how they are describing (testing in the case of the sphericity test) the same structure in this data set, though in different ways? Which do you think does the best job (go ahead and be subjective!)?
- c. You used MINITAB, SAS, APL, and FORTRAN program. (You may also have used some of the Orloci & Kenkel Apple DOS 3.3 programs.) Can you see that the results (e.g. PCA, sphericity test) are basically the same when done by the different languages or packages? Do you have likes and dislikes related to ease of use, clarity of output, or any other characteristic?



- A : No redundancy in 4 dimensions  
= all 4 roots exactly equal  
= sphericity in 4 dimensions  
(all zero correlations)
- B : No redundancy in 3 dimensions  
= 1 root zero, rest exactly equal  
= sphericity in 3 dimensions  
(all zero correlations)
- C : Expected when random  
uncorrelated data are  
used (nonzero correlations  
by chance only).  
.95 cls on a single run  
are shown.  
(In 3 dimensions.)
- D : RSLCT on sediment data.
- E : PCA on sediment data.
- F : Total redundancy  
= only 1 nonzero root  
= all perfect correlations.

7.2. Job Listings and Outputs.

FILE: PCA2      MINITAB A1    VM/SP - CONVERSATIONAL MONITOR SYSTEM

PAGE 001

READ C1-C5

2.350	97.150	0.0	7.325	1.
4.522	94.250	1.123	3.479	1.
3.564	95.744	0.472	8.940	1.
3.117	96.592	0.191	8.753	1.
1.513	94.170	4.317	4.451	1.
2.520	97.430	0.0	6.377	1.
2.374	76.120	0.296	7.058	1.
2.460	77.467	0.073	5.574	1.
31.543	60.702	7.750	5.190	1.
5.695	93.341	0.454	7.330	1.
2.365	95.763	0.172	3.734	1.
8.561	90.075	1.341	5.192	1.
4.237	95.593	0.170	7.303	1.
4.329	95.671	0.0	7.560	1.
2.078	97.344	0.553	7.034	1.
1.751	98.123	0.166	7.397	1.
2.307	97.091	0.0	3.377	1.
4.385	94.620	0.977	7.570	1.
3.295	95.477	1.223	5.972	1.
2.953	97.047	0.0	3.671	1.
46.607	53.003	0.390	5.238	2.
74.374	24.588	0.538	2.356	2.
80.096	19.211	0.673	2.412	2.
81.447	17.401	1.150	2.236	2.
78.150	20.590	1.150	2.589	2.
49.475	49.396	1.127	5.073	2.
47.414	52.413	0.168	5.709	2.
85.553	12.607	0.835	4.360	2.
50.699	43.619	0.692	2.275	2.
80.388	18.146	1.465	2.259	2.
53.262	41.068	0.670	3.572	2.
60.382	38.412	1.206	3.933	2.
55.870	43.314	0.816	2.243	2.
74.645	23.567	1.788	2.638	2.
74.699	23.581	1.720	2.705	2.
42.257	57.207	0.536	5.542	2.
41.767	57.494	0.549	6.297	2.
72.371	27.043	0.586	2.055	2.
76.056	23.502	0.342	2.563	2.
80.534	13.572	0.874	2.534	2.
0.650	99.350	0.0	7.909	3.
0.600	99.400	0.0	7.364	3.
0.550	99.450	0.0	6.337	3.
0.385	99.015	0.0	6.175	3.
0.704	99.296	0.0	7.278	3.
1.203	93.792	0.0	4.300	3.
1.469	98.531	0.0	7.524	3.
0.855	99.145	0.0	7.914	3.
2.193	97.807	0.0	7.535	3.
1.332	93.168	0.0	7.298	3.
0.960	99.005	0.035	6.396	3.
0.754	99.193	0.053	6.333	3.
1.150	98.350	0.0	11.752	3.
2.078	97.922	0.0	7.526	3.

```

1.053  93.147  0.0  7.490  1.
0.344  99.126  0.030  5.069  3.
1.668  93.332  0.0  5.096  3.
1.501  93.179  0.021  5.797  1.
1.387  93.713  0.0  0.627  3.
2.504  97.391  0.175  5.354  3.
PRINT C1-C5
CORR C1-C4, M1
PRINT M1
NOTE M1 IS THE CORRELATION MATRIX
EIGEN M1, C6, M2
PICK 1 3 C6, C7
NOTE C7 CONTAINS THE FIRST THREE (NON-ZERO) ROOTS OF THE CORR MATRIX
LET C8=LUGE(C7)
NOTE C8 CONTAINS THE LOG OF THE FIRST 3 ROOTS OF THE CORRELATION MATRIX
LET K1=EXP(SUM(C8))
PRINT K1
NOTE K1 IS THE PRODUCT OF THE FIRST THREE ROOTS OF THE CORRELATION MATRIX
LET K2=37
NOTE K2 IS THE NUMBER OF OBSERVATIONS MINUS ONE
LET K3=4
NOTE K3 IS THE NUMBER OF VARIABLES
LET K4=-(K2-(2*K3+5)/6)*LUGE(K1)
NOTE K4 IS THE CHI-SQUARE VALUE FOR SPHERICITY TEST
LET K5=K3*(K3-1)/2
NOTE K5 IS THE DEGREES OF FREEDOM
PRINT K4-K5
PRINT C5
NOTE C6 GIVES THE EIGENVALUES OR ROOTS OF THE CORRELATION MATRIX
SUM C6, K1
NOTE K1 IS THE SUM OF THE EIGENVALUES, AND SHOULD HAVE VALUE 4
LET C7=100*C6/4
PRINT C7
NOTE C7 ARE THE EIGENVALUES GIVEN IN PERCENTAGE
PRINT M2
NOTE M2 GIVES THE EIGENVECTORS OF THE CORRELATION MATRIX
LET C1=(C1-AVER(C1))/STAN(C1)
LET C2=(C2-AVER(C2))/STAN(C2)
LET C3=(C3-AVER(C3))/STAN(C3)
LET C4=(C4-AVER(C4))/STAN(C4)
NOTE C1-C4 NOW CONTAIN THE Z TRANSFORMATION OF THE ORIGINAL DATA
COPY C1-C4 INTO M3
PRINT M3
NOTE M3 IS THE MATRIX CONTAINING THE STANDARDISED VALUES OF THE DATA
MULT M3 M2, M4
COPY M4 INTO C8-C11
DESCRIBE C8-C11
NOTE C8-C11 ARE THE PCI-PCIV
WIDTH 100, 50
LPLT C8 C9, C5
NOTE THIS GIVES THE PLOT OF PCI VERSUS PCII
STOP

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 MAY 4, 1985 \*\*\* NATIONAL UNIVERSITY OF SINGAPORE - LOCAL VERSION 02/12/1982  
 STORAGE AVAILABLE 4900

PRINCIPAL COMPONENTS ANALYSIS ON SEDIMENT PARAMETERS  
 OF CANADIAN GRASS SAMPLES USING MINITAB

COLUMN COUNT	C1 60	C2 60	C3 60	C4 60	C5 60	AREA CODE
ROW	% SAND	% SILT-CLAY	% GRAVEL	% O.M.		
1	2.8500	97.1500	0.0	9.3250	1.	
2	4.6200	94.2500	1.1200	8.4700	1.	
3	3.5640	95.9440	0.49200	8.8400	1.	
4	3.1170	96.6920	0.19100	8.7580	1.	
5	1.5130	94.1700	4.31700	8.4510	1.	
6	2.5200	97.4300	0.0	6.8790	1.	
7	2.8940	96.8200	0.28600	7.0580	1.	
8	2.4600	97.4670	0.07300	6.8940	1.	
9	31.5480	60.7020	7.75000	6.1900	1.	
10	5.6350	93.8410	0.46400	7.3300	1.	
11	2.9650	96.9530	0.17200	3.7340	1.	
12	8.5610	90.0790	1.34100	6.1920	1.	
13	4.2370	95.5930	0.17000	7.3030	1.	
14	4.3200	95.6710	0.0	7.5600	1.	
15	2.0980	97.3440	0.55800	7.0340	1.	
16	1.7610	98.1230	0.16600	7.8970	1.	
17	2.3090	97.6910	0.0	8.3770	1.	
18	4.3360	94.6200	0.97400	7.5300	1.	
19	3.2950	95.4770	1.22000	5.9720	1.	
20	2.9530	97.0470	0.0	8.6910	1.	
21	46.6070	53.0030	0.39000	5.2380	2.	
22	74.8740	24.5380	0.53300	2.8560	2.	
23	30.0960	19.2110	0.59300	2.4120	2.	
24	81.4400	17.4010	1.15000	2.2360	2.	
25	78.1500	20.6900	1.16000	2.6690	2.	
26	49.4750	49.3950	1.12000	5.0730	2.	
27	47.4140	52.4130	0.16800	5.7090	2.	
28	56.5590	17.6070	0.83500	4.7600	2.	
29	50.6990	43.6190	0.68200	2.2750	2.	
30	20.3380	13.1460	1.46600	2.2500	2.	
31	59.2520	41.0680	0.57000	3.5920	2.	
32	60.3220	38.4120	1.20600	3.9300	2.	
33	55.8700	43.3140	0.31600	2.2430	2.	
34	74.6450	23.5670	1.78300	2.6300	2.	
35	74.6990	23.5310	1.72000	2.7050	2.	
36	42.2370	57.2070	0.53600	5.5420	2.	
37	41.9370	57.4840	0.54000	6.2990	2.	
38	72.3710	27.0430	0.59600	2.0550	2.	
39	76.0560	23.6020	0.34200	2.6530	2.	
40	80.5340	18.5720	0.39400	2.5840	2.	
41	0.6500	97.3500	0.0	7.9080	3.	
42	0.6000	99.4000	0.0	7.8640	3.	
43	0.5500	97.4500	0.0	6.3370	3.	
44	0.9450	97.0150	0.0	6.8750	3.	
45	0.7040	99.2960	0.0	7.2200	3.	

	% SAND	% SILT-CLAY	% GRAVEL	% O.M.	AREA CODE
46	1.2080	98.7920	0.0	4.3000	3.
47	1.4690	98.5310	0.0	7.5240	3.
48	0.8550	99.1450	0.0	7.9140	3.
49	2.1930	97.8070	0.0	7.5350	3.
50	1.8320	99.1680	0.0	7.2980	3.
51	0.9600	99.0050	0.03500	6.3960	3.
52	0.7540	99.1380	0.05800	6.3330	3.
53	1.1500	93.8500	0.0	11.7620	3.
54	2.0790	97.9220	0.0	7.5260	3.
55	1.6530	98.3470	0.0	7.4900	3.
56	0.8440	99.1260	0.03000	6.0680	3.
57	1.6580	93.3320	0.0	6.0860	3.
58	1.8010	99.1780	0.02100	6.9970	3.
59	1.0370	93.9130	0.0	6.6290	3.
60	2.5040	97.3910	0.10500	5.8540	3.

-- THE FOLLOWING ARE THE CORRELATION COEFFICIENTS BETWEEN CI AND Cj

	C1	C2	C3
C2	-0.999		
C3	0.275	-0.309	
C4	-0.961	0.857	-0.194

-- M1 IS THE CORRELATION MATRIX OF THE SEDIMENT PARAMETERS

MATRIX M1                    4 ROWS BY            4 COLUMNS

1.00000	-0.99936	0.27496	-0.96114
-0.99936	1.00000	-0.30918	0.85901
0.27496	-0.30918	1.00000	-0.19412
-0.96114	0.85901	-0.19412	1.00000

-- K1            3.469345    DETERMINANT OF THE 3-ROOT DIAGONAL MATRIX

-- K4            42.9897    CHI-SQUARE VALUE FOR SPHERICITY TEST

K5            6.00000    ASSOCIATED DEGREES OF FREEDOM

-- C6 GIVES THE EIGENVALUES OR ROOTS OF THE CORRELATION MATRIX

COLUMN	C6
COUNT	4
	2.92073    0.70049    0.17937    0.00700

-- NOTE THAT THE SUM OF THE ROOTS EQUALS THE SUM OF THE DIAGONAL ELEMENTS OF THE CORRELATION MATRIX, I.E.

SUM            =            4.0000

C7 GIVES THE ROOTS OR EIGENVALUES IN PERCENTAGE  
 COLUMN C7  
 COUNT 4

73.0133	22.5223	4.4593	0.0300
---------	---------	--------	--------

--  
 M2 IS THE MATRIX OF EIGENVECTORS OF THE CORRELATION MATRIX  
 MATRIX M2 4 ROWS BY 4 COLUMNS

-0.572831	0.116192	0.405130	0.702979
0.575171	-0.079755	-0.397237	0.710719
-0.223838	-0.768759	-0.091743	0.026452
0.537211	-0.274421	0.913301	-0.000014

--  
 M3 GIVES THE STANDARDISED VALUES OF THE ORIGINAL DATA  
 MATRIX M3 60 ROWS BY 4 COLUMNS

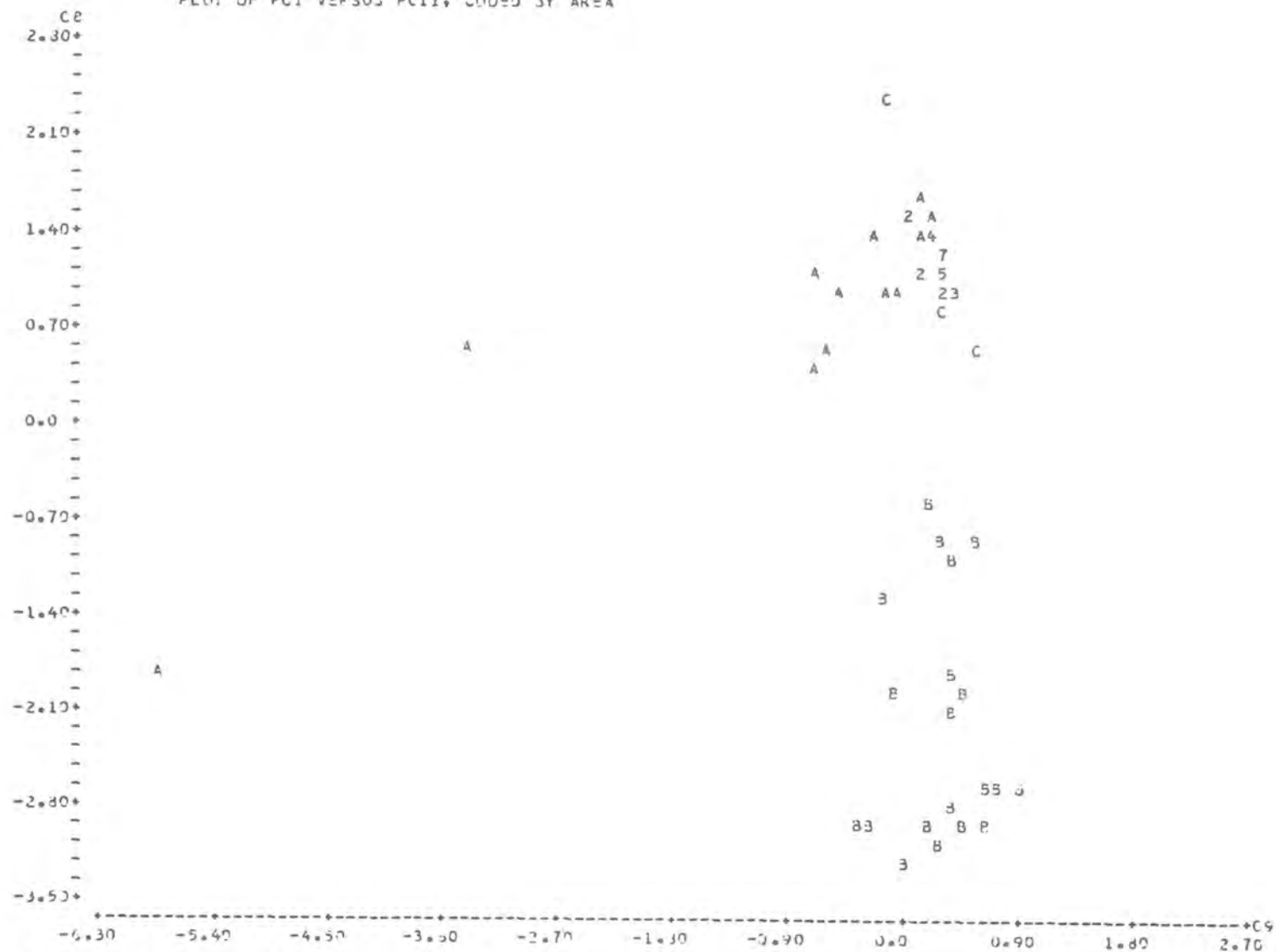
-0.67544	0.63756	-0.52371	1.43570
-0.61907	0.59555	0.43693	1.06141
-0.55256	0.64934	-0.10471	1.22112
-0.60689	0.67304	-0.36105	1.18484
-0.71223	0.59312	3.15279	1.04902
-0.53601	0.69901	-0.52371	0.35352
-0.67403	0.57710	-0.29014	0.43272
-0.68794	0.69760	-0.46154	0.36016
0.24404	-0.45752	6.07645	0.04869
-0.58429	0.53269	-0.12855	0.55306
-0.67496	0.58153	-0.37723	1.17423
-0.49246	0.46407	0.61933	0.04957
-0.63100	0.63821	-0.37893	0.54111
-0.62805	0.64069	-0.52371	0.65482
-0.69953	0.59373	-0.04850	0.42710
-0.71033	0.71939	-0.33234	0.80391
-0.69277	0.70470	-0.52371	1.01628
-0.62623	0.50738	0.32281	0.64154
-0.56119	0.63454	0.52210	-0.04776
-0.57214	0.58420	-0.52371	1.15520
0.72553	-0.71151	-0.17157	-0.37250
1.63221	-1.61201	-0.05553	-1.42636
1.79952	-1.79241	0.05647	-1.62240
1.34237	-1.83977	0.45567	-1.70967
1.73717	-1.73554	0.46419	-1.50025
0.91842	-0.92532	0.43779	-0.44550
0.75239	-0.73025	-0.39063	-0.15412
2.00636	-1.79170	0.19740	-0.75095
0.35764	-0.35044	0.05710	-1.63741
1.89883	-1.81516	0.72479	-1.59049
1.09995	-1.09974	0.04639	-1.10074
1.16749	-1.17391	0.57336	-0.04927
1.02332	-1.01936	0.17122	-1.69757
1.62457	-1.64436	0.97971	-1.52281
1.62660	-1.64792	0.94110	-1.49317
0.53716	-0.57320	-0.00723	-0.23900

0.57787	-0.56950	-0.05616	0.09691
1.55201	-1.53421	-0.02465	-1.78075
1.67008	-1.64325	-0.23245	-1.51175
1.81355	-1.80256	0.23765	-1.54670
-0.74593	0.75728	-0.52371	0.80878
-0.74753	0.75988	-0.52371	0.78931
-0.74913	0.75045	-0.52371	0.11373
-0.73519	0.74666	-0.52371	0.36060
-0.74420	0.75557	-0.52371	0.50793
-0.72805	0.73959	-0.52371	-0.78750
-0.71969	0.73132	-0.52371	0.63999
-0.73936	0.75078	-0.52371	0.81144
-0.69649	0.70338	-0.52371	0.64376
-0.70906	0.71702	-0.52371	0.53990
-0.73600	0.74634	-0.49390	0.13983
-0.74260	0.75214	-0.47431	0.11196
-0.72991	0.74143	-0.52371	2.51390
-0.70018	0.71202	-0.52371	0.63977
-0.71379	0.72549	-0.52371	0.62385
-0.73971	0.75018	-0.47915	-0.00529
-0.71331	0.72502	-0.52371	0.00268
-0.70905	0.72014	-0.50582	0.40573
-0.73193	0.74343	-0.52371	0.24292
-0.68653	0.69519	-0.43429	-0.07997

--  
 C8-C11 ARE THE PRINCIPAL COMPONENTS, I.E. PC I-PC IV  
 THE DESCRIPTIVE STATISTICS FOR THE FOUR PRINCIPAL COMPONENTS ARE:

C8	N = 60	MEAN = 0.000011224	ST.DEV. = 1.71
C9	N = 60	MEAN = -0.000005957	ST.DEV. = 0.949
C10	N = 60	MEAN = 0.000010253	ST.DEV. = 0.422
C11	N = 60	MEAN = 0.000007924	ST.DEV. = 0.000144

PLOT OF PCI VERSUS PCII, CODED BY AREA



\*\*\* MINITAB WWW STATISTICS DEPT + PENN STATE UNIV. \* RELEASE 31.1 \*  
 STOPPAGE AVAILABLE 4300

TITLE SAS ANALYSIS ON SEDIMENT DATA;  
 DATA SEDABC;  
 INPUT PERDAND PERSLTCL PERGRAV PERORG LOCATION;  
 CARDS;

2.45J	97.150	0.0	9.325	1.
4.522	94.250	1.124	4.477	1.
3.564	93.944	0.492	8.840	1.
3.117	92.092	0.191	8.793	1.
1.513	94.170	4.317	3.451	1.
2.520	97.480	0.0	6.379	1.
2.574	96.120	0.296	7.053	1.
2.460	97.487	0.073	6.394	1.
31.544	63.702	7.750	0.190	1.
3.095	93.341	0.464	7.330	1.
2.565	96.363	0.172	8.734	1.
3.561	92.091	1.341	6.192	1.
4.277	95.593	0.170	7.303	1.
4.329	95.071	0.0	7.560	1.
2.304	97.344	0.593	7.034	1.
1.761	93.123	0.156	7.397	1.
2.309	97.691	0.0	3.377	1.
4.386	94.520	0.974	7.530	1.
3.275	95.477	1.233	5.972	1.
2.953	97.047	0.0	8.691	1.
46.607	53.003	0.300	5.233	1.
74.874	24.533	0.538	2.356	2.
80.096	17.211	0.693	2.412	2.
81.449	17.401	1.150	2.236	2.
73.150	20.690	1.160	2.689	2.
49.475	47.396	1.129	5.073	2.
47.414	52.413	0.153	5.709	2.
85.558	12.607	0.935	4.360	2.
50.699	43.519	0.632	2.275	2.
30.383	19.146	1.466	2.259	2.
58.262	41.063	0.670	3.592	2.
60.382	33.412	1.206	3.733	2.
55.870	43.314	0.810	2.243	2.
74.545	23.567	1.783	2.633	2.
74.699	23.591	1.720	2.705	2.
42.257	57.207	0.536	5.542	2.
41.767	57.434	0.549	6.299	2.
72.371	27.043	0.596	2.055	2.
76.056	23.002	0.342	2.663	2.
80.534	18.572	0.894	2.584	2.
0.650	99.350	0.0	7.703	3.
0.600	99.400	0.0	7.364	3.
0.550	99.450	0.0	6.337	3.
0.485	99.015	0.0	6.895	3.
0.704	97.296	0.0	7.221	3.
1.203	98.792	0.0	4.300	3.
1.469	98.531	0.0	7.524	3.
0.855	99.145	0.0	7.914	3.
2.193	97.807	0.0	7.535	3.
1.332	98.168	0.0	7.278	3.
0.960	99.005	0.035	6.396	3.

0.754	97.184	0.053	5.333	1.
1.150	93.150	0.0	11.762	3.
2.073	97.922	0.0	7.526	3.
1.553	93.347	0.0	7.490	3.
0.344	99.125	0.033	5.063	3.
1.063	93.332	0.0	5.095	3.
1.301	94.174	0.021	6.997	3.
1.027	94.913	0.0	5.629	1.
2.504	97.391	0.175	5.354	3.

```
PROC PRINT;
PROC PLOT; PLOT PERSAND*PERSLTCL=LOCATION;
PROC PLOT; PLOT PERSAND*PERGRAV=LOCATION;
PROC PLOT; PLOT PERSAND*PERORG=LOCATION;
PROC PLOT; PLOT PERSLTCL*PERGRAV=LOCATION;
PROC PLOT; PLOT PERSLTCL*PERORG=LOCATION;
PROC PLOT; PLOT PERGRAV*PERORG=LOCATION;
PROC PRINT; OUT=CQVPCS; VAR PERSAND PERSLTCL PERGRAV PERORG;
PROC PRINT; DATA=SEDARC; OUT=CQRPCS; VAR PERSAND PERSLTCL PERGRAV
PERORG;
PROC PLOT; DATA=CQVPCS; PLOT PRIN1*PRIN2=LOCATION;
PROC PLOT; DATA=CQRPCS; PLOT PRIN1*PRIN2=LOCATION;
PROC CLUSTER; DATA=SEDARC; OUTTREE=TREE;
VAR PERSAND PERSLTCL PERGRAV PERORG; ID LOCATION;
PROC PRINT; DATA=TREE;
PROC PLOT; PLOT _CCC_=_HCL_;
```

## SAS ANALYSIS ON SEDIMENT DATA

17:09 TUESDAY, MAY 7, 1985

1

055	PERSAND	PERSLTC	PERGRAV	PERORG	LOCATION
1	2.950	97.150	0.000	9.325	1
2	4.622	74.250	1.129	8.479	1
3	3.564	75.944	0.492	8.840	1
4	3.117	96.692	0.191	8.758	1
5	1.513	94.170	4.317	8.451	1
6	2.520	97.430	0.000	6.879	1
7	2.894	96.820	0.286	7.058	1
8	2.460	97.467	0.073	6.894	1
9	31.548	50.702	7.750	6.190	1
10	5.695	93.841	0.464	7.330	1
11	2.865	76.053	0.172	8.734	1
12	3.561	90.098	1.341	6.192	1
13	4.237	95.593	0.170	7.303	1
14	4.329	95.671	0.000	7.560	1
15	2.099	97.344	0.558	7.034	1
16	1.761	73.123	0.165	7.897	1
17	2.309	97.691	0.000	8.377	1
18	4.386	94.620	0.994	7.530	1
19	3.295	95.477	1.228	5.972	1
20	2.953	97.047	0.000	8.691	1
21	46.607	53.003	0.390	5.239	2
22	74.874	24.588	0.538	2.856	2
23	80.096	19.211	0.593	2.412	2
24	81.449	17.401	1.150	2.236	2
25	78.150	20.690	1.150	2.689	2
26	49.475	49.396	1.129	5.073	2
27	47.414	52.418	0.168	5.709	2
28	86.558	12.607	0.835	4.360	2
29	50.099	48.619	0.682	2.275	2
30	80.388	18.146	1.466	2.259	2
31	58.262	41.050	0.670	3.592	2
32	60.392	39.412	1.200	3.933	2
33	55.370	43.314	0.316	2.243	2
34	74.645	23.567	1.788	2.638	2
35	74.697	23.591	1.720	2.705	2
36	42.257	57.207	0.536	5.542	2
37	41.957	57.484	0.549	6.279	2
38	72.371	27.043	0.586	2.055	2
39	75.056	23.602	0.342	2.653	2
40	80.534	19.572	0.994	2.584	2
41	0.650	99.350	0.000	7.908	3
42	0.600	99.400	0.000	7.254	3
43	0.550	99.450	0.000	6.337	3
44	0.925	99.075	0.000	6.895	3
45	0.704	99.296	0.000	7.228	3
46	1.279	98.722	0.000	4.300	3
47	1.469	98.531	0.000	7.524	3
48	0.355	99.645	0.000	7.914	3
49	2.193	97.807	0.000	7.535	3
50	1.932	98.068	0.000	7.279	3
51	0.950	99.050	0.000	6.296	3
52	0.754	99.246	0.000	6.333	3
53	1.150	98.850	0.000	11.752	3
54	2.078	97.922	0.000	7.526	3
55	1.653	98.347	0.000	7.490	3
56	0.844	99.156	0.000	6.059	3



OBS	SAS ANALYSIS ON SEDIMENT DATA				LOCATION
	PERSAND	PERSLTCL	PERGRAV	PERORG	
57	1.668	78.332	0.000	6.086	3
58	1.801	78.178	0.021	6.997	3
59	1.087	98.913	0.000	6.629	3
60	2.504	97.391	0.105	5.854	3

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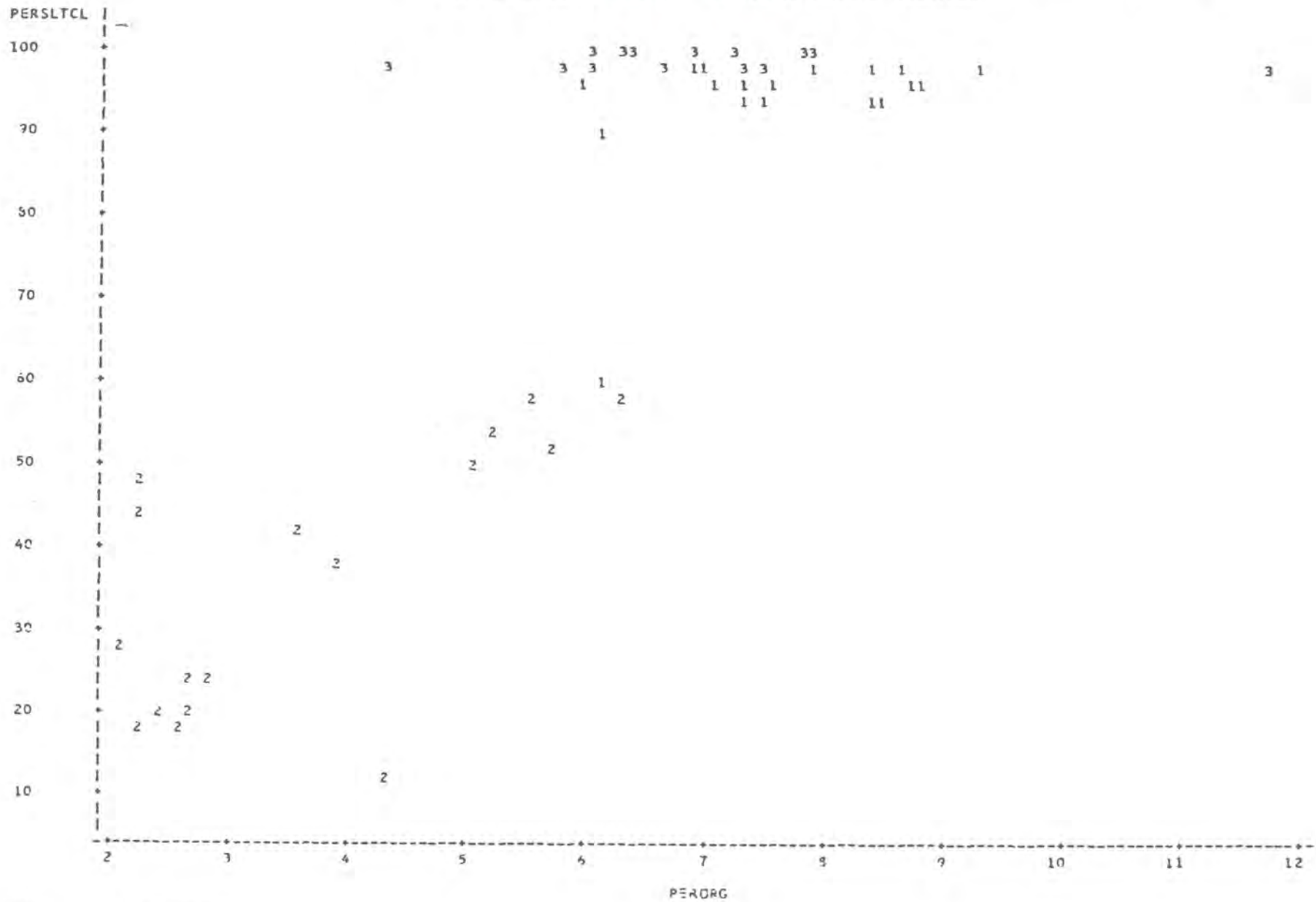




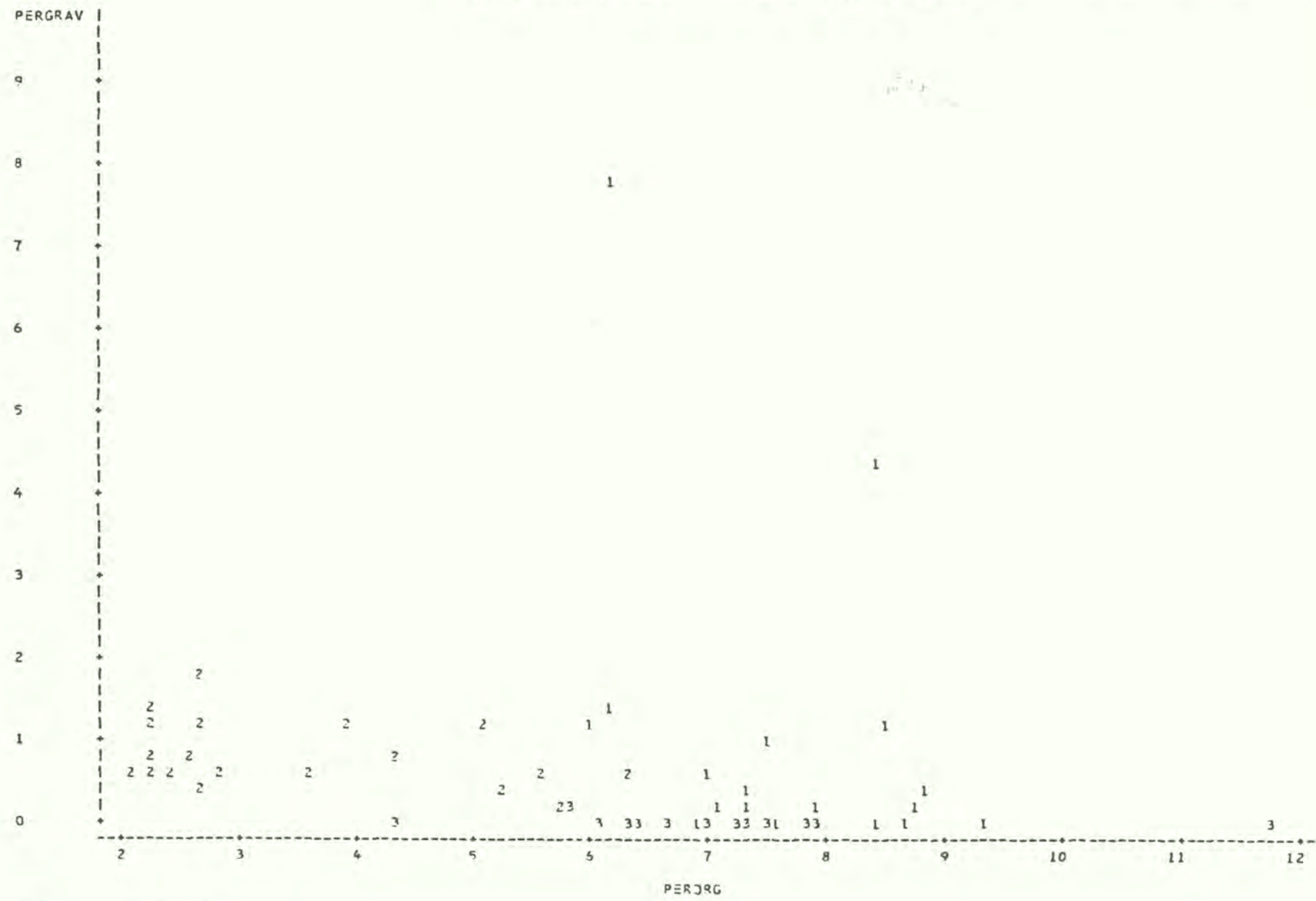


SAS ANALYSIS ON SEDIMENT DATA  
 PLOT OF PERSLTCL\*PERORG SYMBOL IS VALUE OF LOCATION

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NOTE: 11 OBS HIDDEN



NOTE: 10 OBS HIDDEN

SAS ANALYSIS ON SEDIMENT DATA  
 PRINCIPAL COMPONENT ANALYSIS

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60 OBSERVATIONS  
 4 VARIABLES

SIMPLE STATISTICS

	PERSAND	PERSLTCL	PERGRAV	PERORG
MEAN	23.93125	75.45463	0.614950	6.079983
ST DEV	31.21103	31.55473	1.174217	2.260263

COVARIANCES

	PERSAND	PERSLTCL	PERGRAV	PERORG
PERSAND	974.13	-984.2	10.077	-60.75
PERSLTCL	-984.2	995.7	-11.46	61.266
PERGRAV	10.077	-11.46	1.3788	-5.152
PERORG	-60.75	61.266	-5.152	5.1088

TOTAL VARIANCE = 1976.315

	EIGENVALUE	DIFFERENCE	PROPORTION	CUMULATIVE
PRIN1	1973.098	1971.150	0.998	0.998
PRIN2	1.943	0.577	0.001	0.999
PRIN3	1.270	1.270	0.001	1.000
PRIN4	0.000	.	0.000	1.000

EIGENVECTORS

	PRIN1	PRIN2	PRIN3	PRIN4
PERSAND	-.702525	-.368596	0.148886	0.577315
PERSLTCL	0.710263	-.393550	0.085707	0.577315
PERGRAV	-.007729	0.752076	-.234269	0.577421
PERORG	0.043400	0.237187	0.956373	-.000170



SAS ANALYSIS ON SEDIMENT DATA  
 PRINCIPAL COMPONENT ANALYSIS

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60 OBSERVATIONS  
 4 VARIABLES

SIMPLE STATISTICS

	PERSAND	PERSLTCL	PERGRAV	PERORG
MEAN	23.73125	75.45463	0.614950	6.079983
ST DEV	31.21193	31.55470	1.174217	2.260263

CORRELATIONS

	PERSAND	PERSLTCL	PERGRAV	PERORG
PERSAND	1.0000	-.9994	0.2750	-.8611
PERSLTCL	-.9994	1.0000	-.3092	0.8590
PERGRAV	0.2750	-.3092	1.0000	-.1941
PERORG	-.8611	0.8590	-.1941	1.0000

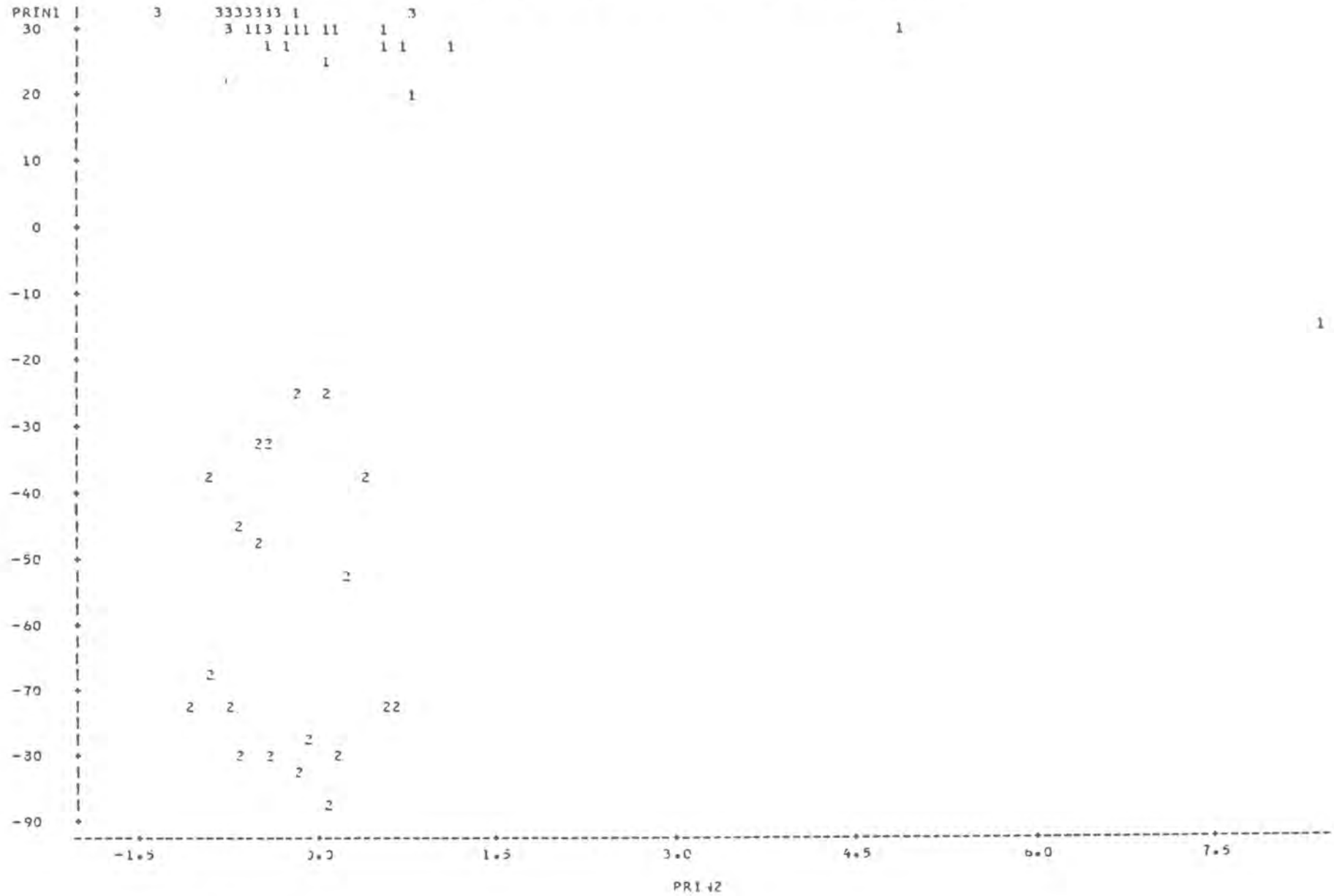
	EIGENVALUE	DIFFERENCE	PROPORTION	CUMULATIVE
PRIN1	2.920733	2.019341	0.730133	0.730133
PRIN2	0.900892	0.722518	0.225223	0.955406
PRIN3	0.178375	0.173375	0.044594	1.000000
PRIN4	0.000000		0.000000	1.000000

EIGENVECTORS

	PRIN1	PRIN2	PRIN3	PRIN4
PERSAND	-.572831	-.116183	0.405131	0.702978
PERSLTCL	0.575171	0.079856	-.397286	0.710719
PERGRAV	-.228839	0.969759	-.091744	0.026452
PERORG	0.537211	0.204422	0.818301	-.000015

SAS ANALYSIS OF SEDIMENT DATA  
 PLOT OF PRIN1\*PRIN2 SYMBOL IS VALUE OF LOCATION

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NOTE: 11 OBS HIDDEN



EIGENVALUES OF THE COVARIANCE MATRIX

EIGENVALUE	DIFFERENCE	PROPORTION	CUMULATIVE
1973.098	1971.150	0.998	0.998
1.948	0.677	0.001	0.999
1.270	1.270	0.001	1.000
0.000	.	0.000	1.000

ROOT-MEAN-SQUARE TOTAL-SAMPLE STANDARD DEVIATION = 22.2279  
 ROOT-MEAN-SQUARE DISTANCE BETWEEN OBSERVATIONS = 44.4558

NUMBER OF CLUSTERS	FREQUENCY OF NEW CLUSTER	RMS STD OF NEW CLUSTER	SEMI-PARTIAL R-SQUARED	R-SQUARED	APPROXIMATE EXPECTED R-SQUARED	CUBIC CLUSTERING CRITERION
10	8	1.57967	0.000335	0.998145	0.992461	7.7405
9	31	0.927458	0.000383	0.997762	0.990491	8.0046
8	6	2.02742	0.000597	0.997164	0.967673	8.1599
7	6	2.70773	0.000976	0.995139	0.923463	8.1911
6	39	1.47425	0.001349	0.994840	0.976815	8.4546
5	11	2.93514	0.002092	0.992748	0.965504	8.8987
4	7	4.24037	0.002444	0.990304	0.944134	10.2639
3	10	5.85681	0.006683	0.983621	0.896731	11.4793
2	21	11.5844	0.078526	0.905095	0.757624	7.1540
1	60	22.2279	0.905095	0.000000	0.000000	0.0000

SAS ANALYSIS ON SEDIMENT DATA

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OBS	_NAME_	_PARENT_	_NCL_	_FREQ_	_RMSSTD_	_DIST_	_AVLINK_	_SPRSQ_
1		3	CL59	60	1	0.000000	0.0000000	0.0000000
2		3	CL59	60	1	0.000000	0.0000000	0.0000000
3		1	CL59	60	1	0.000000	0.0000000	0.0000000
4		1	CL58	60	1	0.000000	0.0000000	0.0000000
5		2	CL57	60	1	0.000000	0.0000000	0.0000000
6		2	CL57	60	1	0.000000	0.0000000	0.0000000
7		3	CL56	60	1	0.000000	0.0000000	0.0000000
8		3	CL56	60	1	0.000000	0.0000000	0.0000000
9		1	CL55	60	1	0.000000	0.0000000	0.0000000
10		1	CL55	60	1	0.000000	0.0000000	0.0000000
11		3	CL54	60	1	0.000000	0.0000000	0.0000000
12		3	CL54	60	1	0.000000	0.0000000	0.0000000
13		3	CL53	60	1	0.000000	0.0000000	0.0000000
14		3	CL53	60	1	0.000000	0.0000000	0.0000000
15		3	CL52	60	1	0.000000	0.0000000	0.0000000
16		3	CL52	60	1	0.000000	0.0000000	0.0000000
17		3	CL51	60	1	0.000000	0.0000000	0.0000000
18		3	CL51	60	1	0.000000	0.0000000	0.0000000
19		1	CL50	60	1	0.000000	0.0000000	0.0000000
20		1	CL50	60	1	0.000000	0.0000000	0.0000000
21	CL59		CL49	59	2	0.0229445	0.0013734	0.0018734
22		3	CL49	60	1	0.000000	0.0000000	0.0000000
23		3	CL48	60	1	0.000000	0.0000000	0.0000000
24	CL53		CL48	53	2	0.100067	0.0063665	0.0063665
25		1	CL47	60	1	0.000000	0.0000000	0.0000000
26	CL55		CL47	55	2	0.075021	0.0048367	0.0048367

2.97420E-03  
3.43574E-07  
1.9E250E-07

OBS	_RSQ_	_ERSQ_	_RATIO_	_LOGR_	_CCC_	PER SAND	PER SILTCL	PER GRAV	PER ORG	LOCATION
1	1.00000	*	*	*	*	0.6500	99.3500	0.00000	7.90800	3
2	1.00000	*	*	*	*	0.6000	99.4000	0.00000	7.86400	3
3	1.00000	*	*	*	*	2.5200	97.4900	0.00000	6.87900	1
4	1.00000	*	*	*	*	2.4600	97.4670	0.07300	6.89400	1
5	1.00000	*	*	*	*	74.6450	23.5670	1.78800	2.63800	2
6	1.00000	*	*	*	*	74.6990	23.5910	1.72000	2.70500	2
7	1.00000	*	*	*	*	2.1930	97.8070	0.00000	7.53500	3
8	1.00000	*	*	*	*	2.0790	97.9220	0.00000	7.52600	3
9	1.00000	*	*	*	*	2.5650	96.9630	0.17200	8.73400	1
10	1.00000	*	*	*	*	2.7530	97.0470	0.00000	8.69100	1
11	1.00000	*	*	*	*	1.4690	98.9310	0.00000	7.52400	3
12	1.00000	*	*	*	*	1.3530	98.3470	0.00000	7.49000	3
13	1.00000	*	*	*	*	0.9600	99.0050	0.03500	6.39600	3
14	1.00000	*	*	*	*	1.0870	98.9130	0.00000	6.62900	3
15	1.00000	*	*	*	*	0.7540	99.1890	0.05800	5.33300	3
16	1.00000	*	*	*	*	0.8440	99.1260	0.03000	6.06200	3
17	1.00000	*	*	*	*	1.8320	98.1690	0.00000	7.29900	3
18	1.00000	*	*	*	*	1.8010	98.1780	0.02100	6.99700	3
19	1.00000	*	*	*	*	4.2370	95.5930	0.17000	7.30300	1
20	1.00000	*	*	*	*	4.3290	95.6710	0.00000	7.56000	1
21	1.00000	1.00000	14.3934	2.66600	25.2623	0.6250	99.3750	0.00000	7.88600	*
22	1.00000	*	*	*	*	0.3550	99.1450	0.00000	7.91400	3
23	1.00000	*	*	*	*	0.7850	97.0150	0.00000	6.89500	3
24	1.00000	0.99999	21.0662	3.04707	28.9789	1.0235	99.9590	0.01750	6.51250	*
25	1.00000	*	*	*	*	3.1170	96.6920	0.19100	8.75900	1
26	1.00000	0.99999	25.9870	3.25347	30.9235	2.9290	97.0050	0.08600	8.71250	*

SAS ANALYSIS ON SEDIMENT DATA

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OBS	_NAME_	_PARENT_	_NCL_	_FREQ_	_RMSSTD_	_DIST_	_AVLINK_	_SPRSQ_
27		3	CL46	60	1	0.000000	0.0000000	0.000000000
28	CL52		CL46	52	2	0.101827	0.0064787	0.0000003557
29		1	CL45	60	1	0.000000	0.0000000	0.000000000
30	CL54		CL45	54	2	0.092782	0.0059031	0.0000002953
31	CL58		CL44	58	2	0.034138	0.0021719	0.0000000400
32		1	CL44	60	1	0.000000	0.0000000	0.000000000
33		?	CL43	60	1	0.000000	0.0000000	0.000000000
34		2	CL43	60	1	0.000000	0.0000000	0.000000000
35	CL49		CL42	49	3	0.096517	0.0073438	0.0000006094
36		3	CL42	60	1	0.000000	0.0000000	0.000000000
37		1	CL41	60	1	0.000000	0.0000000	0.000000000
38	CL47		CL41	47	3	0.125500	0.0088368	0.0000009824
39	CL56		CL40	56	2	0.057598	0.0036639	0.0000001138
40	CL51		CL40	51	2	0.107293	0.0068267	0.0000003949
41		2	CL39	60	1	0.000000	0.0000000	0.000000000
42		2	CL39	60	1	0.000000	0.0000000	0.000000000
43	CL46		CL38	46	3	0.138634	0.0092314	0.0000009629
44	CL48		CL38	48	3	0.132703	0.0087477	0.0000008647
45	CL44		CL37	44	3	0.198147	0.0153251	0.0000002658
46		1	CL37	60	1	0.000000	0.0000000	0.000000000
47	CL45		CL36	45	3	0.175870	0.0127150	0.0000008268
48	CL40		CL36	40	4	0.183267	0.0131881	0.00000029479
49		1	CL35	60	1	0.000000	0.0000000	0.000000000
50		1	CL35	60	1	0.000000	0.0000000	0.000000000
51		2	CL34	60	1	0.000000	0.0000000	0.000000000
52		2	CL34	60	1	0.000000	0.0000000	0.000000000

OBS	_RSJ_	_EPSO_	_RATIO_	_LOGR_	_CCC_	PERSAND	PERSLTCL	PERGRAV	PEPORG	LOCATION
27	1.00000	.	.	.	.	0.5500	99.4500	0.00000	6.33700	3
28	1.00000	0.99997	20.9430	3.04130	29.8235	0.7990	99.1570	0.04400	6.20050	.
29	1.00000	.	.	.	.	1.7610	98.1230	0.16600	7.89700	1
30	1.00000	0.99999	22.4683	3.11210	29.4892	1.5610	98.4390	0.00000	7.50700	.
31	1.00000	1.00000	24.8404	3.21247	30.4397	2.4900	97.4735	0.03650	6.98650	.
32	1.00000	.	.	.	.	2.0780	97.3440	0.55900	7.03400	1
33	1.00000	.	.	.	.	80.3850	18.1460	1.46600	2.25900	2
34	1.00000	.	.	.	.	80.5340	18.5720	0.89400	2.58400	2
35	1.00000	0.99994	20.3204	3.01162	29.5185	0.7917	99.2983	0.00000	7.89533	.
36	1.00000	.	.	.	.	0.7040	99.2960	0.00000	7.22800	3
37	1.00000	.	.	.	.	2.4500	97.1500	0.00000	9.32500	1
38	1.00000	0.99992	19.2251	2.90290	22.4603	2.9753	95.7007	0.12100	8.72767	.
39	1.00000	0.99999	30.0937	3.40449	37.2593	2.1355	97.5645	0.00000	7.53050	.
40	1.00000	0.99996	21.0503	3.04694	29.9725	1.9165	99.1730	0.01050	7.14750	.
41	1.00000	.	.	.	.	42.2570	57.2070	0.53600	5.54200	2
42	1.00000	.	.	.	.	41.9470	57.4340	0.54900	6.29900	2
43	0.99999	0.99990	17.7694	2.87748	22.2648	0.7160	99.2547	0.02933	6.24600	.
44	1.00000	0.99993	19.8949	2.93989	22.7392	1.00107	99.9777	0.01167	6.64000	.
45	0.99994	0.99987	13.3097	2.59947	20.0290	2.3593	97.4303	0.21033	6.93567	.
46	1.00000	.	.	.	.	2.9740	96.8200	0.28500	7.05900	1
47	0.99997	0.99989	15.4242	2.74932	21.2677	1.5277	98.3377	0.05533	7.63700	.
48	0.99993	0.99978	10.4317	2.34435	18.1462	1.9750	98.0137	0.00525	7.33900	.
49	1.00000	.	.	.	.	4.0227	94.2500	1.12800	8.47900	1
50	1.00000	.	.	.	.	4.3950	94.6200	0.99400	7.53000	1
51	1.00000	.	.	.	.	46.6070	59.0030	0.39000	5.23800	2
52	1.00000	.	.	.	.	47.4140	52.4190	0.16800	5.70900	2

SAS ANALYSIS ON SEDIMENT DATA

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ONS	_NAME_	_PARENT_	_NCL_	_FREQ_	_RMSSTO_	_DIST_	_AVLINK_	_SPRSQ_	_RSC_	
53		2	CL33	60	1	0.00000	0.000000	0.000000	0.000000000	1.00000
54	CL43		CL33	43	2	0.28187	0.017933	0.017933	0.000002726	0.99999
55		3	CL32	60	1	0.00000	0.000000	0.000000	0.000000000	1.00000
56		3	CL32	60	1	0.00000	0.000000	0.000000	0.000000000	1.00000
57	CL41		CL31	41	4	0.19655	0.015091	0.015779	0.000002895	0.99998
58		1	CL31	60	1	0.00000	0.000000	0.000000	0.000000000	1.00000
59		2	CL30	60	1	0.00000	0.000000	0.000000	0.000000000	1.00000
50		2	CL30	60	1	0.00000	0.000000	0.000000	0.000000000	1.00000
61	CL35		CL29	35	2	0.37269	0.023711	0.023711	0.000004765	0.99996
62		1	CL29	60	1	0.00000	0.000000	0.000000	0.000000000	1.00000
63	CL31		CL28	31	5	0.29731	0.024583	0.025743	0.000008194	0.99993
64		1	CL28	60	1	0.00000	0.000000	0.000000	0.000000000	1.00000
65	CL37		CL27	37	4	0.26195	0.018537	0.019914	0.000004368	0.99997
66	CL32		CL27	32	2	0.45404	0.028888	0.028888	0.000007072	0.99994
67	CL33		CL26	33	3	0.36206	0.023553	0.025203	0.000006269	0.99995
68		2	CL26	60	1	0.00000	0.000000	0.000000	0.000000000	1.00000
69	CL30		CL25	30	2	0.55233	0.035173	0.035173	0.000010484	0.99992
70	CL57		CL25	57	2	0.03909	0.002487	0.002487	0.000000052	1.00000
71	CL50		CL24	50	2	0.11699	0.007443	0.007443	0.000000470	1.00000
72		1	CL24	60	1	0.00000	0.000000	0.000000	0.000000000	1.00000
73	CL42		CL23	42	4	0.18451	0.015011	0.015424	0.000002865	0.99996
74	CL38		CL23	38	6	0.19663	0.012707	0.014531	0.000004105	0.99997
75	CL27		CL22	27	6	0.42943	0.027325	0.032970	0.000017497	0.99987
76	CL36		CL22	36	7	0.22221	0.012561	0.015827	0.000004584	0.99996
77		2	CL21	60	1	0.00000	0.000000	0.000000	0.000000000	1.00000
78		2	CL21	60	1	0.00000	0.000000	0.000000	0.000000000	1.00000

ONS	_ERSQ_	_RATIO_	_LOGR_	_CCC_	PERSAND	PEPSLTCL	PERGRAV	PERORG	LOCATION
53	.	.	.	.	80.0960	19.2110	0.69300	2.4120	2
54	0.99995	12.0217	2.48571	19.2424	80.4610	13.3590	1.13000	2.4215	.
55	.	.	.	.	1.6680	93.3320	0.00000	6.0860	3
56	.	.	.	.	2.5040	97.3910	0.10500	5.8540	3
57	0.99930	10.7159	2.37182	18.3544	2.9463	96.9630	0.09075	8.8770	.
58	.	.	.	.	2.3090	97.6910	0.00000	8.3770	1
59	.	.	.	.	74.8740	24.5880	0.53800	2.8560	2
60	.	.	.	.	76.0560	23.6020	0.34200	2.6030	2
61	0.99961	9.1094	2.20731	17.1005	4.5040	94.4350	1.00100	8.0045	.
62	.	.	.	.	5.6950	93.8410	0.45400	7.3300	1
63	0.99943	8.2129	2.10570	11.5274	2.8188	97.1096	0.07250	3.7770	.
64	.	.	.	.	3.5640	95.9440	0.49200	8.8400	1
65	0.99969	7.4938	2.24254	17.4109	2.4930	97.2777	0.22925	6.9662	.
66	0.99943	8.4059	2.13353	13.5554	2.0960	97.8615	0.35250	5.9700	.
67	0.99953	8.7130	2.15482	16.7530	30.3393	18.6430	1.71767	2.4183	.
68	.	.	.	.	81.4490	17.4010	1.15000	2.2360	2
69	0.99937	7.3209	2.05679	11.2404	75.4650	24.0950	0.44000	2.7595	.
70	1.00000	32.2953	3.47494	32.2268	74.6720	23.5740	1.75400	2.6715	.
71	0.99995	20.9569	3.04247	29.5305	4.2930	95.6220	0.08500	7.4315	.
72	.	.	.	.	3.2950	95.4770	1.22800	5.9720	1
73	0.99992	11.2003	2.41594	18.6953	0.7023	99.2977	0.00000	7.7295	.
74	0.99972	9.7778	2.28235	17.6539	0.8633	97.1162	0.02050	6.4430	.
75	0.99915	6.4754	1.86471	10.2297	2.3573	97.4723	0.17033	6.5342	.
76	0.99965	9.2531	2.22550	17.2252	1.3267	98.1537	0.02671	7.4667	.
77	.	.	.	.	49.4750	49.3950	1.12900	5.0730	2
78	.	.	.	.	50.6970	43.6140	0.68200	2.2750	2

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OBS	_NAME_	_PARENT_	_NCL_	_FREQ_	_RMSSTO_	_DIST_	_AVLINK_	_SPRSQ_	_RSQ_
79	CL29	CL20	29	3	0.5335	0.03615	0.03804	0.000015	0.99991
80	CL24	CL20	24	3	0.6137	0.04738	0.04753	0.000025	0.99981
81		2 CL19	60	1	0.0000	0.00000	0.00000	0.000000	1.00000
82		2 CL19	60	1	0.0000	0.00000	0.00000	0.000000	1.00000
83	CL23	CL18	23	10	0.3833	0.02943	0.03135	0.000035	0.99977
84		3 CL18	60	1	0.0000	0.00000	0.00000	0.000000	1.00000
85	CL26	CL17	26	4	0.5137	0.03731	0.04008	0.000018	0.99985
86		2 CL17	60	1	0.0000	0.00000	0.00000	0.000000	1.00000
87	CL28	CL16	28	6	0.3976	0.03253	0.03467	0.000015	0.99989
88		3 CL16	60	1	0.0000	0.00000	0.00000	0.000000	1.00000
89	CL25	CL15	25	4	0.5674	0.03651	0.04055	0.000023	0.99983
90		2 CL15	60	1	0.0000	0.00000	0.00000	0.000000	1.00000
91	CL19	CL14	19	2	1.2223	0.07777	0.07777	0.000061	0.99959
92		2 CL14	60	1	0.0000	0.00000	0.00000	0.000000	1.00000
93	CL20	CL13	20	6	0.7288	0.04242	0.04198	0.000066	0.99964
94		1 CL13	60	1	0.0000	0.00000	0.00000	0.000000	1.00000
95	CL22	CL12	22	13	0.4470	0.02718	0.03369	0.000040	0.99973
96	CL18	CL12	18	11	0.5473	0.06113	0.06329	0.000058	0.99953
97	CL34	CL11	34	2	0.3976	0.02530	0.02530	0.000005	0.99995
98	CL21	CL11	21	2	1.1253	0.07160	0.07160	0.000043	0.99969
99	CL13	CL10	13	7	1.1333	0.10920	0.11323	0.000173	0.99938
100		1 CL10	60	1	0.0000	0.00000	0.00000	0.000000	1.00000
101	CL16	CL9	16	7	0.8349	0.08949	0.09097	0.000116	0.99934
102	CL12	CL9	12	24	0.6741	0.04143	0.05140	0.000173	0.99871
103	CL17	CL8	17	5	0.8864	0.07712	0.07968	0.000081	0.99945
104		2 CL8	60	1	0.0000	0.00000	0.00000	0.000000	1.00000

OBS	_ERSQ_	_RATIO_	_LNCR_	_CCC_	PER SAND	PER SLTCL	PER GRAY	PER ORG	LOCATION
79	0.999315	7.2338	1.97877	10.8341	4.9010	94.2370	0.86200	7.7797	.
80	0.999907	5.6574	1.73296	9.4929	3.9537	95.5903	0.46600	6.9450	.
81	.	.	.	.	58.2620	41.0630	0.57000	3.5920	2
82	.	.	.	.	60.3820	39.4120	1.20600	3.9330	2
83	0.998793	5.2800	1.66393	9.1160	0.7989	99.1888	0.01230	6.9572	.
84	.	.	.	.	1.2030	98.7920	0.00000	4.3000	3
85	0.999036	6.2186	1.82755	10.0097	30.6157	18.3325	1.05075	2.3728	.
86	.	.	.	.	78.1500	20.6900	1.16000	2.6899	2
87	0.999249	5.8474	1.92397	10.5343	2.9430	96.9145	0.14250	8.7375	.
88	.	.	.	.	1.1500	98.8500	0.00000	11.7620	3
89	0.999707	5.7143	1.77746	9.7355	75.0685	23.8345	1.09700	2.7155	.
90	.	.	.	.	72.3710	27.0430	0.58600	2.0550	2
91	0.998150	4.5183	1.50813	8.2684	57.3220	39.7400	0.93300	3.7625	.
92	.	.	.	.	55.8700	43.3140	0.91600	2.2430	2
93	0.998343	4.6125	1.52873	8.3900	4.4273	94.9087	0.66400	7.3623	.
94	.	.	.	.	1.5130	94.1700	4.31700	8.4510	1
95	0.979665	4.9639	1.60217	9.7790	2.0716	97.8392	0.09300	7.0825	.
96	0.997917	4.4596	1.49507	8.1995	0.8361	99.1527	0.01118	6.7156	.
97	0.997972	9.7311	2.18754	16.9495	47.0105	52.7105	0.27900	5.4715	.
98	0.998518	4.7444	1.55676	8.5327	30.0870	40.0075	0.90550	3.6740	.
99	0.995733	3.7932	1.33320	7.3307	4.0110	94.8031	1.18596	7.5179	.
100	.	.	.	.	3.5610	90.0990	1.34100	5.1720	1
101	0.997707	4.0557	1.40013	7.5845	2.8869	97.1710	0.12214	9.2124	.
102	0.994951	3.7130	1.36429	7.5090	1.5053	99.4412	0.05550	6.9143	.
103	0.997641	4.7096	1.46061	9.0132	80.1234	18.8040	1.07260	2.4360	.
104	.	.	.	.	89.5580	12.6070	0.83500	4.3500	2



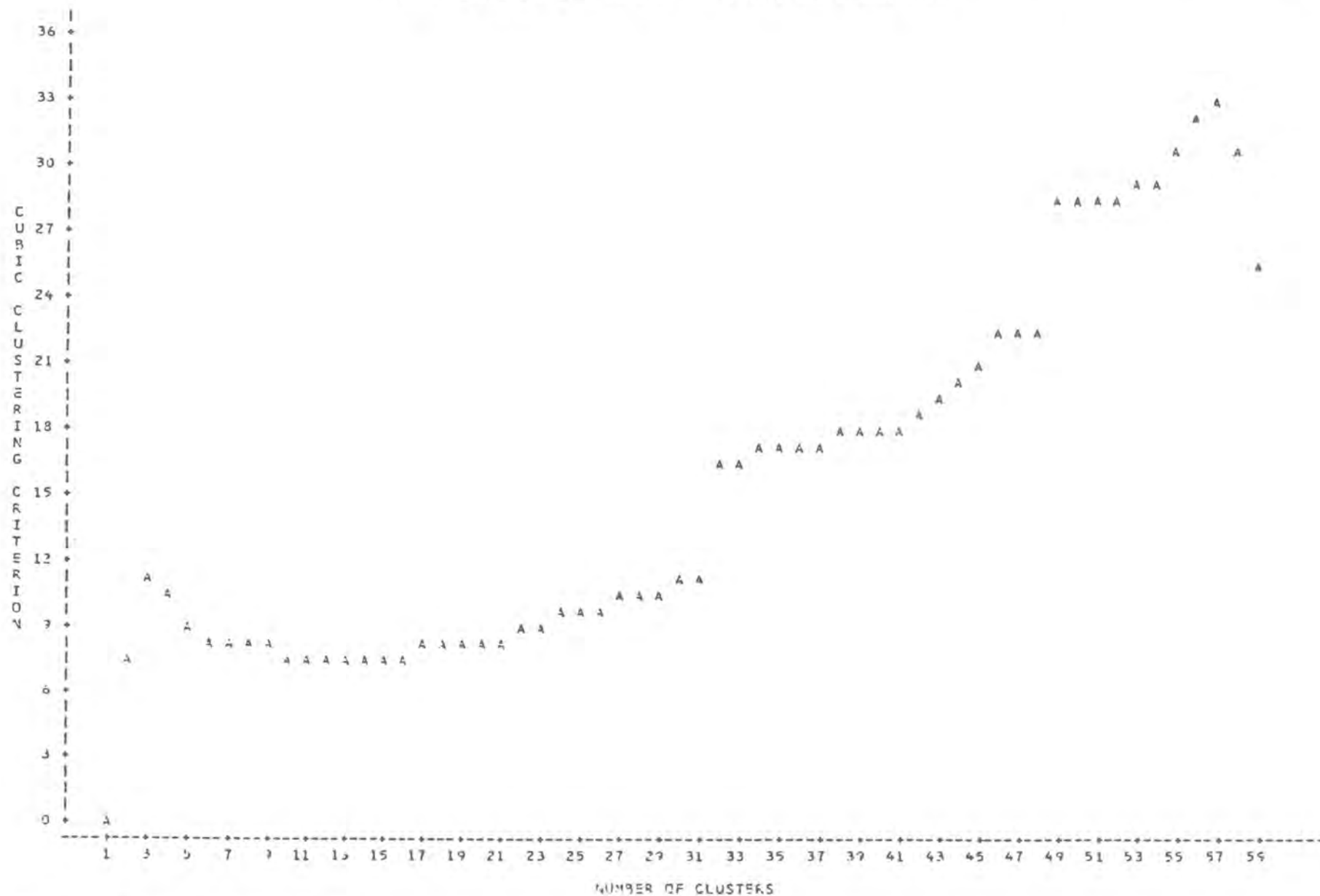
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OBS	_NAME_	_PARENT_	_NCL_	_FREQ_	_RMSSTD_	_DIST_	_AVLINK_	_SPRSQ_	_RSQ_
105	CL11	CL7	11	4	1.6453	0.11647	0.12250	0.000230	0.99848
106	CL39	CL7	39	2	0.3029	0.01927	0.01927	0.000003	0.99998
107	CL9	CL6	9	31	0.9275	0.06459	0.07914	0.000383	0.99776
108	CL10	CL6	10	8	1.5797	0.15026	0.15750	0.000335	0.99814
109	CL15	CL5	15	5	1.0746	0.09614	0.09865	0.000125	0.99921
110	CL8	CL5	8	6	2.0274	0.20563	0.20870	0.000597	0.99716
111		CL4	60	1	0.0000	0.00000	0.00000	0.000000	1.00000
112	CL7	CL4	7	6	2.7077	0.20778	0.21766	0.000976	0.99619
113	CL4	CL3	4	7	4.2409	0.41018	0.42493	0.002444	0.99030
114	CL14	CL3	14	3	1.7315	0.11691	0.12321	0.000154	0.99406
115	CL3	CL2	3	10	5.9568	0.43331	0.47224	0.006683	0.98362
116	CL5	CL2	5	11	2.9351	0.21273	0.23250	0.002092	0.99275
117	CL6	CL1	6	39	1.4742	0.11197	0.13645	0.001349	0.99484
118	CL2	CL1	2	21	11.5344	0.94047	0.98124	0.078526	0.90510
119	CL1		1	60	22.2279	1.97791	2.04331	0.905095	0.00000

OBS	_ERSQ_	_RATIO_	_LOGR_	_CCC_	PER SAND	PER SLTCL	PER GRAV	PER ORG	LOCATION
105	0.993897	4.0206	1.39144	7.6692	49.5497	50.8590	0.59225	4.57375	.
106	0.999749	10.2192	2.32427	17.9376	42.1120	57.3455	0.54250	5.92050	.
107	0.990471	4.2483	1.44652	8.0046	1.7721	99.1589	0.07055	7.43326	.
108	0.992461	4.0639	1.40213	7.7405	4.5798	94.2150	1.20525	7.35212	.
109	0.995900	3.9271	1.36790	7.5112	74.5290	24.4762	0.99480	2.53340	.
110	0.987673	4.3473	1.46956	8.1599	91.1958	17.7712	1.03300	2.75567	.
111	.	.	.	.	31.5430	60.7020	7.75000	6.19000	1
112	0.993463	4.3390	1.46764	8.1911	46.4032	53.0212	0.57567	5.02267	.
113	0.944134	5.7617	1.75123	10.2639	44.2810	54.1184	1.60057	5.13943	.
114	0.976377	3.8180	1.33974	7.3510	58.1713	40.9313	0.89733	3.25600	.
115	0.896731	6.3049	1.84133	11.4793	48.4431	50.1623	1.38960	4.60040	.
116	0.955504	4.7569	1.55959	8.9997	73.1655	20.8189	1.01564	2.67791	.
117	0.976815	4.4933	1.50259	3.4545	2.3491	97.3499	0.30331	7.41562	.
118	0.757624	2.5539	0.93761	7.1540	64.0143	34.7920	1.19371	3.59757	.
119	0.000000	1.0000	0.00000	0.0000	23.9312	75.4546	0.61495	6.07998	.

PLOT OF \_CCC\*\_NCL\_ LEGEND: A = 1 OBS, B = 2 OBS, ETC.



NOTE: 60 OBS HAD MISSING VALUES

EXAMPLE OF SELECTING THE "BEST" VARIABLES FROM A MULTIVARIATE DATA SET

N = 60 P = 4 CYCLE = 4 BRIEF = 1 CPCT = 100.

THE N SAMPLES-BY-P VARIABLES INPUT DATA ARE:

%SAND	%SILT-CLAY	%GRAVEL	%O.M.
2.950	97.150	0.000	9.325
4.622	94.250	1.123	8.479
3.554	95.944	0.492	8.840
3.117	96.092	0.191	9.758
1.513	94.170	4.317	8.451
46.607	53.003	0.390	5.238
74.874	24.598	0.533	2.856
80.095	19.211	0.693	2.412
81.449	17.401	1.150	2.236
78.150	20.590	1.160	2.689
0.650	99.350	0.000	7.908
0.600	99.400	0.000	7.964
0.550	97.450	0.000	6.337
0.995	99.015	0.000	6.895
0.704	99.296	0.000	7.228
2.520	97.480	0.000	6.879
2.394	96.320	0.286	7.058
2.460	97.467	0.073	6.894
31.548	60.702	7.750	6.190
5.695	93.641	0.464	7.330
49.475	49.395	1.129	5.073
47.414	52.413	0.168	5.709
80.553	12.507	0.835	4.360
50.699	43.619	0.682	2.275
90.338	18.146	1.466	2.259
1.209	98.792	0.000	4.300
1.469	93.531	0.000	7.524
0.355	99.145	0.000	7.914
2.193	97.807	0.000	7.535
1.132	92.166	0.000	7.293
2.865	90.763	0.172	3.734
3.561	90.095	1.341	5.192
4.237	95.593	0.170	7.303
4.329	95.671	0.000	7.560
2.099	97.344	0.558	7.034
52.262	41.063	0.670	3.592
60.382	33.412	1.206	3.933
55.370	43.314	0.816	2.243
74.645	23.567	1.791	2.639
74.699	23.931	1.720	2.705
0.760	99.005	0.000	6.390
0.754	99.196	0.053	6.333
1.150	98.050	0.000	11.762
2.078	97.422	0.000	7.526
1.553	99.347	0.000	7.490
1.761	95.123	0.165	7.697
2.307	97.691	0.000	5.377

%SAND	%SILT-CLAY	%GRAVEL	%M.
4.386	94.620	0.794	7.530
3.295	95.477	1.223	5.972
2.453	77.047	0.000	3.691
42.257	57.207	0.536	3.542
41.767	57.494	0.547	5.297
72.371	27.043	0.596	2.055
76.056	23.502	1.342	2.663
80.534	19.572	0.394	2.594
0.844	99.126	0.000	5.063
1.563	90.332	0.000	5.096
1.301	93.179	0.021	5.997
1.027	93.313	0.000	5.627
2.504	77.391	0.103	3.954

THE P-BY-P CORRELATION MATRIX IS:

1.0000	-0.7793	0.2750	-0.9511
-0.9793	1.0000	-0.3032	0.3590
0.2750	-0.3032	1.0000	-0.1941
-0.9511	0.3590	-0.1941	1.0000

THE VARIABLES (TOP) ARE ORDERED BY THE SUM-OF-SQUARED CRITERION (BOTTOM) :

2	1	4	3
2.8322	2.3153	2.5171	1.2099

1 VARIABLES & THEIR CORRELATIONS HAVE BEEN REMOVED.

THE P-BY-P CORRELATION MATRIX IS:

0.0013	0.0000	-0.0340	-0.0027
0.0000	0.0000	0.0000	0.0000
-0.0340	0.0000	0.9044	0.0715
-0.0027	0.0000	0.0715	0.2521

THE VARIABLES (TOP) ARE ORDERED BY THE SUM-OF-SQUARED CRITERION (BOTTOM) :

3	4	1	2
0.9242	0.0733	0.0012	0.0000

2 VARIABLES & THEIR CORRELATIONS HAVE BEEN REMOVED.

THE P-BY-P CORRELATION MATRIX IS:

0.0000	0.0000	-0.0000	0.0000
0.0000	0.0000	0.0000	0.0000
-0.0000	0.0000	0.0000	0.0000
0.0000	0.0000	0.0000	0.2564

THE VARIABLES (TOP) ARE ORDERED BY THE SUM-OF-SQUARED CRITERION (BOTTOM) :

4	1	3	2
0.0658	0.0000	0.0000	0.0000

3 VARIABLES & THEIR CORRELATIONS HAVE BEEN REMOVED.

THE P-PY-P CORRELATION MATRIX IS:

0.0000	0.0000	-0.0000	0.0000
0.0000	0.0000	0.0000	0.0000
-0.0000	0.0000	0.0000	0.0000
0.0000	0.0000	0.0000	0.0000

THE VARIABLES (TOP) ARE ORDERED BY THE SUM-OF-SQUARED CRITERION (BOTTOM) :

1	3	4	2
0.0000	0.0000	0.0000	0.0000

4 VARIABLES & THEIR CORRELATIONS HAVE BEEN REMOVED.

THE CORRELATION MATRIX IS NOW EXHAUSTED.

THE BEST ORDER IN WHICH TO CHOOSE VARIABLES, FOR MAXIMUM INFORMATION ABOUT STRUCTURE IN THE DATA SET, IS:

2	3	4	1
---	---	---	---

THE TRACES ASSOCIATED WITH THE RESIDUAL CORRELATION MATRICES ARE (BEGINNING WITH 0 VARIABLES REMOVED):

4.000	1.163	0.255	0.000
-------	-------	-------	-------

THE TRACES, AS A PERCENTAGE OF P, ARE:

100.0	29.2	6.4	0.0
-------	------	-----	-----

## 8. MULTIVARIATE ANALYSIS OF VARIANCE (MANOVA) AND DISCRIMINANT ANALYSIS (DA)

### 8.1 Introduction

We will use the data set 'SEDABC' again but this time our analysis will be based on the a priori partitioning of the samples into 3 groups (the 3 locations). Thus the data matrix leads to the MANOVA and DA

$$\begin{array}{c}
 n=60 \\
 \left[ \begin{array}{c}
 \text{---} \\
 \text{---} \\
 \text{---}
 \end{array} \right]
 \end{array}
 \begin{array}{l}
 p=4 \\
 n_1 = 20 \\
 n_2 = 20 \\
 n_3 = 20
 \end{array}$$

We will do the analysis in APL and MINITAB.  
In principle MANOVA and DA is simple:

- (1) We calculate the deviation squares and cross-products matrix for each of the 3 groups, and then we sum them (matrix addition). That is, we calculate  $W_1$ ,  $W_2$  and  $W_3$  and then  $W=W_1 + W_2 + W_3$ . Then we calculate the deviation squares and cross products matrix for all the data regardless of group membership, to get the T matrix. Then the among-group matrix A is obtained by  $A=T-W$ .
- (2) We can think of a ANOVA table, as in the univariate ANOVA:

Source	df	SS	MS
Among groups	$g-1$	$A = P \begin{bmatrix} P \\ \end{bmatrix}$	Among group pxp covariance matrix
Within groups	$n-g$	$W = P \begin{bmatrix} P \\ \end{bmatrix}$	Within group pxp covariance matrix
Total	$n-1$	$T = P \begin{bmatrix} P \\ \end{bmatrix}$	Total pxp covariance matrix

(3) In MANOVA we test for group differences by evaluating the ratio of  $W$  to  $T$  and seeing whether it is significantly less than one, instead of evaluating the ratio of the among-group to the within-group variance and seeing whether it is greater than one as we do in the univariate ANOVA. To be specific we evaluate "Wilk's lambda" which is  $\Delta = |W|/|T|$ , which is the determinant of  $W$  divided by the determinant of  $T$ . One test is

$$X^2(p(g-1)df) = -(n-1-p+g/2)\log\Delta.$$

(4) If the null hypothesis  $H_0 =$  "groups have similar mean vectors" is rejected, and we conclude that the groups are different, then we proceed to do a DA to describe that differences. In matrix algebra the calculations for a DA are simple: we find the roots and vectors of  $W^{-1}A$ . That is, we invert the matrix  $W$  to obtain  $W^{-1}$ . Then we do the matrix multiplication to obtain  $W^{-1}A$ . Then we find the roots and vectors of the  $W^{-1}A$  matrix (which is not symmetric). The vectors contain the coefficients in the "discriminant functions" which describe the relationships between the new rotated axes and the original axes, much as the principal component vectors did. However in PCA we were attempting to "most efficiently" describe the

variation and covariation in the data, and to do that we found the roots and vectors of either the covariance or the correlation matrix. In DA we want to most efficiently describe the ratio of among-group to within-group variation and covariation, and to do this we find the roots and vectors of  $W^{-1}A$ .

## 8.2 Assignment

To save you the time and bother, here are the matrices and parameters:

$$g = 3 \qquad n = 60 \qquad p = 4$$

$$W1 = \begin{bmatrix} 796.46 & -987.11 & 190.50 & -51.69 \\ -987.11 & 1244.40 & -257.10 & 62.99 \\ 190.50 & -257.10 & 66.56 & 18.83 \\ -51.69 & 62.99 & -11.29 & 18.83 \end{bmatrix}$$

$$W2 = \begin{bmatrix} 4358.21 & -4413.81 & 55.60 & -284.32 \\ -4413.81 & 4473.19 & -59.38 & 289.04 \\ 55.60 & -59.38 & 3.78 & -4.72 \\ -284.32 & 289.04 & -4.72 & 36.77 \end{bmatrix}$$

$$W3 = \begin{bmatrix} 6.49 & -6.58 & 0.085 & -0.97 \\ -6.58 & 6.67 & -0.099 & 1.20 \\ 0.085 & -0.099 & 0.014 & -0.23 \\ -0.97 & 1.20 & -0.23 & 37.82 \end{bmatrix}$$

$$T = \begin{bmatrix} 57473.6 & -58069.2 & 594.53 & -3584.21 \\ -58069.2 & 58746.3 & -675.90 & 3614.69 \\ 594.53 & -675.90 & 81.35 & -30.40 \\ -3584.21 & 3614.69 & -30.40 & 301.42 \end{bmatrix}$$

8.2.1 Calculate  $W$  and  $A$  using APL. Then calculate  $\Delta$  (use PDET to find the determinants) and then the  $X^2$  and the degrees of freedom. Calculate  $W^{-1}$  and the  $W^{-1}A$ . Then use GEIG to find at least the first two roots, and the associated vectors, of  $W^{-1}A$ .



### 8.2.2 Now do it all in SAS:

```
TITLE  MANOVA AND DA ON SEDIMENT DATA;  
DATA  SEDABC;  
INPUT PERSAND PERSLTCL PERGRAV PERORG LOCATION;  
CARDS;
```

(the SEDABC data go here - use the 'GET SEDABC DATA' command)

```
PROC  GLM; CLASS LOCATION;  
MODEL PERSAND PERSLTCL PERGRAV PERORG=LOCATION;  
MANOVA H=LOCATION/PRINTH PRINTE;  
PROC  CANDISC OUT=DISC; CLASS LOCATION;  
VAR  PERSAND PERSLTCL PERGRAV PERORG  
PROC  PLOT; PLOT CAN2*CAN1=LOCATION;
```

8.2.3 Try to interpret these results. Compare your APL results with the SAS results. Also compare this MANOVA/DFA analysis with the PCA analysis.

### 8.3. Job Listings & Outputs.

FILE: MANOVA SAS AL VM/SP - CONVERSATIONAL MONITOR SYSTEM

PAGE 001

TITLE SAS ANALYSIS JM SEDIMENT DATA;  
 DATA SEDJBC;  
 INPUT PERSANO PERSLTCL PERGRAV PERORG LOCATION;  
 CAFOS;

2.850	97.15J	0.0	9.325	1.
4.622	94.250	1.124	3.477	1.
3.564	95.944	0.492	3.840	1.
3.117	96.592	0.191	8.753	1.
1.513	94.170	4.317	9.451	1.
2.52J	97.480	0.0	6.379	1.
2.374	96.320	0.235	7.053	1.
2.460	97.467	0.073	6.874	1.
31.543	60.702	7.750	6.190	1.
5.675	93.341	0.464	7.330	1.
2.365	96.753	0.172	8.734	1.
3.541	90.076	1.341	5.192	1.
4.237	95.593	0.170	7.303	1.
4.327	95.671	0.0	7.550	1.
2.393	97.344	0.553	7.034	1.
1.761	90.123	0.165	7.377	1.
2.309	97.091	0.0	8.377	1.
4.385	94.620	0.974	7.530	1.
3.295	95.477	1.223	5.972	1.
2.753	97.047	0.0	9.591	1.
46.607	53.003	0.370	5.232	2.
74.374	24.593	0.533	2.956	2.
80.096	17.211	0.693	2.412	2.
81.449	17.401	1.150	2.230	2.
73.150	20.690	1.160	2.589	2.
49.475	49.396	1.127	5.073	2.
47.414	52.418	0.163	5.707	2.
86.553	12.607	0.835	4.360	2.
50.699	43.619	0.592	2.275	2.
80.383	12.143	1.455	2.257	2.
59.262	41.063	0.670	3.572	2.
60.362	38.412	1.205	3.773	2.
55.37J	43.314	0.915	2.243	2.
74.645	23.567	1.737	2.633	2.
74.679	23.531	1.777	2.705	2.
42.257	57.207	0.536	5.542	2.
41.767	57.484	0.547	5.279	2.
72.371	27.743	0.595	2.055	2.
70.056	23.502	0.342	2.663	2.
80.534	13.372	0.874	2.584	2.
0.650	99.350	0.0	7.903	3.
0.600	99.400	0.0	7.964	3.
0.550	99.450	0.0	5.337	3.
0.795	99.015	0.0	6.975	3.
0.774	99.290	0.0	7.228	3.
1.203	98.792	0.0	4.300	3.
1.467	98.531	0.0	7.524	3.
0.355	99.145	0.0	7.914	3.
2.193	97.307	0.0	7.535	3.
1.332	98.163	0.0	7.278	3.
0.960	99.000	0.035	6.376	3.

0.754	99.188	0.058	6.333 3.
1.150	99.850	0.0	11.762 3.
2.078	97.722	0.0	7.526 3.
1.653	98.347	0.0	7.470 3.
0.844	99.126	0.030	0.068 3.
1.668	98.332	0.0	6.086 3.
1.801	93.178	0.021	6.797 3.
1.087	98.913	0.0	6.629 3.
2.504	97.391	0.105	5.854 3.

```
PROC GLM; CLASS LOCATION;  
  MODEL PERSAND PERSLTCL PERGRAV PERORG=LOCATION;  
  MANOVA H=LOCATION/PRINTH PRINTE;  
PROC CANDISC OUT=DISC; CLASS LOCATION;  
  VAR PERSAND PERSLTCL PERGRAV PERORG;  
PROC PLOT; PLOT CAN2*CAN1=LOCATION;
```

SAS ANALYSIS ON SEDIMENT DATA  
 GENERAL LINEAR MODELS PROCEDURE

13:57 THURSDAY, MAY 9, 1985 2

DEPENDENT VARIABLE: PERSAND

SOURCE	DF	SUM OF SQUARES	MEAN SQUARE	F VALUE	PR > F	R-SQUARE	C.V.
MODEL	2	52312.42925440	26156.21462720	288.87	0.0001	0.910199	39.7622
ERROR	57	5161.15760285	90.54662461				PERSAND MEAN
CORRECTED TOTAL	59	57473.58685725			9.51559901		23.93125000

SOURCE	DF	TYPE I SS	F VALUE	PR > F	DF	TYPE III SS	F VALUE	PR > F
LOCATION	2	52312.42925440	288.87	0.0001	2	52312.42925440	288.87	0.0001

SAS ANALYSIS ON SEDIMENT DATA  
 GENERAL LINEAR MODELS PROCEDURE

13:57 THURSDAY, MAY 9, 1985 3

DEPENDENT VARIABLE: PERSLTCL

SOURCE	DF	SUM OF SQUARES	MEAN SQUARE	F VALUE	PR > F	R-SQUARE	C.V.
MODEL	2	53021.98620823	26510.99310412	263.99	0.0001	0.902559	13.2812
ERROR	57	5724.26648170	100.42572775				
CORRECTED TOTAL	59	58746.25268993					
						ROOT MSE	PERSLTCL MEAN
						10.02126378	75.45463333

SOURCE	DF	TYPE I SS	F VALUE	PR > F	DF	TYPE III SS	F VALUE	PR > F
LOCATION	2	53021.98620823	263.99	0.0001	2	53021.98620823	263.99	0.0001

SAS ANALYSIS ON SEDIMENT DATA  
 GENERAL LINEAR MODELS PROCEDURE

13:57 THURSDAY, MAY 9, 1985 4

DEPENDENT VARIABLE: PERGRAV

SOURCE	DF	SUM OF SQUARES	MEAN SQUARE	F VALUE	PR > F	R-SQUARE	C.V.
MODEL	2	10.99139110	5.49569553	4.45	0.0160	0.135115	180.5659
ERROR	57	70.35588975	1.23433160				PERGRAV MEAN
CORRECTED TOTAL	59	81.34728085			1.11103468		0.61495000

SOURCE	DF	TYPE I SS	F VALUE	PR > F	DF	TYPE III SS	F VALUE	PR > F
LOCATION	2	10.99139110	4.45	0.0160	2	10.99139110	4.45	0.0160



SAS ANALYSIS ON SEDIMENT DATA  
 GENERAL LINEAR MODELS PROCEDURE

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E = ERROR SSCP MATRIX

DF=57	PERSAND	PERSLTCL	PERGRAV	PERORG
PERSAND	5161.15760235	-5407.49233490	246.17888955	-336.97633555
PERSLTCL	-5407.47238490	5724.26648170	-316.57580430	353.23004425
PERGRAV	246.17888955	-316.57580430	70.35688975	-16.24259370
PERORG	-336.97633555	353.23004425	-16.24259370	93.41757835

PARTIAL CORRELATION COEFFICIENTS FROM THE ERROR SSCP MATRIX / PROB > |R|

DF=56	PERSAND	PERSLTCL	PERGRAV	PERORG
PERSAND	1.000000 0.0000	-0.994351 0.0001	0.403530 0.0015	-0.435301 0.0001
PERSLTCL	-0.994351 0.0001	1.000000 0.0000	-0.473844 0.0001	0.483040 0.0001
PERGRAV	0.403530 0.0015	-0.473844 0.0001	1.000000 0.0000	-0.200349 0.1316
PERORG	-0.435301 0.0001	0.483040 0.0001	-0.200349 0.1316	1.000000 0.0000



SAS ANALYSIS ON SEDIMENT DATA  
GENERAL LINEAR MODELS PROCEDURE

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H = TYPE III SSEC MATRIX FOR: LOCATION

DF=2	PERSAND	PERSLTCL	PERGRAV	PERORG
PERSAND	52312.42725440	-52561.72757150	348.34769720	-3247.23085120
PERSLTCL	-52561.72757150	53021.79520823	-359.32151090	3261.46488833
PERGRAV	348.34769720	-359.32151090	10.99139110	-14.15430135
PERORG	-3247.23085120	3261.46488833	-14.15430135	208.00093453

CHARACTERISTIC ROOTS AND VECTORS OF: E INVERSE \* H, WHERE H = TYPE III SSEC MATRIX FOR: LOCATION E = ERROR SSEC MATRIX

CHARACTERISTIC ROOT	PERCENT	CHARACTERISTIC VECTOR V*EV=1			
		PERSAND	PERSLTCL	PERGRAV	PERORG
11.56690295	77.92	-0.34202934	-0.35741547	-0.40504314	0.00248915
0.25057483	2.13	-8.01961179	-8.01460641	-9.10501852	-0.07043560
0.00000000	0.00	13.79276148	18.97411355	18.95771705	-0.03019553
0.00000000	0.00	0.05834420	0.05223041	-0.02554548	0.09013353

MANOVA TEST CRITERIA FOR THE HYPOTHESIS OF NO OVERALL LOCATION EFFECT

H = TYPE III SSCP MATRIX FOR: LOCATION  
 E = ERROR SSCP MATRIX  
 P = DEP. VARIABLES = 4  
 Q = HYPOTHESIS DF = 2  
 NE = DF OF E = 57  
 S = MIN(P,Q) = 2  
 M = .5(AJS(P-Q)-1) = 0.5  
 Y = .5(NE-P-1) = 26.0

---

HOTELLING-LAWLEY TRACE =  $TR(E^{-1}H)$  = 11.92747793 (SEE PILLAI'S TABLE #3)  
 F APPROXIMATION =  $2(S*N+1)*TR(E^{-1}H)/(S*S*(2M+S+1))$  WITH  $S(2M+S+1)$  AND  $2(S*M+1)$  DF  
 $F(3,106) = 79.02$  PROB > F = 0.0001

---

PILLAI'S TRACE  $V = TR(H*INV(H+E))$  = 1.12776525 (SEE PILLAI'S TABLE #2)  
 F APPROXIMATION =  $(2N+S+1)/(2M+S+1) * V/(S-V)$  WITH  $S(2M+S+1)$  AND  $S(2N+S+1)$  DF  
 $F(8,110) = 17.78$  PROB > F = 0.0001

---

WILKS' CRITERION  $L = DET(E)/DET(H+E)$  = 0.06252690 (SEE RAO 1973 P 555)  
 EXACT F =  $(1-SQRT(L))/SQRT(L)*(NE+Q-P-1)/P$  WITH  $2P$  AND  $2(NE+Q-P-1)$  DF  
 $F(8,109) = 40.45$  PROB > F = 0.0001

---

ROY'S MAXIMUM ROOT CRITERION = 11.56693295 (SEE AMS VGL 31 P 625)  
 FIRST CANONICAL VARIABLE YIELDS AN F UPPER BOUND  
 $F(2,57) = 332.51$  (UPPER BOUND)

---

SAS ANALYSIS ON SEDIMENT DATA  
CANONICAL DISCRIMINANT ANALYSIS

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60 OBSERVATIONS 59 DF TOTAL  
4 VARIABLES 57 DF WITHIN CLASSES  
3 CLASSES 2 DF BETWEEN CLASSES

CANONICAL CORRELATIONS AND TESTS OF H0: THE CANONICAL CORRELATION IN THE CURRENT ROW AND ALL THAT FOLLOW ARE ZERO

	CANONICAL CORRELATION	ADJUSTED CAN CORR	APPROX STD ERROR	VARIANCE RATIO	CANONICAL R-SQUARED	LIKELIHOOD RATIO	F STATISTIC	NUM DF	DEN DF	PROB>F
1	0.959715637	0.754743941	0.711277393	11.6567	0.921054104	0.062626899	40.4453	8	108	0.0000
2	0.454354773	0.369637739	0.103277412	0.2606	0.206711149	0.793238951	4.7772	3	55	0.0050

MULTIVARIATE TEST STATISTICS AND F APPROXIMATIONS

STATISTIC	VALUE	F	NUM DF	DEN DF	PROB>F
WILKS' LAMBDA	0.0625259	40.44525	8	108	4.08791E-29
PILLAI'S TRACE	1.127765	17.77321	3	110	9.50357E-17
HOTELLING-LAHEY TRACE	11.92743	79.01954	8	106	3.75164E-41
ROY'S GREATEST ROOT	11.6669	160.4199	4	55	1.25025E-29

NOTE: F STATISTIC FOR ROY'S GREATEST ROOT IS AN UPPER BOUND  
F STATISTIC FOR WILKS' LAMBDA IS EXACT

TOTAL CANONICAL STRUCTURE

	CAN1	CAN2
PEPSAND	0.9933	0.0841
PEPSLTCL	-0.9385	-0.1103
PERGRAV	0.1522	0.7324
PERORG	-0.8575	0.2490

BETWEEN CANONICAL STRUCTURE

	CAN1	CAN2
PEPSAND	0.9992	0.0401
PEPSLTCL	-0.9986	-0.0528
PERGRAV	0.4234	0.9059
PERORG	-0.9907	0.1363

WITHIN CANONICAL STRUCTURE

	CAN1	CAN2
PEPSAND	0.9313	0.2499
PEPSLTCL	-0.8399	-0.3143
PERGRAV	0.0490	0.7015
PERORG	-0.4323	0.3984

SAS ANALYSIS ON SEDIMENT DATA  
CANONICAL DISCRIMINANT ANALYSIS

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STANDARDIZED CANONICAL COEFFICIENTS

	CAN1	CAN2
PERSAND	-80.5952	1849.49
PERSLTCL	-85.1493	1909.34
PERGRAV	-3.5903	71.8523
PERJRG	0.0425	1.2020

RAW CANONICAL COEFFICIENTS

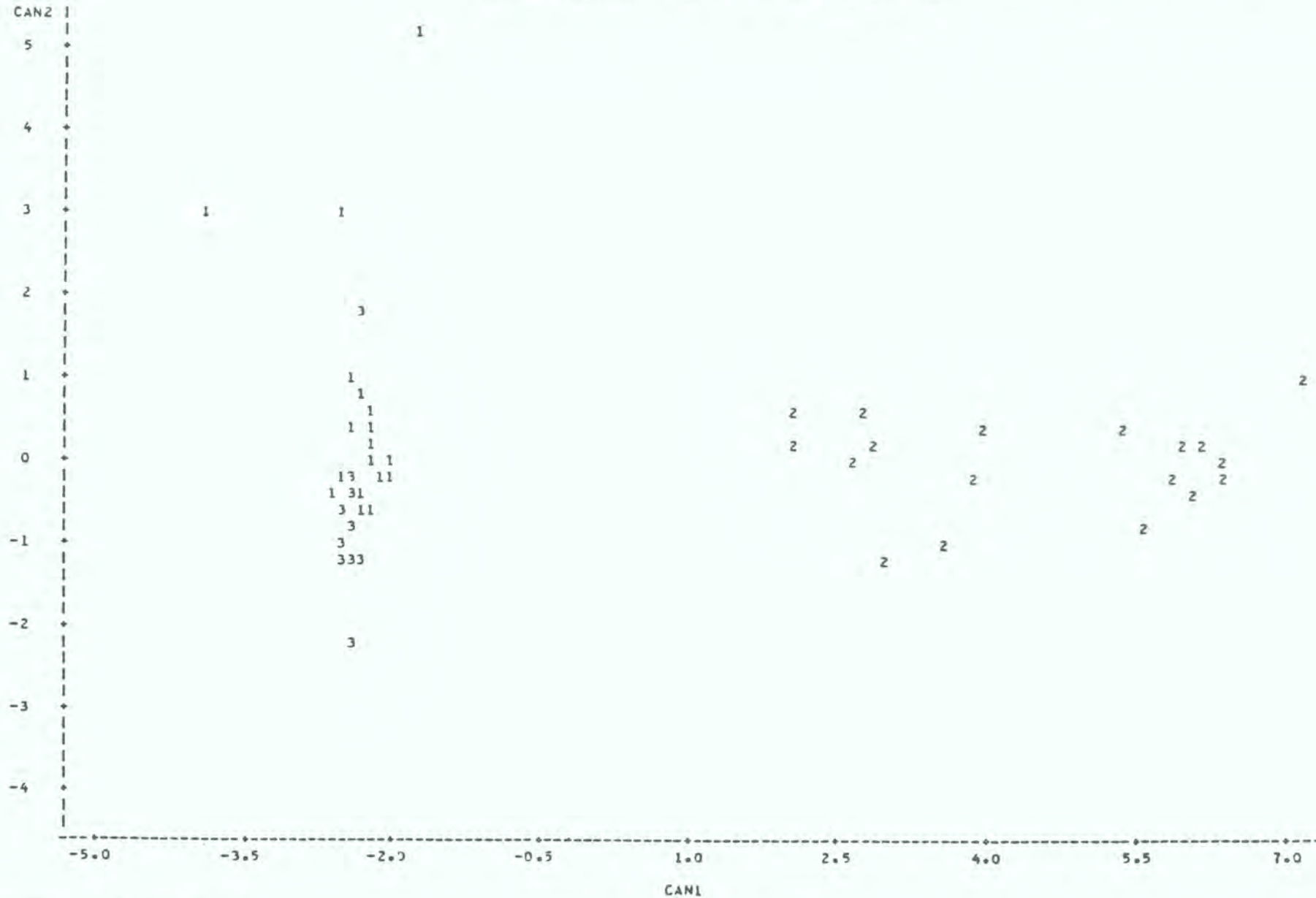
	CAN1	CAN2
PERSAND	-2.592265143	60.539147745
PERSLTCL	-2.593435433	60.536937759
PERGRAV	-3.059038984	61.191655247
PERJRG	0.013772464	0.531777736

CLASS MEANS ON CANONICAL VARIABLES

LOCATION	CAN1	CAN2
1	-2.3143	0.6129
2	4.7080	-0.0069
3	-2.3937	-0.6059

SAS ANALYSIS ON SEDIMENT DATA  
 PLOT OF CAN2\*CAN1 SYMBOL IS VALUE OF LOCATION

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NOTE: 14 OBS HIDDEN

## 9. PRINCIPLES OF SAMPLING DESIGN

### 9.1 Ten Principles

1. Be able to state concisely to someone else what question you are asking. Your results will be as coherent and as comprehensible as your initial conception of the problem.
2. Take replicate samples within each combination of time, location, and any other controlled variable. Differences among can only be demonstrated by comparison to differences within.
3. Take an equal number of randomly allocated replicate samples for each combination of controlled variables. Putting samples in "representative" or "typical" places is not random sampling.
4. To test whether a condition has an effect, collect samples both where the condition is present and where the condition is absent but all else is the same. An effect can only be demonstrated by comparison with a control.
5. Carry out some preliminary sampling to provide a basis for evaluation of sampling design and statistical analysis options. Those who skip this step because they do not have enough time usually end up losing time.
6. Verify that your sampling device or method is sampling the population you think you are sampling, and with equal and adequate efficiency over the entire range of sampling conditions to be encountered. Variation in efficiency of sampling from area to area biases among-area comparisons.
7. If the area to be sampled has a large-scale environmental pattern, break the area up into relatively homogeneous subareas and allocate samples to each in

proportion to the size of the subarea. If it is an estimate of total abundance over the entire area that is desired, make the allocation proportional to the number of organisms in the subarea.

8. Verify that your sample unit size is appropriate to the size, densities, and spatial distributions of the organisms you are sampling. Then estimate the number of replicate samples required to obtain the precision you want.

9. Test your data to determine whether the error variation is homogeneous, normally distributed, and independent of the mean. If it is not, as will be the case for most field data, then (a) appropriately transform the data, (b) a distribution-free (nonparametric) procedure, (c) use an appropriate sequential sampling design, or (d) test against simulated  $H_0$  data.

10. Having chosen the best statistical method to test your hypothesis, stick with the result. An unexpected or undesired result is not a valid reason for rejecting the method and hunting for a "better" one.

## 9.2 Estimation of sample number

### 9.2.1 Based on preliminary sampling:

Say that preliminary sampling estimates  $\bar{X}_1 = 18$  and  $S_1^2 = 236$ . If you wish to collect enough samples to estimate  $\bar{X}_2$  so that the true mean  $\mu$  lies within  $\pm 20\%$  of  $\bar{X}$  with a chance of  $\alpha = 0.05$  or less that it doesn't, then

$$\bar{X} \pm t \text{ (S.E.)} = \bar{X} \pm t\sqrt{S^2/n} = \bar{X} \pm tS/\sqrt{n}$$

which should equal  $\bar{X} \pm 0.2\bar{X}$ .

Therefore  $t \frac{S}{\sqrt{n}} = 0.2 \bar{X}$  and, if  $n$  is fairly large,

$$(2) 15.36/\sqrt{n} \approx (0.2)(18) \text{ and } n \approx 73.$$

(The value of  $t_{\alpha=0.5}$  for 72 df is almost exactly 2.)

9.2.2 Without preliminary sampling, but assuming Taylor's Power Law:

$$S^2 = a\bar{X}^{-b}, \text{ with } a \approx 1 \text{ and } b \approx 2:$$

$$\text{Define } D_o = \text{S.E.}/\bar{X} = \frac{S/\sqrt{n}}{\bar{X}}.$$

If  $a \approx 1$  and  $b \approx 2$ , then  $S^2 = a\bar{X}^{-b} \approx \bar{X}^2$  and  $S \approx \bar{X}$ .

$$\text{Therefore } D_o = \frac{S/\sqrt{n}}{\bar{X}} = \frac{X/\sqrt{n}}{\bar{X}} = \frac{1}{\sqrt{n}} \text{ and } n \approx \frac{1}{D_o^2}$$

If we want a precision of  $\pm 20\%$  with  $\alpha=0.05$ ,

$$(2) (\text{S.E.}) \approx 0.2 \bar{X} \text{ as before, and}$$

$$(2) \text{S.E.}/\bar{X} = 2 D_o \approx 0.2 .$$

$$\text{Therefore } D_o \approx 0.1 \text{ and } n \approx \frac{1}{D_o^2} = \frac{1}{(0.1)^2} = 100$$



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## APPENDIX I. - AGENDA

## OPENING SESSION

Addresses by: K. L. Chan, Acting Head, Zoology Department  
National University of Singapore

: J. R. E. Harger, Unesco-Mab/Unep representative  
from ROSTISEA

: H. H. Huang, Deputy Vice-Chancellor National  
University of Singapore

## COURSE SCHEDULE

	Morning	Afternoon
April 22 M	Introductory remarks - hand out schedule.	Tour of facilities, illustrate use of equipment, hand out out Tu-W tutorial.
23 Tu	Course organization, objectives, assumptions re. background, review of simple linear regression analysis.	Doing linear regression & plots using the various hardware & software.
24 W	Principles of linear, regression analysis, Models I & II, "Cookbook" versus matrix algebra solutions - latter in MINITAB & APL.	Continuation - if there is time, try it with larger data set, and/or transformation of variables. Explore MINITAB, APL.
25 Th	Introduction to common bivariate relationships in biology - nonlinear models.	Demonstration of doing regression by matrix algebra in MINITAB and APL. Demonstration of MINITAB plot & regression options. Introduction to SAS.

26	F	Continuation of presentation of biologically important nonlinear bivariate models. Intro. to ratio variables.	Doing linear and nonlinear regression analysis - emphasis on MINITAB, some SAS. Includes matrix algebra approach.	
27	Sa	Unscheduled - please use to your best advantage.	_____	
<hr/>				
29	M	Ratio variables, Taylor's Power Law (re. choosing transformations), introduction to analysis of covariance.	Doing nonlinear regression models with asymptotes, emphasis on MINITAB.	
30	Tu	Analysis of covariance, Walford plots, transforming multivariate data.	Doing a nonlinear analysis using micros. Doing a ratio variable problem.	
May	1	W	Review of material covered so far (please have questions ready!). Overview of statistical models, including multivariate.	Doing a Walford plot problem in MINITAB.
2	Th	Continue introduction to multivariate models - emphasis on tests for structure, ordination & clustering.	Doing an analysis of covariance in MINITAB, SAS, and on micros.	
3	F	Ordination & clustering, and introduction to MV ANOVA and discriminant analysis.	Calculation of matrices for MV analysis, testing for structure - emphasis on MINITAB.	

4 Sa    Unscheduled – please use to  
          your best advantage.

---

6 M	MV ANOVA and discriminant analysis.	Doing MV structure tests, ordination, in MINITAB & SAS.
7 Tu	Canonical correlation analysis Introduction to principles of sampling design.	Doing ordination & clustering-related analyses using custom-written programs, including on micros.
8 W	Continuation on principles of sampling design – sample unit size, estimation of necessary number of samples.	Doing MANOVA and DFA in SAS and APL, and canonical correlation in SAS.
9 Th	Transforming data – rationale, principles, strategy. Begin discussion of examples of design problems.	Doing MANOVA/DFA on APPLES. Doing exercise in sampling random and contagious distributions.
10 F	Review of course. Course evaluation. Discussion of participants' design & biostatistical analysis problems "back home".	Continuation of morning discussion of "back home" case studies. Clean up. Make sure you have disks.

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CLOSING SESSION

Addresses by : Kuswata Kartawinata, Unesco/Mab/Unep  
representative from ROSTSEA.

: T. W. Chen, Acting Head, Zoology Department  
National University of Singapore.

: S. D. Tandjung, course participant  
representative.

## APPENDIX II -- PROGRAMS

FILE: ANCOVA BASIC A1 11/52 - CONVERSATIONAL MONITOR SYSTEM

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100 REM
110 REM ANCOVA PROGRAM WRITTEN BY KEITH SOMEPS, JAN. 1983.
120 REM
130 REM THIS PROGRAM READS DATA FROM A SERIES OF FILES AND THEN
140 REM COMPARES THE SLOPES AND INTERCEPTS IN AN ANALYSIS OF
150 REM COVARIANCE AS OUTLINED IN ZAR (1974) - PP. 234-235.
160 REM
170 PRINT "HOW MANY REGRESSIONS ARE BEING COMPARED?"
180 INPUT N
190 PRINT
200 FOR I=1 TO N
210 PRINT "HOW MANY DATA PAIRS ARE IN REGRESSION #;I"
220 INPUT N1(I)
230 NEXT I
240 PRINT "IF YOU WANT NO TRANSFORMATION OF X INPUT 0"
250 PRINT "IF YOU WANT A LOG(X) TRANSFORMATION INPUT 1"
260 PRINT "IF YOU WANT A LOG(X+1) TRANSFORMATION INPUT 2"
270 INPUT TX
280 PRINT "IF YOU WANT NO TRANSFORMATION OF Y INPUT 0"
290 PRINT "IF YOU WANT A LOG(Y) TRANSFORMATION INPUT 1"
300 PRINT "IF YOU WANT A LOG(Y+1) TRANSFORMATION INPUT 2"
310 INPUT TY
320 DIM X(500),Y(500)
330 FOR I=1 TO N
340 LET C5=N1(I)
350 PRINT "INPUT";N1(I);" X,Y PAIRS FOR REGRESSION";I
360 FOR C=1 TO C5
370 INPUT X(C),Y(C)
380 IF TX=1 THEN X(C)=LOG(X(C))
390 IF TX=2 THEN X(C)=LOG(X(C)+1)
400 IF TY=1 THEN Y(C)=LOG(Y(C))
410 IF TY=2 THEN Y(C)=LOG(Y(C)+1)
420 LET A=A+X(C)
430 LET B=B+X(C)2
440 LET D=D+Y(C)
450 LET E=E+Y(C)2
460 LET F1=F1+X(C)*Y(C)
470 NEXT C
480 LET G(I)=B-A2/C5
490 LET H(I)=F1-(A*D)/C5
500 LET S8(I)=H(I)/G(I)
510 LET J1(I)=A/C5
520 LET K1(I)=D/C5
530 LET I8(I)=K1(I)-S8(I)*J1(I)
540 LET M(I)=H(I)-2/G(I)
550 LET Z1(I)=E-D2/C5
560 LET O(I)=Z1(I)-M(I)
570 LET P(I)=C5-2
580 LET Q(I)=O(I)/P(I)
590 LET S9(I)=SQR(Q(I)/G(I))
600 LET I9(I)=SQR(Q(I)*(1/C5+J1(I)2/G(I)))
610 LET A1=A1+A
620 LET B1=B1+B
630 LET C1=C1+C5
640 LET D1=D1+D

```

A



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650 LET E2=E2+E
660 LET F2=F2+F1
670 LET A3=A3+G(I)
680 LET J3=J3+H(I)
690 LET C3=C3+E1(I)
700 LET S4=S4+D(I)
710 LET F4=F4+P(I)
720 LET A=J
730 LET N=0
740 LET D=0
750 LET E=0
760 LET F1=J
770 NEXT I
780 LET S3=C3-J3~2/A3
790 LET F3=C1-N-1
800 LET A5=B1-A1~2/C1
810 LET S6=F2-(A1*J1)/C1
820 LET C6=E2-D1~2/C1
830 LET S1=C6-B6~2/A6
840 LET F2=C1-2
850 LET R1=B6/A6
860 LET R2=(D1/C1)-(R1*(A1/C1))
870 LET M7=S4/F4
880 LET M3=S3/F3
890 LET F8=((S3-S4)/(N-1))/M7
900 LET F9=(S1-S3)/(N-1)/M3
910 LET F5=N-1
920 PRINT
930 IF N<>2 THEN 1170
940 PRINT "THE ANALYSIS OF COVARIANCE BETWEEN"
950 PRINT "REGRESSION ONE AND REGRESSION TWO"
960 PRINT "HAS PRODUCED THE FOLLOWING: "
970 LET J2=S4/F4
980 LET J3=SQR(J2/G(1)+J2/G(2))
990 LET J4=A3S(S3(1)-S8(2))/J3
1000 LET J8=B3/A3
1010 LET J5=ABS(K1(1)-K1(2))-J8*(ABS(J1(1)-J1(2)))
1020 LET J6=SQR((S3/F3)*(1/N1(1)+1/N1(2)+(J1(1)-J1(2))~2/A3))
1030 LET J7=J5/J6
1040 PRINT
1050 PRINT "THE STUDENT'S T STATISTIC FOR B1=B2 WITH A "
1060 PRINT "TWO-TAILED ALPHA OF 0.05 AND ";F4;" DEGREES OF FREEDOM"
1070 PRINT "IS: ";J4
1080 PRINT
1090 PRINT "IF B1=B2 IS FALSE, THEN TWO DIFFERENT POPULATIONS WERE SAMPLED."
1100 PRINT "IF B1=B2, THEN TEST FOR COMMON INTERCEPTS."
1110 PRINT
1120 PRINT "THE STUDENT'S T STATISTIC FOR COMMON INTERCEPTS WITH A "
1130 PRINT "TWO-TAILED ALPHA OF 0.05 AND ";F3;" DEGREES OF FREEDOM"
1140 PRINT "IS: ";J7
1150 PRINT
1160 GO TO 1290
1170 PRINT "THE ANALYSIS OF COVARIANCE HAS BEEN COMPLETED"
1180 PRINT "FOR THE ";N;" REGRESSION LINES."
1190 PRINT

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```
1200 PRINT "THE ANCOVA TEST AS OUTLINED IN ZAR (1974) CHAPTER 17,"
1210 PRINT "PRODUCES TWO F-STATISTICS THAT DETERMINE IF THE COMBINED"
1220 PRINT "SLOPES AND COMBINED INTERCEPTS ARE DIFFERENT."
1230 PRINT
1240 PRINT "THE F-VALUE FOR THE SLOPES IS: ";F8
1250 PRINT "THE F-VALUE FOR THE INTERCEPTS IS: ";F9
1260 PRINT "BOTH VALUES HAVE A NUMERATOR D. OF F. OF: ";F5
1270 PRINT "AND A DENOMINATOR D. OF F. OF: ";F3
1280 PRINT
1290 PRINT "IF THE SLOPES ARE NOT SIGNIFICANTLY DIFFERENT, "
1300 PRINT "THE COMMON REGRESSION SLOPE IS: ";R1
1310 PRINT
1320 PRINT "AND IF THE INTERCEPTS ARE NOT SIGNIFICANTLY DIFFERENT, "
1330 PRINT "THE COMMON REGRESSION INTERCEPT IS: ";R2
1340 IF N=2 THEN 2030
1350 PRINT
1360 PRINT "IF THE SLOPES OR INTERCEPTS ARE DIFFERENT, THEN "
1370 PRINT "A MULTIPLE RANGE TEST CAN BE USED TO IDENTIFY THE "
1380 PRINT "DIFFERENCES BETWEEN THE SET OF REGRESSIONS. "
1390 PRINT
1400 PRINT "DO YOU WANT TO DO THE MULTIPLE RANGE TESTS?"
1410 PRINT "TYPE Y OR N. "
1420 INPUT A9$
1430 IF A9$="N" THEN 2030
1440 REM TO COMPLETE THE MULTIPLE RANGE TESTS, SEVERAL PROCEDURES
1450 REM ARE AVAILABLE. TYPICALLY THE NEWMAN-KEULS MULTIPLE RANGE
1460 REM TEST IS USED IF EACH REGRESSION IS COMPARED WITH EACH
1470 REM OTHER REGRESSION (OPTION 1). HOWEVER, IF THE REGRESSIONS
1480 REM ARE BASED ON DIFFERENT X-VALUES THEN A DIFFERENT FORMULA
1490 REM MUST BE USED (OPTION 2). ALTERNATIVELY, IF ONE OF THE
1500 REM REGRESSION LINES IS A CONTROL AND ALL OTHERS ARE TO BE
1510 REM COMPARED TO IT, DUNNETT'S TEST IS APPROPRIATE (OPTION 3).
1520 REM AGAIN, IF THE X-VALUES ARE DIFFERENT, AN ALTERNATIVE
1530 REM FORMULA IS REQUIRED (OPTION 4).
1540 PRINT "CHOOSE A MULTIPLE RANGE TEST FROM THIS LIST: "
1550 PRINT "(1) NEWMAN-KEULS WITH THE SAME X-VALUES"
1560 PRINT "(2) NEWMAN-KEULS WITH DIFFERENT X-VALUES"
1570 PRINT "(3) DUNNETT'S TEST WITH THE SAME X-VALUES"
1580 PRINT "(4) DUNNETT'S TEST WITH DIFFERENT X-VALUES"
1590 INPUT A3
1600 PRINT
1610 PRINT
1620 PRINT "THE RESULTS OF THE MULTIPLE RANGE TESTS"
1630 PRINT "ARE AS FOLLOWS: "
1640 PRINT
1650 REM TO ACCURATELY DEFINE THE Q-STATISTIC GIVEN BELOW,
1660 REM YOU MUST RANK THE SLOPES AND INTERCEPTS FOR EACH
1670 REM REGRESSION. THE DIFFERENCE IN ORDER BETWEEN PAIRS
1680 REM OF REGRESSIONS PROVIDES A P-VALUE FOR D. OF F.
1690 REM NEEDED IN THE Q-TABLE (I.E. Q(0.05)(DF;V)(DF;P)).
1700 REM IF HI-TO-LOW RANKING SEPARATES REGS. 1 AND 4 BY 3 VALUES,
1710 REM THEN P=5 (I.E. FOR 135:46, REGS. 1 + 4 HAVE P=5).
1720 PRINT " REGRESSION SLOPE ELEVATION-(INT)"
1730 PRINT
1740 FOR I=1 TO N
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1750 PRINT " ";I;" " ";S1(I);" " ";I8(I)
1760 NEXT I
1770 PRINT
1780 PRINT " REGRESSIONS SLOPE=J ELEVATION=J C.F.(V)"
1790 PRINT
1800 FOR I=1 TO (N-1)
1810 LET K=I+1
1820 FOR J=K TO N
1830 LET J8=(H(I)+H(J))/(G(I)+G(J))
1840 IF A3>1 THEN 1850
1850 LET X8=SQRT(47/G(I))
1860 IF A3<>2 THEN 1830
1870 LET X8=SQRT((47/2)*(1/G(I)+1/G(J)))
1880 IF A3<>3 THEN 1830
1890 LET X8=SQRT(2*47/G(I))
1900 IF A3<>4 THEN 1830
1910 LET X8=SQRT(47*(1/G(I)+1/G(J)))
1920 LET J9=ABS(S8(I)-S8(J))/X8
1930 IF A3=1 THEN 1950
1940 IF A3<>2 THEN 1960
1950 LET Y8=SQRT(M3/2*(1/N1(I)+1/N1(J)+(J1(I)-J1(J))-2/(G(I)+G(J))))
1960 IF A3<3 THEN 1970
1970 IF A3>4 THEN 1990
1980 LET Y8=SQRT(M3*(1/N1(I)+1/N1(J)+(J1(I)-J1(J))-2/(G(I)+G(J))))
1990 LET R8=ABS((K1(I)-K1(J))-B8*(J1(I)-J1(J)))/Y8
2000 PRINT " ";I;" AND ";J;" " ";J8;" " ";R8;" " ";F4
2010 NEXT J
2020 NEXT I
2030 END

```

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100 REM
110 REM LINEAR REGRESSION PROGRAM WRITTEN BY KEITH SOMERS, MAY 1982.
120 REM
130 REM SIMPLE LINEAR REGRESSION PROGRAM THAT RECEIVES DATA FROM THE
140 REM KEYBOARD AND COMPUTES ADVANCED STATISTICS FOR THAT DATA.
150 REM
160 REM THE STATISTICS AND REGRESSION ANALYSIS FOLLOW THE CHAPTER
170 REM ON THAT SUBJECT IN "BIOSTATISTICAL ANALYSIS" BY ZAR
180 REM
190 DIM X(200),Y(200)
200 DIM Y1(200),Y2(200)
210 PRINT "HOW MANY X-Y PAIRS DO YOU WANT TO ENTER?"
220 INPUT N
230 PRINT
240 PRINT "IF YOU WANT NO TRANSFORMATION OF X INPUT 0"
250 PRINT "IF YOU WANT A LOG(X) TRANSFORMATION INPUT 1"
260 PRINT "IF YOU WANT A LOG(X+1) TRANSFORMATION INPUT 2"
270 INPUT TX
280 PRINT "IF YOU WANT NO TRANSFORMATION OF Y INPUT 0"
290 PRINT "IF YOU WANT A LOG(Y) TRANSFORMATION INPUT 1"
300 PRINT "IF YOU WANT A LOG(Y+1) TRANSFORMATION INPUT 2"
310 INPUT TY
320 PRINT
330 PRINT "ENTER THE DATA AS X-Y PAIRS."
340 PRINT
350 FOR C=1 TO N
360 INPUT X(C),Y(C)
370 IF TX=1 THEN X(C)=LOG(X(C))
380 IF TX=2 THEN X(C)=LOG(X(C)+1)
390 IF TY=1 THEN Y(C)=LOG(Y(C))
400 IF TY=2 THEN Y(C)=LOG(Y(C)+1)
410 NEXT C
420 LET Z5=N-2
430 PRINT
440 PRINT "WHAT IS THE T-VALUE FOR THE 95% CONFIDENCE LIMITS"
450 PRINT "WITH ";Z5;" DEGREES OF FREEDOM?"
460 INPUT T1
470 PRINT
480 PRINT
490 FOR C=1 TO N
500 REM A IS THE SUM OF THE X VALUES
510 LET A=A+X(C)
520 REM B IS THE SUM OF SQUARED X VALUES
530 REM ZAR'S BIG X-SQUARED
540 LET B=B+X(C)^2
550 REM D IS THE SUM OF THE Y VALUES
560 LET D=D+Y(C)
570 REM E IS THE SUM OF SQUARED Y VALUES
580 REM ZAR'S BIG Y-SQUARED
590 LET E=E+Y(C)^2
600 REM F IS THE SUM OF X*Y
610 REM ZAR'S BIG XY
620 LET F=F+X(C)*Y(C)
630 NEXT C
640 REM G IS THE SUM OF SQUARES OF X

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```
550 REM ZAR'S LITTLE X-SQUARED
560 LET G=J-A*2/N
570 REM H IS THE SUM OF CROSS-PRODUCT DEVIATIONS
580 REM ZAR'S LITTLE XY
590 LET H=F-(A*J)/N
700 REM I IS THE SLOPE
710 LET I=H/G
720 REM J IS THE MEAN OF X
730 LET J=A/N
740 REM K IS THE MEAN OF Y
750 LET K=J/N
760 REM L IS THE Y-INTERCEPT
770 LET L=K-I*J
780 REM M IS THE SUM OF SQUARES OF THE REGRESSION
790 REM N IS ALSO THE REGRESSION MEAN SQUARE
800 LET N=I*2/G
810 REM S1 IS THE TOTAL SUM OF SQUARES
820 REM ZAR'S LITTLE Y-SQUARED
830 LET E1=E-C*2/N
840 REM M1 IS R-SQUARED, THE COEFFICIENT OF DETERMINATION
850 LET M1=M/E1
860 REM N2 IS R, THE CORRELATION COEFFICIENT
870 LET N2=SQR(M1)
880 REM Q IS THE RESIDUAL SUM OF SQUARES
890 LET Q=E1-M
900 REM P IS THE RESIDUAL DEGREES OF FREEDOM
910 LET P=N-2
920 REM V IS THE RESIDUAL MEAN SQUARE
930 LET V=Q/P
940 REM R IS THE F-STATISTIC TO DETERMINE IF THE SLOPE EQUALS ZERO
950 LET R=M/Q
960 REM S IS THE STANDARD ERROR OF THE ESTIMATE (EPSILON)
970 LET S=SQR(Q)
980 REM I1 IS THE STANDARD ERROR OF THE SLOPE WHICH IS USED TO TEST
990 REM FOR SIGNIFICANCE OF THE SLOPE AS RELATED TO A SPECIFIED VALUE
1000 LET I1=SQR(Q/G)
1010 REM I2 AND I3 ARE 95% CONFIDENCE LIMITS AROUND THE SLOPE
1020 LET I2=I-T1*I1
1030 LET I3=I+T1*I1
1040 REM L1 IS THE STANDARD ERROR OF THE INTERCEPT
1050 LET L1=SQR(Q*(1/N+J*2/G))
1060 REM L2 AND L3 ARE THE 95% C.L. FOR THE INTERCEPT
1070 LET L2=L-T1*L1
1080 LET L3=L+T1*L1
1090 PRINT "THE REGRESSION STATISTICS ARE AS FOLLOWS:"
1100 PRINT
1110 LET I5=ABS(I)
1120 IF I5-I=0 THEN 1150
1130 PRINT "THE EQUATION OF THE LINE IS: Y=";L;"-";I5;"X"
1140 GO TO 1150
1150 PRINT "THE EQUATION OF THE LINE IS: Y=";L;"+";I;"X"
1160 PRINT
1170 PRINT "WHERE THE SLOPE IS: ";I
1180 PRINT "AND THE Y-INTERCEPT IS: ";L
1190 PRINT "THE STANDARD ERROR OF THE REGRESSION IS: (+ OR -) ";S
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```
1200 PRINT "THE STANDARD ERROR OF THE SLOPE IS: (+ OR -) ";I1
1210 PRINT "THE 95% C.L. FOR THE SLOPE ARE: ";I2;" ";I3
1220 PRINT "THE STANDARD ERROR OF THE INTERCEPT IS: (+ OR -) ";L1
1230 PRINT "THE 95% C.L. FOR THE INTERCEPT ARE: ";L2;" ";L3
1240 PRINT "THE CORRELATION COEFFICIENT (R) IS: ";N2
1250 PRINT "THE COEFFICIENT OF DETERMINATION (R**2) IS: ";N1
1260 PRINT
1270 PRINT
1280 PRINT "DO YOU WANT MORE STATISTICS PRINTED?"
1290 PRINT "TYPE Y OR N."
1300 INPUT S5$
1310 IF S5$="N" THEN 1540
1320 PRINT
1330 PRINT
1340 PRINT "THE REGRESSION COMPUTATIONS HAVE PRODUCED THE FOLLOWING: "
1350 PRINT
1360 PRINT "THE MEANS OF X AND Y ARE: ";J;" ";K
1370 PRINT "THE SUM OF X IS: ";A
1380 PRINT "THE SUM OF Y IS: ";D
1390 PRINT "THE SUM OF X-SQUARED IS: ";B
1400 PRINT "THE SUM OF Y-SQUARED IS: ";E
1410 PRINT "THE SUM OF X*Y IS: ";F
1420 PRINT "THE SUM OF SQUARES OF X IS: ";G
1430 PRINT "THE SUM OF CROSS-PRODUCTS IS: ";H
1440 PRINT "THE REGRESSION SUM OF SQUARES IS: ";M
1450 PRINT "THE RESIDUAL SUM OF SQUARES IS: ";J
1460 PRINT "THE TOTAL SUM OF SQUARES IS: ";E1
1470 PRINT "THE REGRESSION MEAN SQUARE IS: ";M
1480 PRINT "THE RESIDUAL MEAN SQUARE IS: ";D
1490 PRINT
1500 PRINT "THE F-VALUE IS: ";R
1510 PRINT "WITH 1 REGRESSION D OF F, AND"
1520 PRINT "A RESIDUAL D OF F OF : ";P
1530 PRINT
1540 PRINT
1550 PRINT "DO YOU WANT 95% CONFIDENCE LIMITS?"
1560 PRINT "TYPE Y OR N."
1570 INPUT S6$
1580 IF S6$="N" THEN 1830
1590 PRINT
1600 PRINT "DO YOU WANT TO SPECIFY THE X VALUES?"
1610 PRINT "TYPE Y OR N."
1620 INPUT S7$
1630 IF S7$="Y" THEN 1660
1640 LET N4=N
1650 GO TO 1740
1660 PRINT
1670 PRINT "HOW MANY X VALUES DO YOU WANT TO ENTER?"
1680 INPUT N4
1690 PRINT "LIST THE EACH X VALUE BELOW."
1700 PRINT
1710 FOR C=1 TO N4
1720 INPUT (C)
1730 NEXT C
1740 PRINT
```

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1750 PRINT "THE PREDICTED VALUES AND 95% C.L. OF Y ARE:"
1760 PRINT
1770 PRINT "GIVEN X VALUE   PREDICTED Y   LOWER Y   UPPER Y   ERROR"
1780 FOR C=1 TO 14
1790 REM T IS THE PREDICTED Y VALUE
1800 LET T=L+X(C)
1810 REM S IS THE STANDARD ERROR OF THE PREDICTED Y FOR THAT X VALUE
1820 LET S=SQ(C*(1/4+((X(C)-J)^2)/G))
1830 REM Y1 AND Y2 ARE THE 95% C.L. AROUND THE PREDICTED Y
1840 LET Y1(C)=T-(T*S)
1850 LET Y2(C)=T+(T*S)
1860 PRINT "  ";X(C);"  ";T;"  ";Y1(C);"  ";Y2(C);"  "
1870 NEXT C
1880 END

```

REAL X,Y,P,Q,A,I,MX,MY,VAPX,VARY,COVXY,SLX2,SLY2,SLXY,RSS,ESS,TSS	REG00010
REAL RMS,EMS,F,R2,P=RR2	REG00020
INTEGER N,RDF,EDF,TDF,TX,TY	REG00030
DIFFUSION X(50),Y(50),YHAT(50),YRES(50)	REG00040
SUMX=0	REG00050
SUMY=0	REG00060
SUMXY=0	REG00070
SUMX2=0	REG00080
SUMY2=0	REG00090
I=1	REG00100
WRITE(5,12)	REG00110
12 FORMAT('THE DATA AS READ IN, BEFORE ANY TRANSFORMATION, ARE:')	REG00120
WRITE(6,14)	REG00130
14 FORMAT(6X,'X',12X,'Y')	REG00140
READ(2,*)TX,TY	REG00150
5 READ(2,*)X(I),Y(I)	REG00160
IF(X(I).EQ.0)GOTO 20	REG00170
WRITE(6,16)X(I),Y(I)	REG00180
16 FORMAT(F10.3,3X,F10.3)	REG00190
IF(TX.EQ.1) X(I)=LOG(X(I))	REG00200
IF(TX.EQ.2) X(I)=LOG(X(I))+1	REG00210
IF(TY.EQ.1) Y(I)=LOG(Y(I))	REG00220
IF(TY.EQ.2) Y(I)=LOG(Y(I))+1	REG00230
SUMX=SUMX+X(I)	REG00240
SUMY=SUMY+Y(I)	REG00250
SUMXY=SUMXY+X(I)*Y(I)	REG00260
SUMX2=SUMX2+X(I)*X(I)	REG00270
SUMY2=SUMY2+Y(I)*Y(I)	REG00280
I=I+1	REG00290
GO TO 5	REG00300
20 N=I-1	REG00310
WRITE(5,15)	REG00320
WRITE(5,22)	REG00330
22 FORMAT('THE DATA AFTER TRANSFORMATION, IF ANY, ARE:')	REG00340
WRITE(6,14)	REG00350
24 FORMAT(6X,'X',12X,'Y')	REG00360
DO 28 I=1,N	REG00370
28 WRITE(6,16)X(I),Y(I)	REG00380
26 FORMAT(F10.3,3X,F10.3)	REG00390
WRITE(5,150)	REG00400
MX=SUMX/N	REG00410
MY=SUMY/N	REG00420
SLX2=SUMX2-SUMX*SUMX/N	REG00430
SLY2=SUMY2-SUMY*SUMY/N	REG00440
SLXY=SUMXY-SUMX*SUMY/N	REG00450
VARX=SLX2/(N-1)	REG00460
VARY=SLY2/(N-1)	REG00470
COVXY=SLXY/(N-1)	REG00480
B=SLXY/SLX2	REG00490
A=MY-B*MX	REG00500
RDF=1	REG00510
EDF=N-2	REG00520
TDF=N-1	REG00530
RSS=SLXY-B*SLXY/SLX2	REG00540
ESS=SLY2-RSS	REG00550



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TSS=GLYZ
RMS=RSS
EMS=ESS/EDF
F=RMS/EMS
R2=RSS/TSS
PERR2=100*R2
WRITE(5,30)'A,MY
30  FORMAT(' ',2X,'X MEAN=',F9.2,3X,'Y MEAN=',F9.2)
WRITE(6,47)VAPX,VARY,COVXY
40  FORMAT('X VARIANCE=',F8.2,3X,'Y VARIANCE=',F8.2,3X,
5  'XY COVARIANCE =',F9.2)
WRITE(5,150)
WRITE(6,150)
WRITE(5,57)A,B
50  FORMAT(' ',4,'THE REGRESSION LINE IS Y=',F9.4,' + ',F9.4,'X')
WRITE(5,150)
WRITE(6,150)
WRITE(6,60)
60  FORMAT(' ',4,'THE ANALYSIS OF VARIANCE TABLE IS:')
WRITE(5,150)
WRITE(6,70)
70  FORMAT(' ',4,'SOURCE',2X,'SUM OF SQUARES',2X,'MEAN SQUARE',2X,
6  'F-STATISTIC')
WRITE(6,30)PDF,RSS,RMS,F
80  FORMAT(' ',15,4X,F8.2,5X,F7.2,5X,F7.2)
WRITE(5,90)EDF,ESS,EMS
90  FORMAT(15,4X,F8.2,5X,F7.2)
WRITE(6,100)
100  FORMAT(1X,'-----',4X,'-----')
WRITE(5,110)TDF,TSS
110  FORMAT(15,4X,F8.2)
WRITE(6,150)
WRITE(5,120)R2,PERR2
120  FORMAT(' ',4,'R-SQUARED=',F6.5,3X,'PERCENT R-SQUARED=',F5.2)
WRITE(6,150)
WRITE(5,150)
WRITE(6,130)
130  FORMAT('Y-PREDICTEDS AND Y-RESIDUALS FOLLOW.')
WRITE(5,150)
WRITE(6,150)
DO 130 I=1,N
YHAT(I)=A+B*X(I)
130  YRES(I)=YHAT(I)-Y(I)
WRITE(5,150)
WRITE(6,140)
140  FORMAT('Y-PREDICTEDS',3X,'Y-RESIDUALS')
DO 170 I=1,N
170  WRITE(5,160)YHAT(I),YRES(I)
160  FORMAT(F10.3,3X,F11.3)
150  FORMAT(' ')
STOP
ENJ
REG0056J
REG0057J
REG0058J
REG0059J
REG0060J
REG0061J
REG0062J
REG0063J
REG0064J
REG0065J
REG0066J
REG0067J
REG0068J
REG0069J
REG0070J
REG0071J
REG0072J
REG0073J
REG0074J
REG0075J
REG0076J
REG0077J
REG0078J
REG0079J
REG0080J
REG0081J
REG0082J
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REG0088J
REG0089J
REG0090J
REG0091J
REG0092J
REG0093J
REG0094J
REG0095J
REG0096J
REG0097J
REG0098J
REG0099J
REG0100J
REG0101J
REG0102J
REG0103J
REG0104J
REG0105J
REG0106J
REG0107J

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C THIS PROGRAM CALCULATES A P VARIABLES-BY-P VARIABLES CORRELATION
C MATRIX, ORDERS THE VARIABLES BY A DECREASING SUM OF
C R SQUARED CRITERION, AND THEN IT CAN CONTINUE, SWEEPING
C THE MATRIX OF ALL CORRELATIONS WITH THE FIRST VARIABLE, AND
C REPEATING THE ORDERING AND SWEEPING PROCESS UNTIL ALL P
C VARIABLES (OR SOME SPECIFIED SUBSET OF VARIABLES) HAVE BEEN
C CHOSEN. THIS PROCEDURE AND THE ALGORITHM FOR DOING IT WAS
C ORIGINALLY PROPOSED BY L. ORLUCI (1973; NATURE, LONDON 244:
C 371-373). PROGRAMS WRITTEN IN BASIC ARE GIVEN IN ORLUCI'S
C 1978 BOOK AND THE ORLUCI & KRIKEL 1984 COURSE MANUAL (SEE THE
C "BIBLIOGRAPHY" YOU WERE GIVEN FOR THE FULL REFERENCES). THIS
C IMPLEMENTATION OF THE ALGORITHM IN FORTRAN IS BY R. M. GREEN.
C
C THERE IS A CONTROL CARD, WHICH SHOULD BE FOLLOWED BY THE DATA
C CARDS. THE N-BY-P DATA ARE ASSUMED TO BE IN "FREE FORMAT".
C
C THE CONTROL CARD SHOULD HAVE IN IT (IN FREE FORMAT) THE
C VARIABLES N, P, CYCLE, BRIEF, AND CPCT, WHERE:
C (A) N = NUMBER OF SAMPLES (NOW DIMENSIONED FOR 200)
C (B) P = NUMBER OF VARIABLES (NOW DIMENSIONED FOR 150)
C (C) CYCLE = NUMBER OF VARIABLES TO BE CHOSEN
C (D) BRIEF (> 0 IF PRINTOUT OF CORRELATION MATRICES IS
C DESIRED, = 0 OTHERWISE)
C (E) CPCT = PERCENTAGE OF CORRELATION STRUCTURE TO BE
C ACCOUNTED FOR
C
C
C DIMENSION X(600,150),RSJ(150,150),SUM(150),
C EPSMJJ(150),R(150,150),SMRZ(150),RANK(150),VARNO(150),FMTIN(20)
C F,FMTOUT(20),H(150),FRNK(150),DIAG(150),RI(150,150),TPCT(150)
C REAL X,RSJ,SUM,SUMSJ,SUMJK,RSJK,RS4JK,RS4JK2,P,CYCLE,LARGE,RANK,
C EDIAG,RI,SIGN,TPCT,CPCT,FLIP
C INTEGER N,P,I,FMTIN,FMTOUT,J,K,L,VARNO,NUM,H,CYCLE,M,FRNK,BRIEF,
C EG,REST,J
C READ(2,*)N,P,CYCLE,BRIEF,CPCT
C FLIP=100-CPCT
C WRITE(6,42)
C 42 FORMAT(' ')
C WRITE(6,62)N,P,CYCLE,BRIEF,CPCT
C 62 FORMAT(' N =',I4,6X,' P =',I4,6X,' CYCLE =',I3,6X,
C ' BRIEF =',I2,6X,' CPCT =',F4.0)
C WRITE(6,*)' '
C WRITE(6,20)
C 20 FJRMAT('THE N SAMPLES-BY-P VARIABLES INPUT DATA ARE:')
C DO 30 I=1,N
C READ(2,*)(X(I,J),J=1,P)
C WRITE(6,1030)(X(I,J),J=1,P)
C 1030 FORMAT(JF10.3)
C 30 CONTINUE
C WRITE(6,*)' '
C DO 40 J=1,P
C DO 50 I=1,N
C SUM(J)=SUM(J)+X(I,J)

```

RSL00010  
RSL00020  
RSL00030  
RSL00040  
RSL00050  
RSL00060  
RSL00070  
RSL00080  
RSL00090  
RSL00100  
RSL00110  
RSL00120  
RSL00130  
RSL00140  
RSL00150  
RSL00160  
RSL00170  
RSL00180  
RSL00190  
RSL00200  
RSL00210  
RSL00220  
RSL00230  
RSL00240  
RSL00250  
RSL00260  
RSL00270  
RSL00280  
RSL00290  
RSL00300  
RSL00310  
RSL00320  
RSL00330  
RSL00340  
RSL00350  
RSL00360  
RSL00370  
RSL00380  
RSL00390  
RSL00400  
RSL00410  
RSL00420  
RSL00430  
RSL00440  
RSL00450  
RSL00460  
RSL00470  
RSL00480  
RSL00490  
RSL00500  
RSL00510  
RSL00520  
RSL00530  
RSL00540  
RSL00550

```

SUMSQ=SUMSQ+(X(I,J))2/(X(I,J))
50 CONTINUE
PS4JJ(J)=SUMSQ-(SUM4(J))2/(C4(J))/N
SUMSQ=0.0
40 CONTINUE
DO 60 J=1,P
DO 50 K=J,P
DO 70 I=1,N
SUMJK=SUMJK+(X(I,J))2/(X(I,K))
70 CONTINUE
RS4JK=SUMJK-((SUM4(J)+SUM4(K))/N)
SIGN=RS4JK/ABS(RS4K)
RS4K2=RS4K2*RS4K
RSQ(J,K)=RS4K2/(RS4JJ(J)+RS4JJ(K))
P(J,K)=SIGN(RS4K)
P(K,J)=SIGN(RS4K)
P(K,J)=RSQ(J,K)
R(K,J)=R(J,K)
SUMJK=0.0
60 CONTINUE
170 IF (BRIEF.EQ.0) GO TO 22
WRITE(5,*) ' *
WRITE(5,3)
80 FORMAT('THE P-Q-P CORRELATION MATRIX IS:')
DO 100 J=1,P
WRITE(5,11)(R(J,K),K=1,P)
110 FORMAT(15F3.4)
100 CONTINUE
WRITE(5,*) ' *
DO 111 J=1,P
DO 111 K=1,P
111 S4R2(J)=0.0
22 DO 120 J=1,P
DO 120 K=1,P
S4R2(J)=S4R2(J)+RSQ(J,K)
H(J)=J
120 CONTINUE
M=1
J=1
130 Q=P-J+1
LARGE=S4R2(1)
NUM=H(1)
L=1
IF (J.EQ.1) GO TO 1320
DO 5 J=2,P
IF (S4R2(J).LE.LARGE) GO TO 4
LARGE=S4R2(J)
NUM=H(J)
L=J
6 CONTINUE
1020 P=Q+J-1
RANK(M)=LARGE
VARNO(M)=NUM
M=M+1
H(L)=H(P-J+1)

```

```

RSL0056J
RSL0057J
RSL0058J
RSL0059J
RSL0060J
RSL0061J
RSL0062J
RSL0063J
RSL0064J
RSL0065J
RSL0066J
RSL0067J
RSL0068J
RSL0069J
RSL0070J
RSL0071J
RSL0072J
RSL0073J
RSL0074J
RSL0075J
RSL0076J
RSL0077J
RSL0078J
RSL0079J
RSL0080J
RSL0081J
RSL0082J
RSL0083J
RSL0084J
RSL0085J
RSL0086J
RSL0087J
RSL0088J
RSL0089J
RSL0090J
RSL0091J
RSL0092J
RSL0093J
RSL0094J
RSL0095J
RSL0096J
RSL0097J
RSL0098J
RSL0099J
RSL0100J
RSL0101J
RSL0102J
RSL0103J
RSL0104J
RSL0105J
RSL0106J
RSL0107J
RSL0108J
RSL0109J
RSL0110J

```

```

      SMR2(L)=SMR2(P-J+1)
      J=J+1
      IF(J.LE.P) GO TO 130
      WRITE(6,165)
135  FORMAT('THE VARIABLES (TOP) ARE ORDERED BY THE SUM-OF-SQUARED
      CRITERION (BOTTOM) :')
      WRITE(6,24)(VARNO(J),J=1,P)
      WRITE(6,23)(FRNK(J),J=1,P)
24  FORMAT(13I10)
23  FORMAT(13F10.4)
      G=G+1
      FRNK(J)=VARNO(J)
      BEST=VARNO(1)
      DO 155 J=1,P
      DO 155 K=1,P
      IF(J.E.K)GO TO 135
      DIAG(G)=DIAG(G)+P(K,J)
155  R(K,J)=R(K,BEST)+R(BEST,J)/R(BEST,BEST)
      DO 160 J=1,P
      SMR2(J)=0.0
      DO 160 K=J,P
      R(J,K)=R(J,K)+R(J,K)
      R(K,J)=R(J,K)
      R(K,J)=R(J,K)
160  CONTINUE
      WRITE(6,*)' '
      WRITE(6,165)G
165  FORMAT('I4,I4,'VARIABLES & THEIR CORRELATIONS HAVE',
      &' BEEN REMOVED.')
      WRITE(6,*)' '
      TPCT(G)=100*(DIAG(G)/DIAG(1))
      IF(G.EQ.P)GO TO 900
      IF(G.EQ.CYCLE)GO TO 950
      IF(TPCT(G).LT.FLI)GO TO 965
      GO TO 170
900  WRITE(6,*)' '
      WRITE(6,910)
910  FORMAT('THE CORRELATION MATRIX IS NOW EXHAUSTED.')
      GO TO 960
965  WRITE(6,*)' '
      WRITE(6,970)CPCT
970  FORMAT('F4.0,I4,'% OF THE CORRELATION STRUCTURE HAS BEEN',
      &' ACCOUNTED FOR.')
      GO TO 960
950  WRITE(6,955)CYCLE
955  FORMAT('-',I4,I4,'CYCLES HAVE BEEN DONE.')
960  WRITE(6,*)' '
      WRITE(6,920)
920  FORMAT('THE BEST ORDER IN WHICH TO CHOOSE VARIABLES, FOR',
      &' MAXIMUM INFORMATION ABOUT STRUCTURE IN THE DATA SET, IS:')
      WRITE(6,930)(FRNK(I),I=1,G)
930  FORMAT(16I8)
      WRITE(6,*)' '
      WRITE(6,935)

```

```

RSL01110
RSL01120
RSL01130
RSL01140
RSL01150
RSL01160
RSL01170
RSL01180
RSL01190
RSL01200
RSL01210
RSL01220
RSL01230
RSL01240
RSL01250
RSL01260
RSL01270
RSL01280
RSL01290
RSL01300
RSL01310
RSL01320
RSL01330
RSL01340
RSL01350
RSL01360
RSL01370
RSL01380
RSL01390
RSL01400
RSL01410
RSL01420
RSL01430
RSL01440
RSL01450
RSL01460
RSL01470
RSL01480
RSL01490
RSL01500
RSL01510
RSL01520
RSL01530
RSL01540
RSL01550
RSL01560
RSL01570
RSL01580
RSL01590
RSL01600
RSL01610
RSL01620
RSL01630
RSL01640
RSL01650

```

```

935 FORMAT('THE TRACES ASSOCIATED WITH THE RESIDUAL',
& ' CORRELATION MATRICES ARE (BEGINNING WITH 0 VARIABLES',
& ' REMOVED):')
WRITE(6,940)((DIAG(I),I=1,6)
940 FORMAT(16F3.3)
WRITE(6,'')
WRITE(6,975)
975 FORMAT('THE TRACES, AS A PERCENTAGE OF P, ARE:')
WRITE(6,930)((PCT(I),I=1,6)
980 FORMAT(16F3.1)
STOP
END

```

```

RSL01660
RSL01670
RSL01680
RSL01690
RSL01700
RSL01710
RSL01720
RSL01730
RSL01740
RSL01750
RSL01760
RSL01770

```



```

DO 52 L=1,20
52 MOR(L,IY)=12
53 TICX=0.1
   TICY=0.1
   IF(XVAL.GT.2) TICX=0.5
   IF(XVAL.GT.12) TICX=1
   IF(XVAL.GT.47) TICX=10
   IF(YVAL.GT.2) TICY=0.5
   IF(YVAL.GT.12) TICY=1
   IF(YVAL.GT.47) TICY=10
   START=0.0
   IF(ICODEX.NE.1) GO TO 54
   IF(ICODEY.NE.1.AND.XMAX.LT.0.0) GO TO 52
61 START=START+TICX
   IF(START.GT.XMAX) GO TO 52
   IF(START.GT.XMAX) GO TO 52
   LX=(START-XMIN+UNX)/JIX
   IF(LX.GT.14) LX=14
   IF(LX.LT.1) LX=1
   MOR(IX,LX)=12
   GO TO 61
62 START=0.0
   IF(ICODEY.NE.1.AND.XMIN.GT.0.0) GO TO 64
63 START=START-TICX
   IF(START.LT.XMIN) GO TO 64
   IF(START.GT.XMAX) GO TO 63
   LX=(START-XMIN+UNX)/JNX
   IF(LX.LT.1) LX=1
   IF(LX.GT.14) LX=14
   MOR(IX,LX)=12
   GO TO 63
64 START=0.0
   IF(ICODEY.NE.1) GO TO 74
   IF(ICODEX.NE.1.AND.YMAX.LT.0.0) GO TO 70
65 START=START+TICX
   IF(START.GT.YMAX) GO TO 70
   IF(START.LT.YMIN) GO TO 65
   LY=18.7-(START-YMIN+UNY)/UNY
   IF(LY.LT.1) LY=1
   IF(LY.GT.18) LY=18
   MOR(LY,IY)=11
   GO TO 65
70 START=0.0
   IF(ICODEX.NE.1.AND.YMIN.GT.0.0) GO TO 74
71 START=START-TICX
   IF(START.LT.YMIN) GO TO 74
   IF(START.GT.YMAX) GO TO 71
   LY=18.0-(START-YMIN+UNY)/UNY
   IF(LY.GT.18) LY=18
   IF(LY.LT.1) LY=1
   MOR(LY,IY)=11
   GO TO 71
74 IF(ICODEY.NE.1) TICX=0.0
   IF(ICODEX.NE.1) TICX=0.0
   WRITE(6,110) I,J,XMIN,XMAX,JNX,TICX,YMIN,YMAX,UNY,TICX

```

```

PL000560
PL000570
PL000580
PL000590
PL000600
PL000610
PL000620
PL000630
PL000640
PL000650
PL000660
PL000670
PL000680
PL000690
PL000700
PL000710
PL000720
PL000730
PL000740
PL000750
PL000760
PL000770
PL000780
PL000790
PL000800
PL000810
PL000820
PL000830
PL000840
PL000850
PL000860
PL000870
PL000880
PL000890
PL000900
PL000910
PL000920
PL000930
PL000940
PL000950
PL000960
PL000970
PL000980
PL000990
PL001000
PL001010
PL001020
PL001030
PL001040
PL001050
PL001060
PL001070
PL001080
PL001090
PL001100

```

```

110 FORMAT (//3X,'HORIZONTAL AXIS IS DIMENSION',I3/
X 3X,'VERTICAL AXIS IS DIMENSION',I5/3X///10X,'HORIZONTAL AXIS'/
X/3X,'MINIMUM VALUE=',F15.5/3X,'MAXIMUM VALUE=',F15.5/3X,
X 'SCALING UNIT =',F15.5/3X,'ONE TICK=',F10.3///10X,
X*VERTICAL AXIS//3X,'MINIMUM VALUE=',F15.5/3X,'MAXIMUM VALUE=',
XF15.5/3X,'SCALING UNIT =',F15.5/3X,'ONE TICK=',F10.3//3X,
X 'OVERLAPPING OBJECTS (NOT PLOTTED)//3X,'IDENTIFIER',3X,
X 'COORDINATES'/)
DO 100 L=1,M
X=V(I,L)
Y=V(J,L)
IX=(X-XMIN+UMX)/UMX
IY=(Y-YMIN+UNY)/UNY
IF(IX.GT.1) IX=1
IF(IX.LT.1) IX=1
IF(IY.GT.1) IY=1
IF(IY.LT.1) IY=1
IF(L.GE.100) GO TO 700
IF(L.GE.10) GO TO 700
IF(MOD(IY,IX).LE.9) GO TO 300
MOD(IY,IX)=L
GO TO 100
750 IF(IX.EQ.1) IX=IX-1
IF(MOD(IY,IX).LE.9) OR MOD(IY,IX+1).LE.9) GO TO 300
MOD(IY,IX)=L/10
MOD(IY,IX+1)=L-(L/10)*10
GO TO 100
790 IF(IX.EQ.1) IX=IX-2
IF(IX.EQ.1) IX=IX-1
IF(MOD(IY,IX).LE.9) OR MOD(IY,IX+1).LE.9) OR MOD(IY,IX+2).LE.9)
X GO TO 800
MOD(IY,IX)=L/100
MOD(IY,IX+1)=(L-(L/100)*100)/10
MOD(IY,IX+2)=L-(L/10)*10
GO TO 100
800 WRITE(6,901) L,X,Y
801 FORMAT (I9,5F15.5)
100 CONTINUE
111 FORMAT (3X,125A1)
115 FORMAT (1H1)
WRITE(6,115)
DO 113 JJ=1,I+1
113 PLOTT(JJ)=PLS(14)
WRITE(6,111) PLS(14), (PLOTT(JJ),JJ=1,I+1), PLS(14)
DO 200 K=1,I+1
DO 150 L=1,I+1
150 PLOTT(L)=PLS(MOD(K,L)+1)
200 WRITE(6,111) PLS(1+), (PLOTT(KL), KL=1,I+1), PLS(1+1)
DO 201 JJ=1,I+1
201 PLOTT(JJ)=PLS(14)
WRITE(6,111) PLS(14), (PLOTT(JJ),JJ=1,I+1), PLS(14)
STOP
END

```

PL001110  
PL001120  
PL001130  
PL001140  
PL001150  
PL001160  
PL001170  
PL001180  
PL001190  
PL001200  
PL001210  
PL001220  
PL001230  
PL001240  
PL001250  
PL001260  
PL001270  
PL001280  
PL001290  
PL001300  
PL001310  
PL001320  
PL001330  
PL001340  
PL001350  
PL001360  
PL001370  
PL001380  
PL001390  
PL001400  
PL001410  
PL001420  
PL001430  
PL001440  
PL001450  
PL001460  
PL001470  
PL001480  
PL001490  
PL001500  
PL001510  
PL001520  
PL001530  
PL001540  
PL001550  
PL001560  
PL001570  
PL001580  
PL001590  
PL001600  
PL001610  
PL001620  
PL001630



APPENDIX III - COUNTRY REPORTS

III.1 Indonesia

THE APPLICATION OF COMPUTER AT THE INDONESIAN  
INSTITUTE OF SCIENCES AND THE UNIVERSITIES IN INDONESIA

Tri Surja Kreshnawati  
(LIPI Jakarta)

S. Djalal Tandjung  
(UGM Yogyakarta)

Introduction

LIPI is a Government body which provides guidance in the field of scientific and technological research. It reports directly to the President of the Republic of Indonesia.

LIPI has ten national research institutions situated in Jakarta, Bogor, Bandung, and Serpong which are conducting research in the natural, technological and social science. There is also a National Scientific Documentation Centre.

The national research institution administered by LIPI are:

- National Biological Institute
- National Institute of Oceanology
- National Institute of Geology and Mining
- National Institute for Chemistry
- National Institute for Physics
- National Institute for Metalurgy
- National Institute for Electrotechniques
- National Institute for Instrumentation
- National Institute for Economic & Social Research
- National Institute for Cultural Studies

Computer in LIPI

The rapid advance of Science and Technology in the last two decades can be attributed mostly to the intelligent use of computers in data handling and analysis.

A computer is used because it does certain task and ability better and more efficiently than mankind. The characteristics of this machine are speed and capacity to handle large volumes of data in a very short time. It is far from exaggeration that computers in the advancement of Science and Technology are indispensable. Each of the national research institute use computers for R & D activities. In this case, we describe one of the institute is National Biological Institute, and in addition some information on the usage of computers in higher education Institutions in Indonesia.

#### National Biology Institute - LIPI

The National Biology Institute has an Apple II computer with 48 K. capacity and a silent type printer. The printer can print 132 characters and has the capacity to print graphics.

Available computer programmes are as follows:

1. Visifile for information on management data.
2. Visitrend for analysis and graphics.
3. Visicalc for genetic pool collection.
4. Abstat for statistic analysis.
5. Utilities for visifile.
6. DOS 3.3.

At the present time the National Biology Institute has computerised diskettes for:

1. Documental ethnobotany collection.
2. Botanical Garden collection.
3. Genetic pool garden collection.
4. Herbarium Bogoriense collection.
5. Zoology Museum collection.
6. Ecology research.
7. Taxonomy.

The data discrete programme storage specifications are:

- a. One data sheet.
- b. 24 column for one file, with 232 characters.

Example :

### Ethnobotany collection

Registration	number	collector	collector	number	date
location	region	Name of thing	material	plant	useful

The steps are:

1. formulate the format.
2. data entry.
3. data storage.

The data storage can be used at any time. Based on the example, the data can be processed as it is needed, for example:

- What kind of matter at the vitrin 7
- What kind of collection from West Kalimantan
- For what purpose the rottan are used, etc.

### Botanical Garden Collection

Family	Species	type/variety	Island	Location	Altitude
No. of plan	Date of plan	Herbarium	Blooming	Fertilization	

From the data entry can be used for:

- What is the number of Herbarium material.
- What kind of collection from Sumatra.
- How many Pterospermum javanicum is grown.
- When was the Eucalyptus alba planted.

### Plant Ecology

Plotting	species	family	diameter	basal area	unbranched trunk
Total high	topography	soil			

The data can be used to determine:

- What kind of species has diameter of 50cm.
- What kind of species belong to the group of Myrtaceae.

Herbarium specimen

Registration Number	Family Number	Species	Local name	Island
location	height	habitat	collector	Date

- How many genus and species are kept in the Herbarium Bogoriense.
- What species are collected from Sumatra Barat.
- What species are found only at high elevation e.g. 750 meters above sea level.
- What are the Orchidacea family fund in Sumatra.

The computer is also used for finalised legume data sheet as below.

Legume data sheetCollection Data

Accession number	Scientific name	Local/English name
Collection number	Collection date	Collection site
Material collected	Occurrence	Uses

Evaluation Data

Habitat	Plant type	Life duration	leaf type	leaflet shape
Flower colour	Pod type	Pod shape	Pod texture	Pod colour
Seed shape	Seed colour	Tuber	Flowering time	Age of first flower
Pod setting	Pod length	Number of seeds per pod	100	
seed weight	Disease resistance	Pest tolerance		

Additional notes

Those are examples of several usages of computer in the National Biology Institute of LIPI.

### The Application of Computer in the Universities in Indonesia

This is not an official information based on any research or survey. To the authors knowledge, some big universities such as Gadjah Mada University in Yogyakarta and Indonesia University in Jakarta have used computers in their work.

Gadjah Mada University has a computer center, which can be used for education and research by students and teaching staffs. Student from Faculty of Mathematics and Science have to take subjects on computer. Other students from other faculty use the computer as it is needed for data processing of their research. So far, computers have been used in many universities for education and research.

While the new generation of students (started with the year 1975) have the ability to operate the computers, their professors are left far behind, because in their age, when they were students, they did not get any computer training. Now the professors have to catch up today's computer technology.

### Conclusion

LIPPI and higher education institutions in Indonesia have started using computers in their work. More staff have to be trained to handle and be familiar with the computer.

## III.2 Malaysia

THE STATUS OF COMPUTER HARDWARE AND SOFTWARE FACILITIES  
AND THE USE OF COMPUTERS FOR RESEARCH IN ENVIRONMENTAL BIOLOGY  
IN MALAYSIA

Kam Suan Pheng  
(Universiti Sains Malaysia)

Roslan bin Ismail  
(Forest Research Institute)

## A. COMPUTER HARDWARE FACILITIES

A number of universities and research institutions in Malaysia are involved in biology and applied biology, and most of these institutions are equipped with some model of mainframes. The following table summarises the mainframes available at the various institutions, to the best of our knowledge. Therefore, this list is not exhaustive.

Institution	Mainframes and superminis
FRI	Data General Eclipse S140
MARDI	IBM
PORIM	HP 3000
RRIM	HP 3000
UM	IBM
UKM	IBM, PRIME
UPM	UNIVAC
USM	IBM 4331 & IBM 4381
UTM	IBM
IMR	IBM
FRI	Forest Research Institute
MARDI	Malaysian Agricultural Research and Development Institute
PORIM	Palm Oil Research Institute of Malaysia
PRIM	Rubber Research Institute of Malaysia

---

UM	Universiti Malaya (University of Malaya)
UKM	Universiti Kebangsaan Malaysia (National University of Malaysia)
UPM	Universiti Pertanian Malaysia (Agricultural University of Malaysia)
USM	Universiti Sains Malaysia (University of Science, Malaysia)
UTM	Universiti Teknologi Malaysia (University of Technology Malaysia)
IMR	Institute for Medical Research

---

Apart from the mainframes and super-minis, some of the institutions have mini-computers and micro-computers. Besides the government and quasi-government bodies listed above, there are also private research laboratories, such as those associated with the plantation and pesticide companies, which utilise computers in their research activities.

## B. SOFTWARE AND PROGRAMS

The main frames available in the institutions listed above are normally used for large projects with big data sets. Programs are either written, in FORTRAN or BASIC, or software packages such as SAS, SPSS, BMDP, and GENSTAT are used. In a number of research organisations we know, the SAS package is preferred by scientists, especially biologists, because of its greater utility for analysing scientific data.

Specifically, we know of the use of computers for research purposes for the two institutions which we come from:

### 1. Forest Research Institute

- a. Prefelling and post-felling inventorisations
- b. Growth and yield studies
- c. Numerous other aspects of forestry research

## III.3 Philippines

STATUS OF COMPUTER APPLICATIONS  
TO ECOLOGICAL RESEARCHES/PROJECTS  
WITHIN THE MAB PROGRAMME IN THE PHILIPPINES

Christian P. Dizon

Jesus P. Bayrante

(Man &amp; the Biosphere Inter-Agency Committee on Ecological Studies)

The Man and the Biosphere Inter-Agency Committee on Ecological Studies (MAB-ICES) in the Philippines implements its programme of research through cooperation and collaboration with its fourteen (14) government member-agencies and three (3) cooperating agencies which are all involved in resource management (see Attachment). As of March 1985, the MAB-ICES has continued to undertake at least nine (9) national ecological field projects for cooperative research all within the framework of the MAB Programme. Such researches are classified under UNESCO-MAB's international themes (i.e. forest areas, coastal zones, pollution, energy utilization, environmental impact assessment and biosphere reserves).

Quantitative analyses of significant ecological/environmental data generated from the various researches/projects within the MAB programme using the computers have not been widely used due to the lack of computer hardwares/machine/gadgets and technical personnel with the proper training who could easily process the data with the computers using the various quantitative/statistical packages being utilized by other countries.

An inventory of the kinds/types of mainframe computer hardwares and the corresponding softwares used by MAB member-agencies in which MAB researchers/scientists could have access to, revealed that there are more or less five (5) agencies with the mainframe computer hardwares. These hardwares are of the IBM (e.g. IBM 1130, etc.) and in most cases, the FORTRAN language is used.



In terms of micro or minicomputers, the Apple II-E, Apple II-Plus and IBM PC Compatible are the most common. The operating systems utilized are the TRS-DOS by Tandy, CP/M, COMMODORE DOS and MS-DOS.

Based on the foregoing, there is a need to expose the researchers/scientists within the MAB Programme in the Philippines to the current and perhaps advanced statistical packages using the above mentioned computers especially the micros. With this, data handling, storage, processing and analysis would be facilitated.

#### MAB PHILIPPINES GOVERNMENT MEMBER AGENCIES

1. Bureau of Plant Industry
2. Bureau of Animal Industry
3. Bureau of Soils
4. Bureau of Lands
5. Bureau of Mines and Geosciences
6. Bureau of Fisheries and Aquatic Resources
7. Bureau of Forest Development
8. Bureau of Coast and Geodetic survey
9. National Institute of Science and Technology
10. National Museum
11. Philippine Atmospheric, Geophysical, and Astronomical Services Administration
12. National Irrigation Administration
13. Ministry of Public Works & Highways
14. Philippine Coast Guard

#### COOPERATING AGENCIES

1. National Pollution Control Commission
2. Forest Research Institute
3. National Water Resources Council

## III.4 Singapore

THE USE OF COMPUTERS IN THE SCHOOLS OF SINGAPORE  
AND IN THE DEPARTMENT OF ZOOLOGY, NUS

Choo Bee Li  
(National University of Singapore)

Tan Siok Cheng  
(Curriculum Development Institute of Singapore)

The Singapore government started promoting computer awareness in the schools in 1980. Ample funds were allocated to be various educational institutions to purchase computers and train personnel to meet the demands of a sophisticated technological era. This report touches on the present computer situation in the various educational institutions of Singapore.

## TEACHER TRAINING

The Institute of Education has a computer laboratory and conducts computer literacy courses for primary school teachers. It also has two terminals attached to the mainframe computer at the Ministry of Education for the use of its staff, trainee teachers and M Ed students.

The Curriculum Development Institute of Singapore's computer department has a computer laboratory which is well stocked with many IBM and a few Apple micro-computers. It conducts computer literacy lessons in BASIC to secondary school teachers and courses on the use of various software packages such as Logo for primary school teachers and Superpilot and dBaseII for secondary school teachers.

## THE SCHOOLS

The junior colleges have their own computer laboratories and student can opt to take computer science as an 'A' Level examination subject.

Each secondary school in Singapore has three to ten micro-computers. Some SAP (Special Assistance Plan) schools have

as many as twenty-five. These computers belong to the schools' computer clubs which normally conduct computer appreciation courses for their members. Some SAP schools give compulsory computer literacy lessons to their students.

The staff of some schools use computers to compute their school records and examination results.

Most of the primary schools do not have computers. The CDIS CAI (Computer Assisted Instruction) Project Team is preparing, for a start, a computer laboratory in one primary school. It should be ready by this July. It will have a mainframe computer and twenty-four on-line terminals. The project team intends to introduce CAI packages in mathematics, mainly of the drill and practice type to the weaker students in the primary schools.

#### THE DEPARTMENT OF ZOOLOGY, NUS

The Department of Zoology of the National University of Singapore has about eighteen micro-computers and two mainframe terminals. The micro-computers are used mainly for teaching. For example, the fisheries courses for third and honours year students make extensive use of the computers. As micro-computers have small memory spaces, they are only used to analyse simple and small data sets. The micro-computer is also used to catalog the specimens in the Zoological Record Collection of the department.

The mainframe terminals with their more powerful software packages such as SAS and Minitab are well utilised by the staff and students of the department. The software packages can perform complex data manipulations such as multivariate statistical analysis.

## III.5 Thailand

## A REPORT FROM THE PARTICIPANTS OF THAILAND

Santad Koompalum  
(National Environment Board)

Rojchai Satrawaha  
(Khon Kaen University)

Air and noise pollution section, environmental quality standard division, office of the National Environment board (ONEB) is responsible for technical data, policy determination and management of air and noise pollution in Thailand.

In a field of technical data, involved the monitoring of ambient air. There are 8 monitoring stations located in Bangkok area and a mobile monitoring unit is used to monitor air quality in other main cities and other areas which have air pollution problems. Other data include air pollution emission from motor vehicles. From industrial plants, noise and vibration data. Most of the analyzed data are assessed to provide input to the special committee for the determination of air quality standards for ambient air quality. Emission from motor vehicles and emission from industries. Some of the data is also used as information for other government unit and public sector which are concerned with air and noise pollution problems and control.

Raw data are collected continuously by automatic air pollutant analyzers for carbon monoxide, hydrocarbons, sulfur dioxide, oxides of nitrogen, oxidants and suspended particulate matter as charts from recorders and as data cassette tape recorder from dataloggers. Other raw data are from the meteorological department. Traffic volume and industrial information are also obtained.

There are three microcomputer systems used in ONEB at the present. Now a Fujitsu micro-8 computer system and data cassette recorder are used in air and noise pollution section. Some of the software are developed in F-Basic language. Other packages include DBase II, Supercalc, Wordstar and Fortran-86 (16 Bit)

under CP/M. There are two programmers with B.Sc. in statistics who operate the computer.

The Apple IIe and victor 4 system are used in water quality section and solid waste section respectively, ONEB is about to purchase the two 16 bit microcomputer systems under eastern seaboard project and ONEB also plans to have a mini computer to be used as environmental information center and data base for Thailand in the near future.

At the Faculty of Science, Khon Kaen University, there are fifteen Apple II microcomputers which can be used by the university staff. Environmental biologists usually do not have much background on computers. Analysis of biological data is mostly done with assistance from the mathematics and statistics department staff. However, Khon Kaen University has a plan to set up a computer center with mainframe facilities in the near future.

## APPENDIX IV - LIST OF PARTICIPANTS

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